

Short Communication

BMI: a simple, rapid and clinically meaningful index of under-nutrition in the oldest old?

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BMI is commonly used as a sole indicator for the assessment of nutritional status. While it is a good predictor of morbidity and mortality among young and middle-aged adults, its predictive ability among the oldest old remains unclear. The objective of the present study was to investigate the relationship between BMI and risk of falls, fractures and all-cause mortality among older Australians in residential aged care facilities. One thousand eight hundred and forty-six residents of fifty-two nursing homes and thirty hostels in northern Sydney, Australia, participated in the present study. Baseline weight and height were measured and BMI (kg/m^2) calculated. For 2 years following the baseline measurements, incidence and date of all falls and fractures were recorded by research nurses who visited the facilities regularly and date of death was documented based on the participants' records at each facility. Cox proportional hazards regression models were calculated to determine the relationship between baseline BMI and time to fall, fracture or death, within 2 years following the baseline measures taken to be the censoring date. After adjustments were made for age, sex and level of care, low BMI ($<22 \text{ kg}/\text{m}^2$) increased the risk of fracture by 38 % (hazard ratio = 1.38, 95 % CI 1.11, 1.73) and all-cause mortality by 52 % (hazard ratio = 1.52, 95 % CI 1.30, 1.79). The magnitude of this effect was only slightly reduced when adjustments were further made to incorporate cognition, number of medications, falls and fracture in the subsequent 2-year period. In conclusion, BMI has predictive ability in the area of fracture and all-cause mortality for residents of aged care facilities. It is a simple and rapid indicator of nutritional status rendering it a useful nutrition screen and goal for nutrition intervention.

Nutrition: Falls: Fracture: Mortality

Ageing of the population is a global phenomenon. By the year 2050, it is predicted that the proportion of the population over the age of 60 years will double to 22 %⁽¹⁾. Health care expenditure for this subgroup of the population is disproportionately high, increasing exponentially as end of life approaches⁽²⁾. The cost of accidental falls and fall-related orthopaedic injuries contribute largely to the overall cost, with increasing frequency and severity with increasing age⁽³⁾ and the association of these events with poorer outcomes overall^(4,5).

There is evidence that achieving and maintaining a desirable nutritional status can assist in the prevention of falls and fractures and increase life expectancy⁽⁶⁾. While there is no 'gold standard' measure to define desirable nutritional status, there have been numerous screening and assessment instruments developed in an effort to capture established risk

factors and clinical signs and symptoms of malnutrition⁽⁷⁾. The complexity of these instruments vary but most require significant resources in terms of training, time to administer and interpret, and few have demonstrated an ability to predict outcomes among the oldest old. At a time when resources are scarce and expected to become even more so, questions should be asked about the need to use these relatively intensive methods of identifying malnutrition.

The BMI (kg/m^2) is commonly used across a variety of settings and age groups to describe the extremes of adiposity and the level of morbidity risk⁽⁸⁾. An index of weight for height, the BMI is included across a range of nutrition screening and assessment instruments for use among older adults despite concerns about its feasibility and predictive ability.

Abbreviations: FREE, Fracture Risk Epidemiology in the Elderly study; HR, hazard ratio.

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Concerns have been raised about the physiological changes and episodes associated with ageing that can lead to inaccuracy in the measurement of both weight and height, in addition to the variability in equipment and possible observer error associated with both measurement and interpretation⁽⁹⁾. While most of these concerns can be overcome with the use of surrogate measures of height (e.g. knee height), regular training and equipment calibration along with the use of nomograms designed to remove the need to calculate BMI, there is still uncertainty about the ability of BMI to predict meaningful outcomes among the oldest old⁽⁹⁾.

The aim of the present study was to investigate the relationship between BMI and risk of falls, fractures and all-cause mortality among older Australians in residential aged care facilities.

Materials and methods

Study participants

The study sample was taken from the Fracture Risk Epidemiology in the Elderly (FREE) study, which aimed to evaluate falls and fracture risk in very frail older people in residential care facilities. The FREE study comprised 2005 participants (473 males and 1532 females), aged 65–104 years (85.7 (SD 7.1)) from thirty hostels (intermediate care facilities) and fifty-two nursing homes in northern Sydney, Australia. Full details of the study protocol have been published elsewhere^(10,11). In brief, all nursing care facilities in the Northern Sydney Area Health Service region were randomly assigned to blocks of ten nursing homes and five hostels. During recruitment, the facilities were approached one block at a time to maintain randomisation. Facilities were invited to participate by the researchers, and out of the ninety-five facilities invited, eighty-two (88%) agreed to participate. Reasons for non-participation included: two rebuilding at the time; seven closed down after the randomisation; four refused without reason. The facilities not participating were similar to participating facilities. Informed consent was obtained from participants, or the person legally entitled to make decisions on behalf of the participant, with an overall consent rate of 55%. Consent rate was higher for residents able to provide consent independently (78%) compared with residents requiring proxy consent (33%). Non-participants were similar in age and sex to participants but had higher care needs. Ethical approval was obtained from the local human research ethics committee for each participating facility and informed consent obtained from the participants or next of kin.

Anthropometry

A single measurement of body weight was performed using calibrated scales to the nearest 1 kg in light clothing and no footwear. Participant's height was estimated using equipment with a fixed footplate and an adjustable, sliding end plate. The distance from the base of the heel to the anterior surface of the thigh above the condyles of the femur and slightly proximal to the patella was measured to the nearest 0.5 cm with knee flexed at 90°. Age- and sex-specific equations were used to estimate height from knee height⁽¹²⁾. BMI was calculated as

weight in kilograms divided by height squared in metres and the participants classified as BMI < 22 kg/m² (underweight) and BMI ≥ 22 kg/m² (desirable weight for height)⁽¹³⁾. Data are also presented to describe the number of participants with a BMI > 27 kg/m² for comparison with other datasets; however, it must be acknowledged that there is very little evidence for an upper limit on BMI for the oldest old, so these data should be used with caution.

Falls, fractures and mortality

Falls were ascertained from incident reports and nursing records at 6-week intervals, and were defined as events resulting in a person coming to rest unintentionally on the ground or other lower level, not as the result of a major intrinsic event or an overwhelming hazard⁽¹⁴⁾. Fracture data were collected via regular liaison with staff at the residential care facilities and confirmed by radiology reports. Deaths were noted according to documentation in the residents' nursing records.

Potential confounding variables

The following baseline variables were considered potential confounders in the first round of statistical analyses: sex (male, female); age (continuous); level of residential care (hostel or nursing home). All models where the relationship between BMI and outcome remained significant were adjusted further to include cognition (continuous, as measured by mini-mental state examination⁽¹⁵⁾), number of medications (continuous, as measured by review of resident files) and falls in the 2-year follow-up period (yes, no) for fracture and cognition, number of medications and fracture in the 2-year follow-up period (yes, no) for mortality.

Statistical methods

Data are expressed as mean or median and 95% CI according to data distribution. Difference in age and sex across level of care were evaluated using *t* tests and χ^2 analysis, respectively. Separate logistic regression analyses were performed, taking baseline BMI as the dependent variable and baseline sex (male, female), age (continuous) or level of care (hostel or nursing home) as potential predictors. Cox proportional hazards regression models using time to fall, fracture and death from the baseline interview as the endpoint was used to analyse the relationship between BMI and falls, fracture and mortality. We used 2 years after the baseline interview as the censoring date. We also calculated incidence rate ratios using negative binomial regression models to assess the association between BMI and number of falls. Time at risk was entered into the model as an offset variable. All analyses were conducted using the SPSS statistical package version 14.0 for Windows, except for the negative binomial regressions which were performed in SPSS version 15.0.1.1 (2007).

Results

Of the 2005 participants enrolled in the FREE study, estimated height and therefore estimated BMI was available for 1846 participants. There were no differences in age, sex or falls

for those 159 FREE participants screened out of subsequent analyses compared with the study sample for this report. There were, however, a greater proportion of FREE participants screened out of subsequent analyses that experienced a fracture ($P=0.035$) or died ($P=0.046$) in the subsequent 2 years. Table 1 shows characteristics of the study participants. Regardless of setting, males were heavier than females ($P<0.001$) and for those in hostel accommodation they also had a higher BMI than females ($P=0.023$). Females were older ($P<0.001$) and had a greater proportion that suffered a fracture in the subsequent 2 years ($P=0.017$ for hostel comparison, $P=0.009$ for nursing home comparison) compared with males across both settings. There was no sex difference for the proportion that suffered a fall in the subsequent 2 years ($P=0.257$ for hostel comparison, $P=0.780$ for nursing home comparison). The most frequently documented cause of death were: cardiac (33%); infection (29%); cerebrovascular (16%).

Older participants (OR = 1.03, 95% CI 1.02, 1.05, $P<0.001$) and those from nursing home (OR = 1.93, 95% CI 1.60, 2.34) were more likely to have a low BMI according to the logistic regression analyses. Sex was not found to independently predict BMI (OR = 1.25, 95% CI 0.99, 1.58).

In the 2-year follow-up, 1246 participants experienced a fall, 321 suffered a fracture and there were a total of 616 deaths. The incidence rate of falls was 2.1 and 1.8 falls per person-year for years 1 and 2, respectively. The incidence rate of fracture was 0.12 and 0.14 fractures per person-year for years 1 and 2, respectively. A low baseline BMI ($<22\text{ kg/m}^2$) was found to predict time to first fall (hazard ratio (HR) = 1.12, 95% CI 1.02, 1.28) and time to first fracture (HR = 1.38, 95% CI 1.10, 1.72) during the 2-year follow-up. After adjusting for age, sex and level of care, the association only remained for time to first fracture (HR = 1.38, 95% CI 1.11, 1.73). Further inclusion of the number of medications, cognition and fall in subsequent 2 years into the model only slightly reduced the increased risk of fracture at 2 years for those participants with a low BMI (HR = 1.37, 95% CI 1.09, 1.71). Figure 1 shows the cumulative survival curve over the 2 years according to BMI $<22\text{ kg/m}^2$ and $\geq 22\text{ kg/m}^2$ after adjusting for age, sex and level of care. A low BMI showed an increased risk of death at 2 years, independent of age, sex and level of care (HR = 1.52, 95% CI 1.30, 1.79). Further inclusion of number of medications, cognition and fracture in subsequent 2 years into the model only slightly reduced the increased risk of death at 2 years for those participants with a low BMI (HR = 1.50, 95% CI 1.28, 1.77). The negative binomial regression demonstrated an 11% increased risk of experiencing a greater number of falls in those with a BMI $<22\text{ kg/m}^2$ (incidence rate ratio = 1.114, $P=0.048$); however, when age, sex and level of care were entered into the model, this relationship did not persist (incidence rate ratio = 1.087, $P=0.136$).

Discussion

The findings of the present study provide clear evidence that a relationship exists between poor nutritional status, as measured by BMI, and subsequent fracture and mortality among the oldest old. At 2-year follow-up, residents of aged care facilities with a low BMI were 38% more likely to

Table 1. Participant characteristics according to level of residential care and stratified by sex (Mean values and 95% confidence intervals)

Characteristic	Overall						Hostel			Nursing home			P*
	Male (n 432)		Female (n 1414)		Male (n 256)		Female (n 819)		Male (n 176)		Female (n 595)		
	Mean	95% CI	Mean	95% CI	Mean	95% CI	Mean	95% CI	Mean	95% CI	Mean	95% CI	
Age (years)	82.7	82.0, 83.4	86.5	86.2, 86.8	82.3	81.4, 83.2	86.5	86.1, 87.0	83.3	82.1, 84.5	86.5	85.9, 87.0	0.472
Weight (kg)	70.8	69.5, 72.0	56.7	56.1, 57.4	73.0	71.4, 74.7	58.2	57.3, 59.1	67.4	65.7, 69.2	54.6	53.7, 55.6	<0.001
BMI (kg/m ²)	24.1	23.7, 24.5	23.5	23.3, 23.8	24.9	24.4, 25.4	24.1	23.8, 24.5	23.0	22.4, 23.5	22.7	22.3, 23.1	<0.001
BMI < 22 kg/m ²	145		581		72		280		73		301		<0.001
n	33.6		41.1		28.1		34.2		41.5		50.6		<0.001
BMI > 27 kg/m ²	94		305		69		206		25		99		<0.001
n	21.8		21.6		27.0		25.2		14.2		16.6		<0.001
Faller	282		964		162		550		120		414		0.174
n	65.3		68.2		63.3		67.2		68.2		69.6		<0.001
Days to first fall	121.5	106, 160	140	120, 155	177	132, 210	175	156, 194	92.5	61, 112	98.5	82, 112	<0.001
Fracture	52		269		40		185		12		84		<0.001
n	12		19		16		23		7		14		<0.001
Days to first fracture	267.5	168, 473	334	278, 371	267.5	161, 474	353	311, 388	310.5	117, 478	242	189, 360	0.031

* Comparison across setting (hostel v. nursing home).

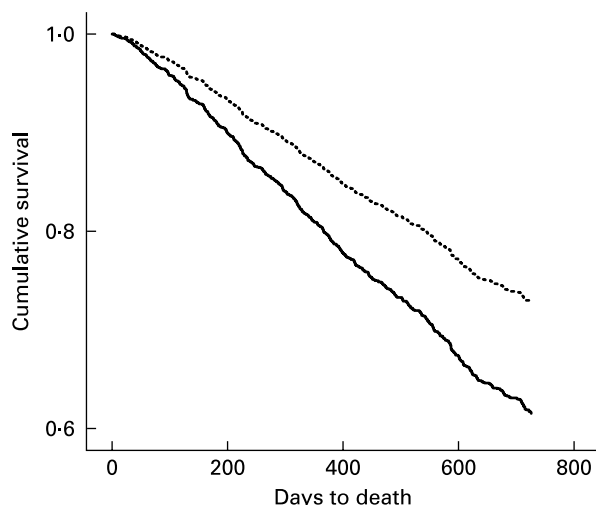


Fig. 1. BMI (—, $<22 \text{ kg/m}^2$; ···, $\geq 22 \text{ kg/m}^2$) and cumulative survival over 2 years (adjusted for age, sex and level of residential care) for the 1846 participants of the Fracture Risk Epidemiology in the Elderly study.

sustain a fracture and 52% more likely to die, even after adjusting for potential confounders, than residents with a BMI $\geq 22 \text{ kg/m}^2$.

The BMI is commonly used across a variety of settings as a measure of nutritional status and is commonly incorporated into nutrition screening protocols; however, BMI categories used to distinguish an acceptable BMI in older adults are inconsistent. Flacker & Kiely⁽¹⁶⁾ identified factors associated with 1-year mortality in nursing home residents and found that a BMI of $<23 \text{ kg/m}^2$ was associated with increased 1-year mortality (HR = 1.29, 95% CI 1.25, 1.34). Volpato *et al.*⁽¹⁷⁾ also investigated the relationship between BMI and mortality risk in nursing home residents, realising a significantly increased risk of mortality within 4 years, in residents with a low BMI ($P < 0.001$). BMI cut-offs of 22 kg/m^2 and 21.6 kg/m^2 were used to classify a low BMI in females and males, respectively. When BMI was classified into tertiles (cut-offs of 21.6 and 25.6 kg/m^2 for men and 22 and 25.4 kg/m^2 for women), a reduced risk of mortality was observed in participants in the higher BMI tertiles (relative risk 0.61 (95% CI 0.43, 0.88) for the highest *v.* the lowest BMI tertile). While the present study does not investigate the impact of high BMI on the outcomes of interest, it is consistent with Volpato *et al.*⁽¹⁷⁾ highlighting that the effect of a low BMI on risk of mortality may persist beyond 1 year.

In addition to the studies described earlier, the WHO conducted a review of a number of European studies that investigated mortality and morbidity risk according to BMI⁽⁸⁾. A Norwegian study found that a BMI between 21 and 27 kg/m^2 and 23 and 27 kg/m^2 for men and women, respectively, resulted in the lowest mortality and morbidity risk, and that the relationship between BMI and mortality is U shaped^(18,19). A Finnish study⁽²⁰⁾ found the most favourable BMI to be 27 and 31 kg/m^2 , and that being overweight did not reduce life expectancy in women aged 65–79 years, a conclusion that supported the results of another study in the same country⁽²¹⁾. In Australia, the nutrition screening initiative⁽¹³⁾ suggested that a BMI range of 22 – 27 kg/m^2 be recommended for the management of chronic disease in older adults.

One of the more common chronic conditions in later life is osteoporosis, with fall-related fractures being the most devastating consequence both in terms of loss of independence and quality of life. In the present study, we found that adjusted BMI predicted the time taken to first fracture within a sample of older adults in residential care. These findings are consistent with the previous work that has evaluated the relationship between nutritional status and risk of fracture among older adults in residential care and independent-living older adults. Sambrook *et al.*⁽¹¹⁾ investigated the influence of body weight on fracture risk in residential aged care residents and found that those who experienced a fracture were significantly lower in weight than those who did not experience a fracture over a median follow-up period of 705 d ($P < 0.001$). Lower body weight (lowest tertile: 27–52 kg) was an independent risk factor for fracture with an incidence rate ratio of 1.99 (95% CI 1.49, 2.66) compared with the highest tertile ($\geq 65 \text{ kg}$). While body weight is a commonly used measure of nutritional status, used in isolation, it can be problematic. Adjustment for height allowing for a standardised index of adiposity or monitoring of weight change are far more valuable alternatives.

Two large prospective studies in Norway also investigated the relationship between BMI and fracture. Meyer *et al.*⁽²²⁾ studied 674 000 Norwegian men and women aged 50–89 years for 16 years with the incidence of hip fracture as an outcome. An inverse relationship between BMI and the incidence of hip fracture was observed. A Cox proportional regression analysis showed a reduced risk of fracture in both men and women in the three highest quartiles of BMI compared with the lowest quartile, with a higher reduction of risk observed in the older age groups. In the 70–79 years age group, the relative risk was 0.57 (95% CI 0.5, 0.65) in women and 0.48 (95% CI 0.39, 0.59) in men for the highest quartile of BMI ($>29.9 \text{ kg/m}^2$ for women and $>27.5 \text{ kg/m}^2$ for men) compared with the lowest quartile ($<24.1 \text{ kg/m}^2$ in women and $<22.9 \text{ kg/m}^2$ in men). The inverse relationship was still present two decades after screening. In the present study, a cut-off of $\leq 22 \text{ kg/m}^2$ was used, lower than that used in the study by Meyer *et al.*⁽²²⁾, and also that used by Flacker & Keely⁽¹⁶⁾, which may indicate that a lower cut-off should be used in our population.

The present study did not provide evidence for a relationship between BMI and falls after adjustments were made for potential confounders. This finding remained robust with negative binomial regression indicating that there was also no difference in the frequency of falls between the two BMI categories. While falls does not appear to be associated with BMI, fracture does, indicating that the increased risk may be due to a reduction in ability to absorb the impact of a fall and increased bone fragility as a consequence of starvation^(23,24). The inability of the present study to demonstrate a relationship between nutritional status and falls could be related to the measure used to define nutritional status or the robust analytical approach used to test the relationship.

While the present study has demonstrated good evidence for the clinical utility of BMI, the feasibility of BMI has been a topic for debate over many years. Much of this debate is centred around pragmatic issues including inability or unwillingness of staff to perform the anthropometric measures necessary and/or to calculate BMI. There is also discussion

around the reliability of measures used to calculate BMI, particularly height. Of much concern is the issue of inability to weigh older adults with mobility limitations. Despite all of these factors being legitimate concerns, there has been much time spent on developing alternative strategies to increase feasibility including the development of knee height equations for estimating stature, easy-to-use nomograms for calculating BMI and advances in technology such as weigh chairs and beds. Increasing recognition of the clinical utility of BMI as a measure of nutritional status in addition to provision of regular training and resources will assist in improving the perception that BMI is not a feasible measure. Agreement on appropriate cut-offs for BMI should also be a priority – the present study would support a BMI ≥ 22 kg/m² being desirable.

While the evidence presented in the present study is convincing, the findings must be interpreted with caution as the study was a secondary analysis of a large epidemiological study with different aims to the present study. The primary concern is that due to the difference in aims, not all potential confounders required for the present study were necessarily measured and hence were unable to be adjusted for in the statistical analyses. Further adjustments are likely to reduce the magnitude of effect found in the present study; however, even after adjusting for some risk factors with large independent effects, the magnitude of effect for BMI and fracture and BMI and mortality was negligible. The present study did, however, include measurement of clearly defined and monitored outcome variables and had a large sample size, therefore increasing the power to detect the differences of interest and allowing for appropriate regression models to be calculated with adjustment for some established confounders.

In the context of an ageing population and the increasing incidence of falls and fractures, it is vital that screening methods are utilised to identify individuals who are at increased risk of fractures and mortality so that strategies can be implemented to reduce these risks. It is well documented that nutritional status has a role in the risk of falls and subsequent fractures, and also that poor nutritional status is a predictor of early mortality in older adults. After adjustment for a small selection of important confounders, the present study found that a relatively simple index of nutritional status, the BMI, can predict early mortality and fractures within older adults in residential care. Despite not all possible confounders being adjusted for, it is recommended that residential care staff be aware of how to measure BMI and that BMI be included as a component of admission procedures for residential care facilities as it is a rapid, inexpensive and non-invasive screen with the potential to identify high-risk individuals who may benefit from strategies to improve their nutritional status to subsequently reduce their fracture and mortality risk.

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References

1. United Nations, Department of Economic and Social Affairs, (2007) *World Population Ageing 2007*. cat. no. 07.XIII.5. New York: United Nations, Department of Economic and Social Affairs.
2. Van Weel C & Michels J (1997) Dying, not old age, to blame for costs of health care. *Lancet* **350**, 1159–1160.
3. Australian Institute of Health and Welfare (AIHW) (2004) *Australia's Health 2004*, AIHW cat. no. AUS-44. Canberra: AIHW.
4. Marottoli RA, Berkman LF & Cooney LM Jr (1992) Decline in physical function following hip fracture. *J Am Geriatr Soc* **40**, 861–866.
5. Magaziner J, Simonsick EM, Kashner Tm, *et al.* (1990) Predictors of functional recovery one year following hospital discharge for hip fracture: a prospective study. *J Gerontol* **45**, M101–M107.
6. Huang Z, Himes JH & McGovern PG (1996) Nutrition and subsequent hip fracture risk among a national cohort of white women. *Am J Epidemiol* **144**, 124–134.
7. Green SM & Watson R (2006) Nutritional screening and assessment tools for older adults: literature review. *J Adv Nurs* **54**, 477–490.
8. World Health Organization (1995) *Physical Status: the Use and Interpretation of Anthropometry. Report of a WHO Expert Committee. Technical Report Series no. 854*. Geneva: WHO.
9. Cook Z, Kirk S, Lawrenson S, *et al.* (2005) Use of BMI in the assessment of undernutrition in older subjects: reflecting on practice. *Proc Nutr Soc* **64**, 313–317.
10. Lord SR, March LM, Cameron ID, *et al.* (2003) Differing risk factors for falls in nursing home and intermediate-care residents who can and cannot stand. *J Am Geriatr Soc* **51**, 1645–1650.
11. Sambrook PN, Cameron ID, Chen JS, *et al.* (2007) Influence of fall related factors and bone strength on fracture risk in the frail elderly. *Osteoporos Int* **18**, 603–610.
12. Chumlea WC & Guo S (1992) Equations for predicting stature in white and black elderly individuals. *J Gerontol* **47**, M197–M203.
13. Lipski P (1996) Australian nutrition screening initiative. *Aust J Aging* **15**, 14–17.
14. Kellogg International Work Group (1987) The prevention of falls in later life. A report of the Kellogg International Work Group on the prevention of falls in the elderly. *Dan Med Bull* **34**, 1–24.
15. Folstein MF, Folstein SE & McHugh BR (1975) Mini-mental state a practical method for grading the cognitive state of patients for the clinician. *J Psychiatr Res* **12**, 189–198.
16. Flacker JM & Kiely DE (2003) Mortality-related factors and 1-year survival in nursing home residents. *J Am Geriatr Soc* **51**, 213–221.

17. Volpato S, Romagnoni F, Soattin L, *et al.* (2004) Body mass index, body cell mass, and 4-year all-cause mortality risk in older nursing home residents. *J Am Geriatr Soc* **52**, 886–891.
18. Waller HT (1984) Height, weight and mortality. The Norwegian experience. *Acta Med Scand Suppl* **679**, 1–56.
19. Waller HT (1988) Hazard of obesity – the Norwegian experience. *Acta Med Scand Suppl* **723**, 17–21.
20. Rissanen A, Knekt P, Heliovaara M, *et al.* (1991) Weight and mortality in Finnish women. *J Clin Epidemiol* **44**, 787–795.
21. Matilla K, Haavisto M & Rajala S (1986) Body weight as a risk factor in the elderly. *Br Med J* **292**, 867–868.
22. Meyer HE, Tverdal A & Falch JA (1995) Body height, body mass index, and fatal hip fractures: 16 years' follow-up of 674,000 Norwegian women and men. *Epidemiology* **6**, 299–305.
23. Coin A, Perissinotto E, Enzi G, *et al.* (2008) Predictors of low bone mineral density in the elderly: the role of dietary intake, nutritional status and sarcopenia. *Eur J Clin Nutr* **62**, 802–809.
24. Coin A, Sergi G, Beninca P, *et al.* (2000) Bone mineral density and body composition in underweight and normal elderly subjects. *Osteoporos Int* **11**, 1043–1050.