

On probabilities of large deviations

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Let $\{X_n\}$ be a sequence of independent identically distributed random variables and let $S_n = \sum_{k=1}^n X_k$. The rate of convergence of probabilities $P\{|S_n| > (n \log n)^{1/r}\}$, where $2 > r > 1$, is studied.

1. Introduction

Let $\{X_n : n \geq 1\}$ denote a sequence of independent identically distributed random variables with common distribution function F . Write $S_n = \sum_{k=1}^n X_k$. If F belongs to the domain of normal attraction of a stable law $V(x)$ with characteristic exponent α ($1 < \alpha < 2$) then for some $a > 0$ and some A_n

$$\lim_{n \rightarrow \infty} P\left\{a^{-1}n^{-1/\alpha}S_n - A_n \leq x\right\} = V(x).$$

(See, for example [3, p. 181].) If, moreover, $EX_1 = 0$ then the constants A_n may be taken to be zero and it follows that $a^{-1}n^{-1/\alpha}(\log n)^{-1/\alpha}S_n \rightarrow 0$ in probability. Clearly

$$P\{|S_n| > \varepsilon(n \log n)^{1/\alpha}\} \geq P\{|S_n| > \varepsilon(n \log n)^{1/r}\}$$

for $0 < r < \alpha$ and it follows that $(n \log n)^{-1/r}S_n \rightarrow 0$ in probability.

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The probability $P\{|S_n| > \epsilon(n \log n)^{1/r}\}$ or either of its one sided components is called a probability of large deviation [7]. If X has a finite variance the probability $P\{|S_n| > \epsilon(n \log n)^{1/2}\}$ is called a probability of moderate deviation (see [1], [5], [6]). We remark that we do not assume that $EX_2^1 < \infty$ but only that $E|X_1|^r < \infty$ for $1 < r < 2$.

In what follows we will assume that $\{X_n\}$ is a sequence of independent identically distributed random variables with common distribution function F . A median for the random variable X is denoted by $\text{med}(X)$ and $\lg x$ is the function defined by $\lg x = \log_e x$ for $x > 1$, $= 0$ otherwise. c denotes a generic (positive) constant and $1 < r < 2$.

2. Results

Lemmas 1 and 2 are stated here for completeness. For proofs we refer to [1].

LEMMA 1. For $r \geq 1$, $E|X|^r < \infty$ if and only

$$\sum n^{r-1} (\log n)^r P\{|X| > n \log n\} < \infty .$$

LEMMA 2. Let $\{A_n\}$ be a sequence of independent events. If

$\sum PA_n < \infty$ then

$$P\{\cup_n A_n\} \geq \sum_n PA_n - \sum_n PA_n \sum_{j=n+1}^{\infty} PA_j .$$

THEOREM. For $1 < r < 2$ the following statements are equivalent:

- (a) $EX_1 = 0$, $E|X_1|^r < \infty$;
- (b) $\sum n^{-1} \log n P\{|S_n| > \epsilon(n \log n)^{1/r}\} < \infty$ for all $\epsilon > 0$;
- (c) $\sum n^{-1} \log n P\{\max_{1 \leq k \leq n} |S_k| > \epsilon(n \lg n)^{1/r}\} < \infty$ for all $\epsilon > 0$;

$$(d) \sum n^{-1} P \left\{ \sup_{k \geq n} \left| S_k / (k \lg k)^{1/r} \right| > \varepsilon \right\} < \infty \text{ for all } \varepsilon > 0 .$$

Proof. The methods of proof parallel those used in [1] and are fairly standard.

Equivalence of (a) and (b)

It is convenient to make the proofs for symmetrized random variables X_n^S , $n = 1, 2, \dots$ and then use the weak symmetrization inequalities [4, p. 245] to transfer to the required results.

Suppose that (a) holds and write $S_n^S = \sum_1^n X_k^S$. Note that $E|X_1^S|^r < \infty$. (See [4, p. 246].) Define

$$X_{kn}^S = \begin{cases} X_k^S & \text{if } |X_k^S| < \varepsilon(n \lg n)^{1/r} \\ 0 & \text{otherwise,} \end{cases} \quad 1 \leq k \leq n$$

and let $S_{nn}^S = \sum_1^n X_{kn}^S$. Since

$$P \left\{ |S_n^S| > \varepsilon(n \lg n)^{1/r} \right\} \leq n P \left\{ |X_1^S| > \varepsilon(n \lg n)^{1/r} \right\} + P \left\{ |S_{nn}^S| > \varepsilon(n \lg n)^{1/r} \right\} ,$$

and from Lemma 1, $E|X_1^S|^r < \infty$ implies $\sum \lg n P \left\{ |X_1^S| > \varepsilon(n \lg n)^{1/r} \right\} < \infty$, it just remains to show that $\sum \frac{\lg n}{n} P \left\{ |S_{nn}^S| > \varepsilon(n \lg n)^{1/r} \right\} < \infty$.

From Markov's inequality [4, p. 148] we have

$$\begin{aligned} & \sum \frac{\lg n}{n} P\left\{ |S_{nn}^s| > \varepsilon(n \lg n)^{1/r} \right\} \\ & \leq c \sum (n \lg n)^{-2/r} \lg n E\left\{ X_{1n}^s \right\}^2 \\ & \leq c \sum_n n^{-2/r} (\lg n)^{1-2/r} \sum_k (k \lg k)^{2/r} P\{(k-1) \lg(k-1) \leq |X^s|^r < k \lg k\} \\ & = c \sum_k (k \lg k)^{2/r} P\{(k-1) \lg(k-1) \leq |X^s|^r < k \lg k\} \sum_{n=k}^{\infty} n^{-2/r} (\lg n)^{1-2/r} \\ & \leq c \sum_k (k \lg k) P\{(k-1) \lg(k-1) \leq |X^s|^r < k \lg k\} \\ & < \infty . \end{aligned}$$

The last series converges because $E|X_1^s|^r < \infty$.

Now note that $E|X_1|^r < \infty$,

$$EX_1 = 0 = n^{-1/r} S_n \xrightarrow{P} 0 \Rightarrow (n \lg n)^{-1/r} S_n \xrightarrow{P} 0 = \text{med} \left(\frac{S_n}{(n \lg n)^{1/r}} \right) \rightarrow 0 .$$

It is now easy to complete the proof of (a) \Rightarrow (b) by a simple use of weak symmetrization inequalities.

Next suppose that (b) holds. By the symmetrization inequalities

$\sum n^{-1} \lg n P\left\{ S_n^s > \varepsilon(n \lg n)^{1/r} \right\} < \infty$ where S_n^s , as before, is the sum of the symmetrized random variables. We first show that $(n \lg n)^{-1/r} S_n^s \xrightarrow{P} 0$. If

not, there exists an $\varepsilon_0 > 0$ such that either

$$P\left\{ S_{n_k}^s > \varepsilon_0 (n_k \lg n_k)^{1/r} \right\} > \varepsilon_0 \text{ or } P\left\{ S_{n_k}^s < -\varepsilon_0 (n_k \lg n_k)^{1/r} \right\} > \varepsilon_0 \text{ for}$$

infinitely many k . For the sake of argument assume

$$P\left\{ S_{n_k}^s > \varepsilon_0 (n_k \lg n_k)^{1/r} \right\} > \varepsilon_0 \text{ for infinitely many } k \text{ and choose}$$

$n_{k+1} > 2n_k$. Then for each j , $n_k \leq j < 2n_k$, we have

$$(j \lg j)^{1/r} < \left[2n_k \lg(2n_k) \right]^{1/r} \leq 2^{2/r} (n_k \lg n_k)^{1/r}$$

and

$$\begin{aligned}
 P\left\{S_j^S > \epsilon_0 \frac{(j \lg j)^{1/r}}{2^{2/r}}\right\} &\geq P\left\{S_j^S > \epsilon_0 (n_k \lg n_k)^{1/r}\right\} \\
 &\geq \frac{1}{2} P\left\{S_{n_k}^S > \epsilon_0 (n_k \lg n_k)^{1/r}\right\} \\
 &\geq \frac{\epsilon_0}{2}.
 \end{aligned}$$

It follows that

$$\begin{aligned}
 \sum \frac{\lg n}{n} P\left\{S_n^S > \frac{\epsilon_0}{2^{2/r}} (n \lg n)^{1/r}\right\} &\geq \sum_k \sum_{i=n_k}^{2n_k} \frac{\lg i}{i} P\left\{S_i^S > \frac{\epsilon_0}{2^{2/r}} (i \lg i)^{1/r}\right\} \\
 &\geq \frac{\epsilon_0}{2} \sum_k \sum_{i=n_k}^{2n_k} \frac{\lg i}{i} = \infty.
 \end{aligned}$$

This contradiction shows that $(n \lg n)^{-1/r} S_n^S \rightarrow 0$ in probability. By the degenerate convergence criterion [4, p. 317] we get

$nP\{X_n^S > \epsilon(n \lg n)^{1/r}\} \rightarrow 0$. Following Erdős [2] we write

$$A_k = \left\{X_k^S > \epsilon(n \lg n)^{1/r}\right\} \text{ and } B_k = \left\{\sum_{j \neq k}^n X_j^S \geq 0\right\} \text{ and see that}$$

$$\begin{aligned}
 P\left\{S_n^S > \epsilon(n \lg n)^{1/r}\right\} &\geq P\left\{\bigcup_{k=1}^n (A_k \cap B_k)\right\} \\
 &\geq \sum_{i=1}^n PA_i [PB_i - nPA_i] \\
 &\geq nPA_i \left|\frac{1}{2} - nPA_i\right|.
 \end{aligned}$$

Thus for $\delta > 0$ and large n we have

$$P\left\{S_n^S > \epsilon(n \lg n)^{1/r}\right\} \geq \left(\frac{1}{2} - \delta\right) nPA_i.$$

It follows that

$$\infty > \sum n^{-1} \lg n P\left\{S_n^S > \epsilon(n \lg n)^{1/r}\right\} \geq c \sum \lg n P\left\{X_1^S > \epsilon(n \lg n)^{1/r}\right\}.$$

By Lemma 1, we obtain $E|X_1^S|^r < \infty$ and thus $E|X_1|^r < \infty$ by Corollary 2 [4,

p. 246].

To show that $EX_1 = 0$ we only have to note that

$$E|X_1|^r < \infty \Rightarrow n^{-1}S_n \xrightarrow{a.s.} EX_1 \text{ and } (n \lg n)^{-1/r} S_n \xrightarrow{P} 0 .$$

Equivalence of (b) and (c)

The (c) \Rightarrow (b) part is trivial and the (b) \Rightarrow (c) part follows from Lévy's inequality [4, p. 247] and the fact that $(b) \Rightarrow (n \lg n)^{-1/r} S_n \xrightarrow{P} 0$.

Equivalence of (d) and (a)

We first show that (a) and (b) \Rightarrow (d). Choose i such that $2^i \leq n < 2^{i+1}$ and again consider the symmetrized random variables X_k^S and S_n^S . We have

$$\begin{aligned} P\{S_n^S > \varepsilon(n \lg n)^{1/r}\} &\geq P\{S_n^S > \varepsilon(2^{i+1} \lg 2^{i+1})^{1/r}\} \\ &\geq \frac{1}{2} P\{S_{2^i}^S > \varepsilon(2^{i+1} \lg 2^{i+1})^{1/r}\} \\ &\geq \frac{1}{2} P\{S_{2^i}^S > \varepsilon 2^{2/r} (2^i \lg 2^i)^{1/r}\} . \end{aligned}$$

Using once again the symmetrization inequalities, we have

$$\begin{aligned} \infty > \sum_i n^{-1} \lg n P\{S_n^S > \varepsilon(n \lg n)^{1/r}\} &= \sum_i \sum_{n=2^i}^{2^{i+1}-1} \frac{\lg n}{n} P\{S_n^S > \varepsilon(n \lg n)^{1/r}\} \\ &\geq \frac{1}{4} \sum_i \lg 2^i P\{S_{2^i}^S > 2^{2/r} \varepsilon (2^i \lg 2^i)^{1/r}\} , \end{aligned}$$

so that

$$\begin{aligned}
 \sum_n n^{-1} P \left\{ \sup_{k \geq n} \frac{S_k^s}{(k \lg k)^{1/r}} > \varepsilon \right\} &\leq c \sum_i P \left\{ \sup_{k \geq 2^i} \frac{S_k^s}{(k \lg k)^{1/r}} > \varepsilon \right\} \\
 &\leq c \sum_i \sum_{j=i}^{\infty} P \left\{ \max_{2^i \geq k < 2^{j+1}} \frac{S_k^s}{(k \lg k)^{1/r}} > \varepsilon \right\} \\
 &\leq c \sum_i \sum_{j=i}^{\infty} P \left\{ S_{2^{j+1}}^s > \varepsilon (2^j \lg 2^j)^{1/r} \right\} \quad (\text{Lévy's inequality}) \\
 &= c \sum_{j=1}^{\infty} j P \left\{ S_{2^{j+1}}^s > \varepsilon (2^j \lg 2^j)^{1/r} \right\} \\
 &\leq c \sum_j \lg 2^{j+1} P \left\{ S_{2^{j+1}}^s > \varepsilon 2^{-2/r} (2^{j+1} \lg 2^{j+1})^{1/r} \right\} \\
 &\leq c \sum_n n^{-1} \lg n P \left\{ S_n^s > \varepsilon \cdot 2^{-4/r} (n \lg n)^{1/r} \right\} \\
 &< \infty .
 \end{aligned}$$

It just remains to use the weak symmetrization inequalities and the fact that $(n \lg n)^{-1/r} S_n \xrightarrow{P} 0$ to see that (a) and (b) = (d).

Next we show that (d) \Rightarrow (a).

Let $A_k = \left\{ |X_k| > \varepsilon (k \lg k)^{1/r} \right\}$. Then the events A_k are independent and satisfy

$$\bigcup_{k=m+1}^{\infty} A_k \subset \bigcup_{k=m}^{\infty} \left\{ |S_k| > \frac{\varepsilon}{2} (k \lg k)^{1/r} \right\} ,$$

so that

$$\begin{aligned}
 P \left\{ \bigcup_{m+1}^{\infty} A_k \right\} &\leq P \left\{ \bigcup_{k=m}^{\infty} \left\{ |S_k| > \frac{\varepsilon}{2} (k \lg k)^{1/r} \right\} \right\} \\
 &= P \left\{ \sup_{k \geq m} \left| \frac{S_k}{(k \lg k)^{1/r}} \right| > \frac{\varepsilon}{2} \right\} .
 \end{aligned}$$

which $\rightarrow 0$ as $m \rightarrow \infty$ because of the fact that the sequence

$$P\left\{\sup_{k \geq m} \left| \frac{S_k}{(k \lg k)^{1/r}} \right| > \frac{\varepsilon}{2} \right\}$$

is non-increasing in m and (d) holds. An

application of the Borel zero-one criterion [4, p. 228] now shows that

$$\sum_{k=1}^{\infty} P\left\{|X_1| > \varepsilon(k \lg k)^{1/r}\right\} = \sum_{k=1}^{\infty} PA_k < \infty .$$

By Lemma 2 therefore

$$\begin{aligned} \infty &> \sum_{n=2}^{\infty} n^{-1} P\left\{\sup_{k \geq n} \left| \frac{S_k}{(k \lg k)^{1/r}} \right| > \frac{\varepsilon}{2} \right\} \\ &\geq \sum_{n=2}^{\infty} n^{-1} P\left\{\bigcup_{k=0}^{\infty} \left\{|X_{n+k}| > \varepsilon((n+k) \lg(n+k))^{1/r}\right\}\right\} \\ &\geq c \sum_{k=0}^{\infty} n^{-1} P\left\{|X_{n+k}| > \varepsilon((n+k) \lg(n+k))^{1/r}\right\} (1-\eta) \\ &\geq c \sum_{k=2}^{\infty} P\left\{|X_1| > \varepsilon(k \lg k)^{1/r}\right\} \sum_{n=2}^k \frac{1}{n} . \\ &\geq c \sum_k \lg k P\left\{|X_1| > \varepsilon(k \lg k)^{1/r}\right\} . \end{aligned}$$

It follows by Lemma 1 that $E|X_1|^2 < \infty$. From (d) we see that

$$n^{-1} S_n \xrightarrow{a.s.} 0 \text{ so that we must have } EX_1 = 0 . \text{ Thus (d) } \Rightarrow \text{(a)} .$$

This completes the proof of the theorem.

References

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