

# AN APPROXIMATE STOCHASTIC MODEL FOR PHAGE REPRODUCTION IN A BACTERIUM

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## 1. Introduction

The stochastic birth-death process considered in this paper provides an approximate model for phage reproduction in a bacterium. In a recent paper, Hershey [1] has discussed reproduction and recombination in phage crosses, and a deterministic model for the reproductive process has been the subject of a previous note by the author [2]. A very readable account of the process is given by Sanders [3] in his recent article, "*The life of viruses*".

Consider a medium initially containing  $N < \infty$  bacteria, which we may for simplicity assume to reproduce as a birth process with constant parameter. Into this medium is inserted one phage particle (or several), which immediately penetrates (infects) a bacterium, and proceeds to reproduce within it as a birth-death process with constant parameters  $\lambda > \mu$ . After a phage has invaded a bacterium, changes occur on the surface of the bacterium to prevent its penetration by further phages. Usually a single phage invades a bacterium, but towards the end of the infective process when the number of phages and bacteria are of the same order, there is a larger probability that a bacterium is infected by two or more phages simultaneously; in our model, this probability is taken to be negligible. The death of a phage corresponds in fact to its reaching maturity, after which it no longer reproduces. When a fixed number  $r$  (two to three hundred) of these mature phages have been produced, the bacterium, which is itself incapable of fission after infection by a phage, breaks open as it dies, releasing the phage offspring. Immature phages cannot attack bacteria, but the  $r$  mature phages immediately penetrate a further  $r$  uninfected bacteria; the phages reproduce faster than the bacteria, and this sequence of processes usually continues until the bacteria are all dead.

Let us now discuss the phage reproduction occurring in a single bacterium. We should like to find the distribution of time  $T_r$  up to the occurrence of the  $r$ -th death in an ordinary birth-death process: this is not in fact known. Assuming for the moment that the process continues indefinitely, without

stopping when  $r$  mature phages are produced, it is possible to write the differential equations for the probabilities.

$$P_{ij}(t) = Pr\{i \text{ survivals and } j \text{ deaths in time } t\}$$

in the form

$$(1.1) \quad P'_{ij}(t) = (i - 1)\lambda P_{i-1,j}(t) - i(\lambda + \mu)P_{ij}(t) + (i + 1)\mu P_{i+1,j-1}(t)$$

for all values of  $i, j = 0, 1, 2, \dots$ , with  $P_{ij}(t)$  identically zero for  $i = j = 0$ , and  $i$  or  $j < 0$ .

If we further define the joint probability generating function (p.g.f.)  $\phi(u, v, t)$  as

$$(1.2) \quad \phi(u, v, t) = \sum_{i,j=0}^{\infty} P_{ij}(t)u^i v^j \quad (|u|, |v| \leq 1),$$

the equations (1.1) lead to the partial differential equation

$$(1.3) \quad \frac{\partial \phi}{\partial t} = \{\lambda u^2 - (\lambda + \mu)u + \mu v\} \frac{\partial \phi}{\partial u},$$

which, when  $v = 1$ , reduces to the well-known equation for the generating function of probabilities of survivals. The equation (1.3) can be solved: solutions have in fact been obtained for somewhat different but equivalent forms of it by Kendall [4] and Bartlett [5]. However, expansion of the p.g.f. is unwieldy, and it seems difficult to obtain the  $P_{ij}(t)$  explicitly from it. A different approach using the Laplace transforms with respect to time

$$(1.4) \quad q_{ij}(s) = \int_0^{\infty} e^{-st} P_{ij}(t) dt \quad (R(s) > 0)$$

leads to the relations

$$(1.5) \quad \{(\lambda + \mu) + s\}q_{10}(s) = 1 \quad \text{for } i = 1, j = 0,$$

$$\{i(\lambda + \mu) + s\}q_{ij}(s) = (i - 1)\lambda q_{i-1,j}(s) + (i + 1)\mu q_{i+1,j-1}(s) \text{ for all other } i, j,$$

with  $q_{ij}(s)$  identically zero for  $i = j = 0$ , and  $i$  or  $j < 0$ . These are again not readily solved.

## 2. An approximation to the birth-death process

It is natural at this point to approximate to the standard birth-death process by one for which it is possible to obtain the probabilities  $P_{ij}(t)$  explicitly. Such a process is that where births occur in a non-homogeneous Poisson process, the probability of a birth in the interval  $(t, t + \delta t)$  being

$$(2.1) \quad \lambda e^{\alpha t} \delta t + o(\delta t)$$

where  $\alpha = \lambda - \mu > 0$  and  $e^{\alpha t}$  is the mean number of survivals at time  $t$  in the standard birth-death process. The death process remains unchanged. We should now strictly refer to probabilities  $P_{ij}(0, t)$  and the p.g.f.,  $\phi(u, v; 0, t)$ , since the process is no longer homogeneous in time; for simplicity, however, we retain the notations  $P_{ij}(t)$ ,  $\phi(u, v, t)$  which are quite clear in this case.

The forward differential equations for these probabilities are now

$$(2.2) \quad P'_{ij}(t) = \lambda e^{\alpha t} P_{i-1,j}(t) - (\lambda e^{\alpha t} + i\mu) P_{ij}(t) + (i+1)\mu P_{i+1,j-1}(t)$$

for all values of  $i, j = 0, 1, 2, \dots$ , with  $P_{ij}(t)$  identically zero for  $i = j = 0$  and  $i$  or  $j < 0$ . The joint p.g.f.,  $\phi(u, v, t)$ , satisfies the equation

$$(2.3) \quad \frac{\partial \phi}{\partial t} + \mu(u - v) \frac{\partial \phi}{\partial u} = \lambda e^{\alpha t} (u - 1) \phi,$$

which we proceed to solve. If we perform the transformation  $T = (u - v)e^{-\mu t}$  leaving  $u, v$  unchanged, and write

$$F(u, v, T) = \phi(u, v, t)$$

we obtain from (2.3) that

$$(2.4) \quad \frac{\partial F}{\partial u} = \frac{\lambda}{\mu} (u - 1)(u - v)^{(\alpha/\mu)-1} T^{-\alpha/\mu} F(u, v, T).$$

The solution to this equation is of the form

$$(2.5) \quad \begin{aligned} F(u, v, T) &= \left\{ \exp \frac{\lambda}{\mu} T^{-\alpha/\mu} \int_{u_0}^u (x - 1)(x - v)^{(\alpha/\mu)-1} dx \right\} f(v, T) \\ &= \left\{ \exp T^{-\alpha/\mu} \left[ (u - v)^{\lambda/\mu} + \frac{\lambda}{\alpha} (v - 1)(u - v)^{\alpha/\mu} - A \right] \right\} f(v, T) \end{aligned}$$

where  $A$  is a function of  $v$  which later vanishes. Rewriting this as  $\phi(u, v, t)$ , we find that

$$(2.6) \quad \phi(u, v, t) = \left\{ \exp e^{\alpha t} \left[ (u - v) + \frac{\lambda}{\alpha} (v - 1) - A (u - v)^{-\alpha/\mu} \right] \right\} f(v, (u - v)e^{-\mu t})$$

The initial condition  $P_{10}(0) = 1$  results in

$$(2.7) \quad \phi(u, v, 0) = u = \left\{ \exp \left[ (u - v) + \frac{\lambda}{\alpha} (v - 1) - A (u - v)^{-\alpha/\mu} \right] \right\} f(v, u - v)$$

whence it follows that

$$(2.8) \quad f(v, T) = (T + v) \exp - \left\{ T + \frac{\lambda}{\alpha} (v - 1) - AT^{-\alpha/\mu} \right\}.$$

Thus  $\phi(u, v, t)$  can finally be written as

$$\begin{aligned} \phi(u, v, t) &= \{v + (u - v)e^{-\mu t}\} \exp\left\{e^{\alpha t} \left[ (u - v) + \frac{\lambda}{\alpha} (v - 1) - A(u - v)^{-\alpha/\mu} \right] \right. \\ &\quad \left. - \left[ (u - v)e^{-\mu t} + \frac{\lambda}{\alpha} (v - 1) - A(u - v)^{-\alpha/\mu} e^{\alpha t} \right] \right\} \\ &= \{ue^{-\mu t} + v(1 - e^{-\mu t})\} \exp\left\{-\frac{\lambda}{\alpha} (e^{\alpha t} - 1)\right. \\ (2.9) \quad &\quad \left. + u(e^{\alpha t} - e^{-\mu t}) + v \left[ \frac{\lambda}{\alpha} (e^{\alpha t} - 1) - (e^{\alpha t} - e^{-\mu t}) \right] \right\} \\ &= \{ue^{-\mu t} + v(1 - e^{-\mu t})\} \exp\{-\rho(t) + u\Lambda(t) + v[\rho(t) - \Lambda(t)]\} \\ &= e^{-\rho(t)} \left\{ \sum_{i=0}^{\infty} \sum_{j=0}^{\infty} u^i v^{j+1} (1 - e^{-\mu t}) \frac{\Lambda^i (\rho - \Lambda)^j}{i! j!} \right. \\ &\quad \left. + \sum_{i=0}^{\infty} \sum_{j=0}^{\infty} u^{i+1} v^j e^{-\mu t} \frac{\Lambda^i (\rho - \Lambda)^j}{i! j!} \right\} \end{aligned}$$

where for  $t > 0$ ,  $\rho(t) = (\lambda/\alpha)(e^{\alpha t} - 1) > 0$ ,  $\Lambda(t) = e^{\alpha t} - e^{-\mu t} = (1 + \alpha\rho/\lambda)^{-\mu/\alpha} \{ (1 + \alpha\rho/\lambda)^{\lambda/\alpha} - 1 \} > 0$ , and it is easily shown that  $\rho(t) - \Lambda(t) > 0$ .

The p.g.f. for the probabilities of survivals, is given by

$$(2.10) \quad \phi(u, 1, t) = \{ue^{-\mu t} + (1 - e^{-\mu t})\} e^{-\Lambda(t)(1-u)},$$

while that for the probabilities of deaths is

$$(2.11) \quad \phi(1, v, t) = \{v(1 - e^{-\mu t}) + e^{-\mu t}\} e^{-\{\rho(t) - \Lambda(t)\}(1-v)}.$$

### 3. Explicit probabilities of births, survivals and deaths

For the birth-death process continuing indefinitely, without stopping when  $r$  mature phages have been produced, it follows from (2.9) directly that the probabilities  $P_{ij}(t)$  are

$$\begin{aligned} P_{0j}(t) &= e^{-\rho(t)} (1 - e^{-\mu t}) \frac{\{\rho(t) - \Lambda(t)\}^{j-1}}{(j-1)!} \quad (j \geq 1) \\ (3.1) \quad P_{i0}(t) &= e^{-\rho(t) - \mu t} \frac{\{\Lambda(t)\}^{i-1}}{(i-1)!} \quad (i \geq 1) \\ P_{ij}(t) &= e^{-\rho(t)} \left\{ (1 - e^{-\mu t}) \frac{\Lambda^i (\rho - \Lambda)^{j-1}}{i! (j-1)!} + e^{-\mu t} \frac{\Lambda^{i-1} (\rho - \Lambda)^j}{(i-1)! j!} \right\} \quad (i, j \geq 1) \end{aligned}$$

Clearly the probability  $B_k(t)$  of  $k$  births in time  $t$  is of the Poisson form

$$(3.2) \quad \begin{aligned} B_k(t) &= \exp\left\{-\int_0^t \lambda e^{\alpha\tau} d\tau\right\} \left\{\int_0^t \lambda e^{\alpha\tau} d\tau\right\}^k / k! \\ &= e^{-\rho(t)} \{\rho(t)\}^k / k! \quad (k \geq 0), \end{aligned}$$

the mean number of births in time  $t$  being  $\rho(t)$ . From (2.10), the probabilities  $S_m(t)$  of  $m$  survivals after time  $t$  are

$$(3.3) \quad \begin{aligned} S_0(t) &= (1 - e^{-\mu t})e^{-\Lambda(t)} \\ S_m(t) &= e^{-\Lambda(t)} \left\{ e^{-\mu t} \frac{\Lambda^{m-1}}{(m-1)!} + (1 - e^{-\mu t}) \frac{\Lambda^m}{m!} \right\} \quad (m \geq 1) \end{aligned}$$

It should be noted that the mean number of survivals after time  $t$  for this process is

$$(3.4) \quad \frac{\partial}{\partial u} \phi(1, 1, t) = e^{\alpha t},$$

exactly as for the standard birth-death process, while its variance  $e^{\alpha t}(1 - e^{-(\lambda+\mu)t})$  is less than the variance  $(\lambda + \mu)(\lambda - \mu)^{-1}e^{\alpha t}(e^{\alpha t} - 1)$  of the standard process. Similarly from (2.11) the probabilities  $D_r(t)$  of  $r$  deaths in time  $t$  are

$$(3.5) \quad \begin{aligned} D_0(t) &= e^{-\mu t - \rho(t) + \Lambda(t)} \\ D_r(t) &= e^{-\{\rho(t) - \Lambda(t)\}} \left\{ e^{-\mu t} \frac{(\rho - \Lambda)^r}{r!} + (1 - e^{-\mu t}) \frac{(\rho - \Lambda)^{r-1}}{(r-1)!} \right\} \quad (r \geq 1), \end{aligned}$$

and the mean number of deaths in time  $t$  is

$$(3.6) \quad \frac{\partial}{\partial v} \phi(1, 1, t) = 1 - e^{\alpha t} + \rho(t) = \frac{\mu}{\alpha} (e^{\alpha t} - 1).$$

Suppose the process now stops at the  $r$ -th death, the probability distribution of this death (or maturing of the  $r$ -th phage), which is improper since there is a non-zero probability that the process ends before, is given by

$$(3.7) \quad \begin{aligned} g(t)dt &= \sum_{i=1}^{\infty} P_{i,r-1}(t) i \mu dt \\ &= \left\{ \text{Term in } v^{r-1} \text{ of } \frac{\partial}{\partial u} \phi(1, v, t) \right\} \mu dt \\ &\doteq \left\{ \text{Term in } v^{r-1} \text{ of } [e^{-\mu t}(1 + \Lambda) + \Lambda v(1 - e^{-\mu t})]e^{-(\rho - \Lambda)(1-v)} \right\} \mu dt \\ &= e^{-(\rho - \Lambda)} \left\{ e^{-\mu t}(1 + \Lambda) \frac{(\rho - \Lambda)^{r-1}}{(r-1)!} + \Lambda(1 - e^{-\mu t}) \frac{(\rho - \Lambda)^{r-2}}{(r-2)!} \right\} \mu dt \end{aligned}$$

We have thus constructed a birth-death process approximating to the

original one, but for which each of the probabilities of births, deaths and survivals is explicitly known.

If the number of bacteria  $N$  is taken to be infinite, an equation for the distribution of the number infected up to time  $t$  (or the number of mature phages released up to  $t$ ) can be obtained from Bellman and Harris' [6] theory of branching processes. If  $Q_n(t)$  is the probability of  $n$  infected bacteria, and  $\psi(s, t) = \sum_{n=1}^{\infty} Q_n(t)s^n$  the p.g.f. of this distribution, this satisfies the integral equation

$$(3.8) \quad \psi(s, t) = s(1 - G(t)) + \int_0^t \{\psi(s, t - \tau)\}^r g(\tau) d\tau$$

where  $g(t)dt$  is the improper probability distribution (3.7) and  $G(t)$  the function  $G(t) = \int_0^t g(\tau) d\tau$ . There seems to be no simple way of solving this equation.

### References

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