




OPTIMAL RISK SHARING FOR LAMBDA VALUE-AT-RISK

ZICHAO XIA,^{***} AND

TAIZHONG HU ^{****} *University of Science and Technology of China*

Abstract

A new risk measure, the Lambda Value-at-Risk (VaR), was proposed from a theoretical point of view as a generalization of the ordinary VaR in the literature. Motivated by the recent developments in risk sharing problems for the VaR and other risk measures, we study the optimization of risk sharing for the Lambda VaR. Explicit formulas of the inf-convolution and sum-optimal allocations are obtained with respect to the left Lambda VaRs, the right Lambda VaRs, or a mixed collection of the left and right Lambda VaRs. The inf-convolution of Lambda VaRs constrained to comonotonic allocations is investigated. Explicit formula for worst-case Lambda VaRs under model uncertainty induced by likelihood ratios is also given.

Keywords: risk measure; inf-convolution; risk allocation; optimal allocation; comonotonicity; model uncertainty

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1. Introduction

Let $(\Omega, \mathcal{F}, \mathbb{P})$ be an atomless probability space, and let L^0 be the set of all random variables defined on $(\Omega, \mathcal{F}, \mathbb{P})$. Let \mathcal{X} be a convex cone of random variables in L^0 , and let L^k be the set of all random variables with finite k th moments, where $k > 0$. For any $X \in L^0$, a positive (negative) value of X represents a financial loss (profit). A risk measure is a functional $\rho : \mathcal{X} \rightarrow (-\infty, +\infty]$; see [3, 14]. In a risk sharing problem, there are m agents equipped with respective risk measures ρ_1, \dots, ρ_m . Let $X \in \mathcal{X}$ denote the total risk, which is shared by m agents. X is splitted into an allocation $(X_1, \dots, X_m) \in \mathbb{A}_m(X)$ among m agents, where $\mathbb{A}_m(X)$ is the set of all possible allocations of X , defined as

$$\mathbb{A}_m(X) = \left\{ (X_1, \dots, X_m) \in \mathcal{X}^m : \sum_{j=1}^m X_j = X \right\}.$$

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* Postal address: International Institute of Finance, School of Management, University of Science and Technology of China, Hefei, Anhui 230026, China.

** Email address: zichaox@mail.ustc.edu.cn

*** Email address: thu@ustc.edu.cn

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The inf-convolution of risk measures ρ_1, \dots, ρ_m is the mapping $\square_{i=1}^m \rho_i : \mathcal{X} \rightarrow (-\infty, \infty]$, defined as

$$\square_{i=1}^m \rho_i(X) = \inf \left\{ \sum_{i=1}^m \rho_i(X_i) : (X_1, \dots, X_m) \in \mathbb{A}_m(X) \right\}, \quad X \in \mathcal{X}.$$

An m -tuple $(X_1, \dots, X_m) \in \mathbb{A}_m(X)$ is *optimal* (also termed as *sum-optimal*) of X for (ρ_1, \dots, ρ_m) if $\square_{i=1}^m \rho_i(X) = \sum_{i=1}^m \rho_i(X_i)$. A sequence of allocations $(X_{1n}, \dots, X_{mn}) \in \mathbb{A}_m(X)$, $n \in \mathbb{N}$, is *asymptotically optimal* if $\sum_{i=1}^m \rho_i(X_{in}) \rightarrow \square_{i=1}^m \rho_i(X)$ as $n \rightarrow \infty$. An allocation $(X_1, \dots, X_m) \in \mathbb{A}_m(X)$ is *Pareto-optimal* if for any $(Y_1, \dots, Y_m) \in \mathbb{A}_m(X)$, $\rho_i(Y_i) \leq \rho_i(X_i)$ for all $i \in [m]$ implies $\rho_i(Y_i) = \rho_i(X_i)$ for all $i \in [m]$, where $[m] = \{1, \dots, m\}$. It is shown in Proposition 1 of [12] that sum-optimality is equivalent to Pareto-optimality for monetary risk measures. For non-monetary risk measures, it is easy to see that sum-optimality implies Pareto-optimality.

Liu *et al.* [23] investigated conditions under which the inf-convolution possesses the property of law invariance. For more on inf-convolution for the case of convex risk measures, see [1], [4], [13], [19 and [26], among others.

Embrechts *et al.* [12], Liu *et al.* [21], and Wang and Wei [27] studied the optimization of risk sharing for non-convex risk measures, for examples, Value-at-Risk (VaR) and Range-Value-at-Risk (RVaR). Explicit formulas of the inf-convolution and Pareto-optimal allocations were obtained with respect to the left VaRs, the right VaRs or a mixed collection of the left and right VaRs for $m \geq 2$. Formal definitions of the left and right VaRs are defined in Subsection 2.1. More precisely, for $m = 2$, Embrechts *et al.* [12] proved that

$$\text{VaR}_{\lambda_1}^- \square \text{VaR}_{\lambda_2}^-(X) = \text{VaR}_{\lambda}^-(X), \quad X \in L^0, \quad (1.1)$$

for $\lambda_1, \lambda_2 \in [0, 1]$ such that $\lambda = \lambda_1 + \lambda_2 - 1 > 0$. Liu *et al.* [21] considered the case of a mixed collection of the left and right VaRs, and proved that

$$\text{VaR}_{\lambda_1}^+ \square \text{VaR}_{\lambda_2}^+(X) = \text{VaR}_{\lambda}^+(X), \quad X \in L^0, \quad (1.2)$$

for $\lambda_1, \lambda_2 \in [0, 1)$ such that $\lambda = \lambda_1 + \lambda_2 - 1 \geq 0$, and that

$$\text{VaR}_{\lambda_1}^- \square \text{VaR}_{\lambda_2}^+(X) = \text{VaR}_{\lambda}^+(X), \quad X \in L^0, \quad (1.3)$$

for $\lambda_1 \in [0, 1], \lambda_2 \in [0, 1)$ such that $\lambda = \lambda_1 + \lambda_2 - 1 \geq 0$. More recently, Lauzier *et al.* [20] investigated the problem of sharing risk among agents with preferences modeled by a general class of comonotonic additive and law-based distortion riskmetrics that need not be either monotone or convex, and solved explicitly Pareto-optimal allocations among agents using the Gini deviation, the mean-median deviation, or the inter-quantile difference as the relevant variability measures.

The Lambda Value-at-Risk (VaR) was proposed by Frittelli *et al.* [15] as a generalization of the usual VaR. The formal definitions of the left and right Lambda VaRs are given in Section 2 (Definition 1). The Lambda VaRs are not monetary risk measures, as can be seen from Proposition 2. One naturally wonders whether an explicit formula also holds for the inf-convolution of the Lambda VaR agents. In this paper, we generalize the formulas (1.1)–(1.3) in several directions within the context of the Lambda VaRs.

The novelty of Lambda VaR is considering a function Λ , called ‘probability loss function’, which can change and adjust according to the profits and losses of a risk variable. The Lambda

VaR can discriminate different risk variables with the same VaR at level λ but with different tail behavior. The function Λ can be either increasing or decreasing in [15]. Burzoni *et al.* [6] focused on the conditions under which the Lambda VaR is robust, elicitable and consistent in the sense of [9]. Hitaj *et al.* [17] applied Lambda VaR in financial risk management as an alternative to VaR to access capital requirements, and their findings show that Lambda VaR estimates are able to capture the tail risk and react to market fluctuations significantly faster than the VaR and expected shortfall. Corbetta and Peri [7] proposed three backtesting methodologies and assessed the accuracy of Lambda VaR from different points of view. Ince *et al.* [18] presented a novel treatment of Lambda VaR on subsets of \mathbb{R}^n , and derived risk contributions of individual assets to the overall portfolio risk, measured via Lambda VaR of the portfolio composition.

The rest of this paper is organized as follows. In Section 2, we provide the formal definitions of the Lambda VaR, collect some basic properties of the Lambda VaR and derive explicit formulas for worst-case Lambda VaRs under model uncertainty induced by likelihood ratios. In Section 3, we introduce the inf-convolution of decreasing functions, and study its detailed properties. These properties will be used in Sections 4 and 5. In Section 4, we obtain explicit formulas of the inf-convolution with respect to the left Lambda VaRs, the right Lambda VaRs and a mixed collection of the left and right Lambda VaRs. Section 5 focuses on the construction of optimal allocations and asymptotically optimal allocations of inf-convolution of several Lambda VaRs. In Section 6, we consider inf-convolution of Lambda VaRs constrained to comonotonic allocations. Section 7 contains some concluding remarks. The proofs of some lemmas and propositions appearing in the previous sections are relegated to Appendices A–D.

2. Properties of Lambda VaRs

2.1. Definitions of Lambda VaRs

Let $X \in L^0$ with distribution function F_X . The (ordinary) left-VaR of X at confidence level $\alpha \in [0, 1]$ is defined as

$$\text{VaR}_\alpha^-(X) = F_X^{-1}(\alpha) = \inf\{x \in \mathbb{R} : F_X(x) \geq \alpha\} = \sup\{x \in \mathbb{R} : F_X(x) < \alpha\},$$

and the (ordinary) right-VaR of X at confidence level $\alpha \in [0, 1]$ is defined as

$$\text{VaR}_\alpha^+(X) = \inf\{x \in \mathbb{R} : F_X(x) > \alpha\} = \sup\{x \in \mathbb{R} : F_X(x) \leq \alpha\}.$$

Here and henceforth, we use the convention that $\inf \emptyset = +\infty$ and $\sup \emptyset = -\infty$. For the role of left-quantile (VaR_α^-) and right quantile (VaR_α^+) with $\alpha \in (0, 1]$ as risk measures, see the discussion in [2] and [21, Remark 5].

Next, we recall the definition of Lambda VaRs from Bellini and Peri [5], which are generalizations of ordinary VaRs.

Definition 1. Let $X \in L^0$ with distribution function F_X , and let $\Lambda : \mathbb{R} \rightarrow [0, 1]$. The Lambda VaRs of X or F_X are defined as follows:

$$\text{VaR}_\Lambda^-(X) = \inf\{x \in \mathbb{R} : F_X(x) \geq \Lambda(x)\},$$

$$\text{VaR}_\Lambda^+(X) = \inf\{x \in \mathbb{R} : F_X(x) > \Lambda(x)\},$$

and

$$\widetilde{\text{VaR}}_{\Lambda}^{-}(X) = \sup\{x \in \mathbb{R} : F_X(x) < \Lambda(x)\},$$

$$\widetilde{\text{VaR}}_{\Lambda}^{+}(X) = \sup\{x \in \mathbb{R} : F_X(x) \leq \Lambda(x)\}.$$

$\text{VaR}_{\Lambda}^{\kappa}(X)$ and $\widetilde{\text{VaR}}_{\Lambda}^{\kappa}(X)$ are also denoted by $\text{VaR}_{\Lambda}^{\kappa}(F_X)$ and $\widetilde{\text{VaR}}_{\Lambda}^{\kappa}(F_X)$, where $\kappa \in \{-, +\}$.

It is known from [5] that $\widetilde{\text{VaR}}_{\Lambda}^{-}(X) = \text{VaR}_{\Lambda}^{-}(X)$ and $\widetilde{\text{VaR}}_{\Lambda}^{+}(X) = \text{VaR}_{\Lambda}^{+}(X)$ for $X \in L^0$ when Λ is decreasing. In this paper, “increasing” and “decreasing” are used in the weak sense. Thus, Lambda VaRs reduce from four to two. In the sequel, we only consider the left and the right Lambda VaRs, VaR_{Λ}^{-} and VaR_{Λ}^{+} .

Instead of a constant confidence level λ in the definition of VaR_{λ} , the function Λ adds flexibility in modeling tail behavior of risks. Under this assumption, properties of Lambda VaRs closely resemble those of the usual VaRs. The financial interpretation of the assumption of a decreasing Λ is well illustrated by a simple two-level Lambda VaR [5, Example 2.7].

2.2. Basic properties of Lambda VaRs

We collect some basic properties of Lambda VaRs from [5]. Throughout, let $\Lambda : \mathbb{R} \rightarrow [0, 1]$ be decreasing to avoid pathological cases, and let \mathcal{M}_1 denote the set of probability measures on $(\mathbb{R}, \mathcal{B}(\mathbb{R}))$. Then

(B1) VaR_{Λ}^{-} and VaR_{Λ}^{+} are finite if and only if $\Lambda \not\equiv 0$ and $\Lambda \not\equiv 1$.

(B2) If $\Lambda_1(x) = \Lambda_2(x)$ on their common points of continuity or $\Lambda_1(x) = \Lambda_2(x)$ almost surely with respect to the Lebesgue measure, then $\text{VaR}_{\Lambda_1}^{\kappa} = \text{VaR}_{\Lambda_2}^{\kappa}$ on L^0 for $\kappa \in \{-, +\}$.

(B3) For $\kappa \in \{-, +\}$, $\text{VaR}_{\Lambda}^{\kappa}$ is quasi-concave on \mathcal{M}_1 , that is,

$$\text{VaR}_{\Lambda}^{\kappa}(\alpha F_1 + (1 - \alpha)F_2) \geq \min\{\text{VaR}_{\Lambda}^{\kappa}(F_1), \text{VaR}_{\Lambda}^{\kappa}(F_2)\}$$

for any $F_1, F_2 \in \mathcal{M}_1$ and $0 < \alpha < 1$.

(B4) For $\kappa \in \{-, +\}$, $\text{VaR}_{\Lambda}^{\kappa}(F)$ has the “convex level set” (CxLS) property. A risk measure $\rho : \mathcal{M}_1 \rightarrow \mathbb{R}$ is said to have the CxLS property if for any $F_1, F_2 \in \mathcal{M}_1$, $\alpha \in (0, 1)$ and $\gamma \in \mathbb{R}$, it holds that

$$\rho(F_1) = \rho(F_2) = \gamma \Rightarrow \rho(\alpha F_1 + (1 - \alpha)F_2) = \gamma.$$

(B5) VaR_{Λ}^{-} is weakly lower semi-continuous, i.e. if $F_n \xrightarrow{d} F$ for $F_n, F \in \mathcal{M}_1$, then

$$\liminf_{n \rightarrow \infty} \text{VaR}_{\Lambda}^{-}(F_n) \geq \text{VaR}_{\Lambda}^{-}(F).$$

VaR_{Λ}^{+} is weakly upper semi-continuous, i.e. if $F_n \xrightarrow{d} F$ for $F_n, F \in \mathcal{M}_1$, then

$$\limsup_{n \rightarrow \infty} \text{VaR}_{\Lambda}^{+}(F_n) \leq \text{VaR}_{\Lambda}^{+}(F).$$

Some further properties of the Lambda VaRs are presented in the following propositions, whose proofs are postponed to Appendix A.

Proposition 1. For any $X \in L^0$ and $\kappa \in \{-, +\}$, we have

$$\begin{aligned} \text{VaR}_\Lambda^\kappa(\lambda X) &\geq \lambda \text{VaR}_\Lambda^\kappa(X), & 0 < \lambda < 1; \\ \text{VaR}_\Lambda^\kappa(\lambda X) &\leq \lambda \text{VaR}_\Lambda^\kappa(X), & \lambda > 1. \end{aligned} \tag{2.1}$$

Consequently, $\text{VaR}_\Lambda^\kappa(\lambda X) / \lambda$ is decreasing in $\lambda \in (0, \infty)$ for any fixed $X \in L^0$.

Han *et al.* [16] in their Remark 3.1 showed that Lambda VaRs are not star-shaped but quasi-star-shaped. Proposition 1 states that the Lambda VaRs possess a “reverse star-shape” property.

Proposition 2. VaR_Λ^- or VaR_Λ^+ is translation invariant on L^0 if and only if Λ is a constant.

Proposition 3. Let $\kappa \in \{-, +\}$. If $\text{VaR}_\Lambda^\kappa$ is positively homogeneous, i.e. $\text{VaR}_\Lambda^\kappa(\lambda X) = \lambda \text{VaR}_\Lambda^\kappa(X)$ for all $X \in L^0$ and $\lambda \in (0, \infty)$, then Λ is constant on intervals $(0, \infty)$ and $(-\infty, 0)$, respectively, that is, there exist $1 \geq \alpha_1 \geq \alpha_2 \geq \alpha_3 \geq 0$ such that

$$\Lambda(x) = \alpha_1 1_{(-\infty, 0)}(x) + \alpha_2 1_{\{0\}}(x) + \alpha_3 1_{(0, \infty)}(x). \tag{2.2}$$

Next, we give three lemmas concerning properties of Lambda VaRs, which will be used in this paper. The first one will be used repeatedly in this paper. The second and the third ones give alternative representations of the Lambda VaRs in terms of the usual VaRs. Here and in the sequel, $\bar{\Lambda} = 1 - \Lambda$, and $\Lambda(x-)$ and $\Lambda(x+)$ denote the left and right limits of function Λ at point x , respectively.

Lemma 1. For $X \in L^0$ and $x \in \mathbb{R}$, we have

$$\mathbb{P}(X > x) \leq \bar{\Lambda}(x+) \iff \text{VaR}_\Lambda^-(X) \leq x, \tag{2.3}$$

$$\mathbb{P}(X > x) < \bar{\Lambda}(x+) \implies \text{VaR}_\Lambda^+(X) \leq x, \tag{2.4}$$

$$\mathbb{P}(X \geq x) > \bar{\Lambda}(x+) \implies \text{VaR}_\Lambda^-(X) \geq x, \tag{2.5}$$

$$\mathbb{P}(X \geq x) \geq \bar{\Lambda}(x-) \iff \text{VaR}_\Lambda^+(X) \geq x. \tag{2.6}$$

Lemma 2. [16, Proposition 3.1] If $\Lambda(t)$ is not constantly 0, that is, $\Lambda(-\infty) > 0$, then

$$\text{VaR}_\Lambda^-(X) = \inf_{y \in \mathbb{R}} \{ \text{VaR}_{\Lambda(y)}^-(X) \vee y \}, \quad X \in L^0.$$

Lemma 3. If $\Lambda(t)$ is not constantly 0, that is, $\Lambda(-\infty) > 0$, then

$$\text{VaR}_\Lambda^+(X) = \inf_{y \in \mathbb{R}} \{ \text{VaR}_{\Lambda(y)}^+(X) \vee y \}, \quad X \in L^0. \tag{2.7}$$

2.3. Worst-case Lambda VaR under model uncertainty

Let \mathcal{P} be the set of all probability measures that are absolutely continuous with respect to \mathbb{P} , where \mathbb{P} is a common benchmark for all agents. For any $Q \in \mathcal{P}$, let $\text{VaR}_\Lambda^{-,Q}$ and $\text{VaR}_\Lambda^{+,Q}$ be the VaR_Λ^- and VaR_Λ^+ evaluated under the probability measure Q instead of \mathbb{P} . We consider the worst-case Lambda VaR risk measures

$$\overline{\text{VaR}}_\Lambda^{-,Q} = \sup_{Q \in \mathcal{Q}} \text{VaR}_\Lambda^{-,Q} \text{ and } \overline{\text{VaR}}_\Lambda^{+,Q} = \sup_{Q \in \mathcal{Q}} \text{VaR}_\Lambda^{+,Q},$$

where \mathcal{Q} is the subset of \mathcal{P} , describing model uncertainty. We call \mathcal{Q} an *uncertainty set* of probability measures. A particular choice of \mathcal{Q} is induced by likelihood ratios, which is the following set of probability measures whose Randon-Nikodym derivatives with respect to \mathbb{P} do not exceed a constant, i.e.

$$\mathcal{P}_\beta = \left\{ Q \in \mathcal{P} : \frac{dQ}{d\mathbb{P}} \leq \frac{1}{\beta} \right\} \text{ for } \beta \in (0, 1].$$

Liu *et al.* [21] considered the special cases VaR_λ^- and VaR_λ^+ with $\Lambda \equiv \lambda \in (0, 1)$ under uncertainty set \mathcal{P}_β , and obtained that

$$\overline{\text{VaR}}_\lambda^{-, \mathcal{P}_\beta} = \text{VaR}_{1-(1-\lambda)\beta}^-, \quad \overline{\text{VaR}}_\lambda^{+, \mathcal{P}_\beta} = \text{VaR}_{1-(1-\lambda)\beta}^+.$$

Proposition 4. *Let $\Lambda : \mathbb{R} \rightarrow [0, 1]$ be decreasing. For $\beta \in (0, 1]$, define $\Lambda_\beta = 1 - \beta\bar{\Lambda}$. Then*

$$\overline{\text{VaR}}_{\Lambda}^{+, \mathcal{P}_\beta}(X) = \text{VaR}_{\Lambda_\beta}^+(X), \quad X \in L^0, \tag{2.8}$$

Furthermore, if $\Lambda > 0$, then

$$\overline{\text{VaR}}_{\Lambda}^{-, \mathcal{P}_\beta}(X) = \text{VaR}_{\Lambda_\beta}^-(X), \quad X \in L^0. \tag{2.9}$$

Proof. We give the proof for the left Lambda VaR since the proof for the right Lambda VaR is similar. First, note that for any given $X \in L^0$ and $Q \in \mathcal{Q}$, we have $Q(X > x) \leq \mathbb{P}(X > x)/\beta$ for any $x \in \mathbb{R}$ and, hence,

$$\begin{aligned} \text{VaR}_{\Lambda_\beta}^-(X) &= \inf\{x : \mathbb{P}(X > x) \leq \bar{\Lambda}_\beta(x)\} = \inf\left\{x : \frac{1}{\beta}\mathbb{P}(X > x) \leq \bar{\Lambda}(x)\right\} \\ &\geq \inf\{x : Q(X > x) \leq \bar{\Lambda}(x)\} = \text{VaR}_\Lambda^{-, Q}(X). \end{aligned}$$

Thus,

$$\overline{\text{VaR}}_{\Lambda}^{-, \mathcal{P}_\beta}(X) \leq \text{VaR}_{\Lambda_\beta}^-(X), \quad X \in L^0. \tag{2.10}$$

To prove the reverse inequality of (2.10), we choose a special $Q_0 \in \mathcal{P}_\beta$ such that $dQ_0/d\mathbb{P} = (1/\beta)1_{\{U_X > 1-\beta\}}$, where $U_X \sim U(0, 1)$ such that $X = F_X^{-1}(U_X)$, a.s. Then

$$\begin{aligned} \text{VaR}_\Lambda^{-, Q_0}(X) &= \inf\{x : Q_0(X \leq x) \geq \Lambda(x)\} \\ &= \inf\{x : \mathbb{P}(X \leq x, U_X > 1 - \beta) \geq \beta \Lambda(x)\} \\ &= \inf\{x : \mathbb{P}(1 - \beta < U_X \leq F_X(x)) \geq \beta \Lambda(x)\} \\ &= \inf\{x : \max\{F_X(x) - 1 + \beta, 0\} \geq \beta \Lambda(x)\}. \end{aligned}$$

Since $\Lambda > 0$, it follows that

$$\text{VaR}_\Lambda^{-, Q_0}(X) = \inf\{x : F_X(x) \geq 1 - \beta\bar{\Lambda}(x)\} = \text{VaR}_{\Lambda_\beta}^-(X). \tag{2.11}$$

Therefore, $\overline{\text{VaR}}_{\Lambda}^{-, \mathcal{P}_\beta}(X) \geq \text{VaR}_{\Lambda_\beta}^-(X)$ for $X \in L^0$. This proves (2.9) for left Lambda VaR.

Remark 1. Eq. (2.11) cannot be true without the assumption $\Lambda > 0$. A counterexample is as follows. Let $\Lambda(x) = 1_{(-\infty, 2]}(x)$, $X \sim U(0, 4)$ under probability measure \mathbb{P} , and set $\beta = 1/4$. Choose $Q_0 \in P_\beta$ such that $dQ_0/d\mathbb{P} = 41_{\{U_X > 3/4\}}$, that is, $X \sim U(3, 4)$ under probability measure Q_0 . Then $\text{VaR}_{\Lambda}^{-, Q_0}(X) = 2$. However, $\text{VaR}_{\Lambda_\beta}^{-}(X) = 3 > \text{VaR}_{\Lambda}^{-, Q_0}(X)$. Thus, (2.11) does not hold in this case.

Further properties of Lambda VaRs under model uncertainty induced by Wasserstein metrics can be found in Xia [29].

3. Inf-convolution of real functions

In order to study inf-convolution of Lambda VaRs, we introduce the following inf-convolution of real functions. We restrict ourselves to consider bounded and decreasing functions.

Definition 2. Let $\Lambda_i : \mathbb{R} \rightarrow \mathbb{R}$ be a bounded and decreasing function for each $i \in [m]$. The inf-convolution of $\Lambda_1, \dots, \Lambda_m$ is denoted by $\bigcirc_{i=1}^m \Lambda_i(y)$, defined as

$$\bigcirc_{i=1}^m \Lambda_i(y) := \inf_{y_1, \dots, y_m \in \mathbb{R}, \sum_{i=1}^m y_i = y} \left\{ 1 - \sum_{i=1}^m \overline{\Lambda}_i(y_i) \right\}. \tag{3.1}$$

Throughout, we denote $\Lambda^*(y) = \bigcirc_{i=1}^m \Lambda_i(y)$.

It is easy to see that $\Lambda^*(y)$ is also decreasing and that

$$\overline{\Lambda^*}(y) = \overline{\bigcirc_{i=1}^m \Lambda_i(y)} = \sup_{y_1, \dots, y_m \in \mathbb{R}, \sum_{i=1}^m y_i = y} \sum_{i=1}^m \overline{\Lambda}_i(y_i). \tag{3.2}$$

That is, $\overline{\Lambda^*}$ is the sup-convolution of $\overline{\Lambda}_1, \dots, \overline{\Lambda}_m$. The next proposition justifies the simple fact that the inf-convolution of m functions can be seen as the repeated applications of the inf-convolution of two functions. In the expression $\Lambda_1 \bigcirc \Lambda_2 \cdots \bigcirc \Lambda_m$ below, the convention is to perform the operations \bigcirc from left to right.

Proposition 5. Let $\Lambda_i : \mathbb{R} \rightarrow \mathbb{R}$ be bounded and decreasing for each $i \in [m]$. For any $y \in \mathbb{R}$, we have $\bigcirc_{i=1}^m \Lambda_i(y) = \Lambda_1 \bigcirc \Lambda_2 \cdots \bigcirc \Lambda_m(y)$, where $\Lambda_1 \bigcirc \Lambda_2(y) = \bigcirc_{i=1}^2 \Lambda_i(y)$.

Several further properties of inf-convolution of real functions are listed in the following propositions, whose proofs are presented in Appendix B. The first proposition, Proposition 6, will be used repeatedly to prove other results in this paper, which gives the expressions of Λ^* at positive infinity and negative infinity. We denote $\Lambda(+\infty) = \lim_{x \uparrow \infty} \Lambda(x)$ and $\Lambda(-\infty) = \lim_{x \downarrow -\infty} \Lambda(x)$ for any decreasing function Λ .

Proposition 6. Let $\Lambda_i : \mathbb{R} \rightarrow \mathbb{R}$ be bounded and decreasing for $i \in [m]$. Then

$$\Lambda^*(-\infty) = \min_{1 \leq i \leq m} \left(1 - \overline{\Lambda}_i(-\infty) - \sum_{j \neq i} \overline{\Lambda}_j(+\infty) \right),$$

$$\Lambda^*(+\infty) = 1 - \sum_{i=1}^m \overline{\Lambda}_i(+\infty).$$

Proposition 7. Let $\Lambda_i : \mathbb{R} \rightarrow \mathbb{R}$ be bounded, right-continuous and decreasing for $i \in [m]$. For any $y \in \mathbb{R}$, $\Lambda^*(y)$ has either one of the following properties:

- (P1) There exists $(y_1, \dots, y_m) \in \mathbb{R}^m$ such that $\sum_{i=1}^m y_i = y$ and $\overline{\Lambda^*}(y) = \sum_{i=1}^m \overline{\Lambda_i}(y_i)$.
- (P2) There exists a sequence $\{(y_{1n}, \dots, y_{mn})\}_{n \in \mathbb{N}}$ such that

$$\sum_{i=1}^m y_{in} = y \text{ and } \sum_{i=1}^m \overline{\Lambda_i}(y_{in}) \rightarrow \overline{\Lambda^*}(y) \text{ as } n \rightarrow \infty,$$

where $\{(y_{1n}, \dots, y_{mn})\}_{n \in \mathbb{N}}$ does not have a cluster point in \mathbb{R}^m . In this case, $\overline{\Lambda^*}(y) = \overline{\Lambda^*}(-\infty)$, and $\sum_{i=1}^m \overline{\Lambda_i}(y_i) < \overline{\Lambda^*}(y)$ whenever $\sum_{i=1}^m y_i = y$.

Furthermore, if $\Lambda^*(y_0)$ has property (P2), then so does $\Lambda^*(x)$ for any $x < y_0$.

The next proposition gives sufficient conditions on $\{\Lambda_i\}$ under which Λ^* is right-continuous or continuous.

Proposition 8. (Continuity.) Let $\Lambda_i : \mathbb{R} \rightarrow \mathbb{R}$ be bounded and decreasing for $i \in [m]$.

- (1) If Λ_i is continuous for some i , then so is Λ^* .
- (2) If all Λ_i are right-continuous, then so is Λ^* .

Proposition 9 gives a sufficient condition under which Property (P1) holds. The condition is that the right tail of each Λ_i is a constant. The special case of Λ^* being constant is investigated in Proposition 10.

Proposition 9. Let Λ_i be bounded, right-continuous and decreasing for $i \in [m]$. If, for each $i \in [m]$, $\Lambda_i(y_i) = \Lambda_i(+\infty)$ for some $y_i \in \mathbb{R}$, then for any $x \in \mathbb{R}$, there exists $(x_1, \dots, x_m) \in \mathbb{R}^m$ such that $\sum_{i=1}^m x_i = x$ and $\overline{\Lambda^*}(x) = \sum_{i=1}^m \overline{\Lambda_i}(x_i)$.

Proposition 10. Let Λ_i be bounded, right-continuous and decreasing for each $i \in [m]$.

- (1) Λ^* is constant if and only if at least one Λ_i is constant.
- (2) Let Λ^* be a constant function. Then $\overline{\Lambda^*}(x) > \sum_{i=1}^m \overline{\Lambda_i}(x_i)$ for any $(x_1, \dots, x_m) \in \mathbb{R}^m$ with $x = \sum_{i=1}^m x_i$ if and only if there exists Λ_{i_0} such that $\Lambda_{i_0}(y) > \Lambda_{i_0}(+\infty)$ for any $y \in \mathbb{R}$.

In view of property (B2) in Subsection 2.2, we always assume that all Λ_i are right-continuous in the next sections.

4. Inf-convolution of several Lambda VaRs

Theorem 1. Let $\Lambda_i : \mathbb{R} \rightarrow (0, 1]$ be decreasing for $i \in [m]$, and let Λ^* be defined by (3.2). If $\overline{\Lambda^*}(-\infty) > 0$, then

$$\square_{i=1}^m \text{VaR}_{\overline{\Lambda_i}}(X) \geq \text{VaR}_{\overline{\Lambda^*}}(X), \quad X \in L^0. \tag{4.1}$$

The proof of Theorem 1 requires the following lemma, which was pointed out to us by an anonymous referee.

Lemma 4. For $\lambda_i \in [0, 1]$ and $(y_1, \dots, y_m) \in \mathbb{R}^m$, we have

$$\inf_{(X_1, \dots, X_m) \in \mathbb{A}_m(X)} \left\{ \sum_{i=1}^m \text{VaR}_{\lambda_i}^-(X_i) \vee y_i \right\} = \bigcap_{i=1}^m \text{VaR}_{\lambda_i}^-(X) \vee \sum_{i=1}^m y_i, \quad X \in L^0. \tag{4.2}$$

Proof of Theorem 1. By Lemma 2, for $X \in L^0$, we have

$$\begin{aligned} \bigcap_{i=1}^m \text{VaR}_{\Lambda_i}^-(X) &= \inf_{(X_1, \dots, X_m) \in \mathbb{A}_m(X)} \sum_{i=1}^m \text{VaR}_{\Lambda_i}^-(X_i) \\ &= \inf_{(X_1, \dots, X_m) \in \mathbb{A}_m(X)} \sum_{i=1}^m \inf_{y_i \in \mathbb{R}} \left\{ \text{VaR}_{\Lambda_i(y_i)}^-(X_i) \vee y_i \right\} \\ &= \inf_{(y_1, \dots, y_m) \in \mathbb{R}^m} \inf_{(X_1, \dots, X_m) \in \mathbb{A}_m(X)} \left\{ \sum_{i=1}^m \text{VaR}_{\Lambda_i(y_i)}^-(X_i) \vee y_i \right\} \\ &= \inf_{(y_1, \dots, y_m) \in \mathbb{R}^m} \left\{ \bigcap_{i=1}^m \text{VaR}_{\Lambda_i(y_i)}^-(X) \vee \sum_{i=1}^m y_i \right\} \end{aligned} \tag{4.3}$$

$$= \inf_{(y_1, \dots, y_m) \in \mathbb{R}^m} \left\{ \text{VaR}_{1 - \sum_{i=1}^m \overline{\Lambda}_i(y_i)}^-(X) \vee \sum_{i=1}^m y_i \right\}, \tag{4.4}$$

where (4.3) follows from Lemma 4, and (4.4) follows from the fact

$$\bigcap_{i=1}^m \text{VaR}_{\Lambda_i(y_i)}^-(X) = \text{VaR}_{1 - \sum_{i=1}^m \overline{\Lambda}_i(y_i)}^-(X)$$

by Corollary 2 in [12]. Here, we use the convention that $\text{VaR}_{1 - \sum_{i=1}^m \overline{\Lambda}_i(y_i)}^-(X) = \text{VaR}_0^-(X) = -\infty$ when $1 - \sum_{i=1}^m \overline{\Lambda}_i(y_i) < 0$. Furthermore, by the definition of Λ^* , we have

$$\begin{aligned} &\inf_{(y_1, \dots, y_m) \in \mathbb{R}^m} \left\{ \text{VaR}_{1 - \sum_{i=1}^m \overline{\Lambda}_i(y_i)}^-(X) \vee \sum_{i=1}^m y_i \right\} \\ &= \inf_{y \in \mathbb{R}} \inf_{(y_1, \dots, y_m) \in \mathbb{R}^m, \sum_{i=1}^m y_i = y} \left\{ \text{VaR}_{1 - \sum_{i=1}^m \overline{\Lambda}_i(y_i)}^-(X) \vee y \right\} \\ &\geq \inf_{y \in \mathbb{R}} \left\{ \text{VaR}_{\Lambda^*(y)}^-(X) \vee y \right\} \end{aligned} \tag{4.5}$$

$$= \text{VaR}_{\Lambda^*}^-(X), \tag{4.6}$$

where (4.5) holds since $\Lambda^*(y) \leq 1 - \sum_{i=1}^m \overline{\Lambda}_i(y_i)$ for any $(y_1, \dots, y_m) \in \mathbb{R}^m$ with $\sum_{i=1}^m y_i = y$, and (4.6) follows from Lemma 2. \square

The equality in (4.1) of Theorem 1 does not hold without further assumptions, as shown by the following counterexample. We will investigate sufficient conditions on Λ^* in Theorems 4 and 5, under which the equality in (4.1) is true.

Example 1. Let $X \in L^0$, with $\mathbb{P}(X = 0) = \mathbb{P}(X = 1) = 1/2$, and let $\Lambda_1(x) \equiv 3/4$ and

$$\Lambda_2(x) = \frac{1}{4\pi} \left(\frac{\pi}{2} - \arctan(x) \right) + \frac{3}{4}.$$

By Proposition 10, it follows that $\Lambda^* \equiv 1/2$, and $\overline{\Lambda^*}(x + y) > \overline{\Lambda_1}(x) + \overline{\Lambda_2}(y)$ for all $(x, y) \in \mathbb{R}^2$. Thus, $\text{VaR}_{\Lambda^*}^-(X) = 0$. We claim that, for any $(X_1, X_2) \in \mathbb{A}_2(X)$,

$$\text{VaR}_{\Lambda_1}^-(X_1) + \text{VaR}_{\Lambda_2}^-(X_2) \geq 1 > \text{VaR}_{\Lambda^*}^-(X).$$

In fact, if this is not true, there exists $(Y_1, Y_2) \in \mathbb{A}_2(X)$ such that $y_1 + y_2 < 1$, where $y_1 = \text{VaR}_{\Lambda_1}^-(Y_1)$ and $y_2 = \text{VaR}_{\Lambda_2}^-(Y_2)$. By Lemma 1, we have $\mathbb{P}(X_1 > y_1) \leq \overline{\Lambda_1}(y_1)$ and $\mathbb{P}(X_2 > y_2) \leq \overline{\Lambda_2}(y_2)$. Thus,

$$\frac{1}{2} = \mathbb{P}(X > y_1 + y_2) \leq \sum_{i=1}^2 \mathbb{P}(X_i > y_i) \leq \sum_{i=1}^2 \overline{\Lambda_i}(y_i) < \frac{1}{2},$$

which is a contradiction.

Theorem 2. Let $\Lambda_i : \mathbb{R} \rightarrow (0, 1)$ be decreasing for $i \in [m]$, and let Λ^* be defined by (3.2). If $\Lambda^*(-\infty) > 0$, then

$$\bigsqcup_{i=1}^m \text{VaR}_{\Lambda_i}^+(X) = \text{VaR}_{\Lambda^*}^+(X), \quad X \in L^0. \tag{4.7}$$

The proof of Theorem 2 requires the following lemma, Lemma 5, whose proof is similar to that of Lemma 4 by using cash invariance of VaR and Theorem 1 in [21].

Lemma 5. Let $\lambda_i \in [0, 1]$ and $\kappa_i \in \{-, +\}$ for $i \in [m]$. For any $(y_1, \dots, y_m) \in \mathbb{R}^m$, we have

$$\inf_{(X_1, \dots, X_m) \in \mathbb{A}_m(X)} \left\{ \sum_{i=1}^m \text{VaR}_{\lambda_i}^{\kappa_i}(X_i) \vee y_i \right\} = \bigsqcup_{i=1}^m \text{VaR}_{\lambda_i}^{\kappa_i}(X) \vee \sum_{i=1}^m y_i, \quad X \in L^0. \tag{4.8}$$

Here, the κ_i and the λ_i are chosen to avoid the appearance of $\text{VaR}_0^- \square \text{VaR}_1^+$ in (4.8).

Proof of Theorem 2. The proof is similar to that of Theorem 1. By Lemma 3, for $X \in L^0$, we have

$$\begin{aligned} \bigsqcup_{i=1}^m \text{VaR}_{\Lambda_i}^+(X) &= \inf_{(X_1, \dots, X_m) \in \mathbb{A}_m(X)} \sum_{i=1}^m \text{VaR}_{\Lambda_i}^+(X_i) \\ &= \inf_{(X_1, \dots, X_m) \in \mathbb{A}_m(X)} \sum_{i=1}^m \inf_{y_i \in \mathbb{R}} \left\{ \text{VaR}_{\Lambda_i(y_i)}^+(X_i) \vee y_i \right\} \\ &= \inf_{(y_1, \dots, y_m) \in \mathbb{R}^m} \inf_{(X_1, \dots, X_m) \in \mathbb{A}_m(X)} \left\{ \sum_{i=1}^m \text{VaR}_{\Lambda_i(y_i)}^+(X_i) \vee y_i \right\} \\ &= \inf_{(y_1, \dots, y_m) \in \mathbb{R}^m} \left\{ \bigsqcup_{i=1}^m \text{VaR}_{\Lambda_i(y_i)}^+(X) \vee \sum_{i=1}^m y_i \right\} \end{aligned} \tag{4.9}$$

$$= \inf_{(y_1, \dots, y_m) \in \mathbb{R}^m} \left\{ \text{VaR}_{1 - \sum_{i=1}^m \overline{\Lambda_i}(y_i)}^+(X) \vee \sum_{i=1}^m y_i \right\}, \tag{4.10}$$

where (4.9) follows from Lemma 5, and (4.10) follows from Theorem 1 of [21], which implies that

$$\square_{i=1}^m \text{VaR}_{\Lambda_i(y_i)}^+(X) = \text{VaR}_{1-\sum_{i=1}^m \overline{\Lambda}_i(y_i)}^+(X), \quad X \in L^0, \tag{4.11}$$

since $\Lambda_i(y_i) < 1$ for $i \in [m]$. Here we use the convention that $\text{VaR}_{1-\sum_{i=1}^m \overline{\Lambda}_i(y_i)}^+(X) = -\infty$ when $1 - \sum_{i=1}^m \overline{\Lambda}_i(y_i) < 0$. Note that

$$\begin{aligned} \inf_{(y_1, \dots, y_m) \in \mathbb{R}^m} \left\{ \text{VaR}_{1-\sum_{i=1}^m \overline{\Lambda}_i(y_i)}^+(X) \vee \sum_{i=1}^m y_i \right\} \\ = \inf_{y \in \mathbb{R}} \inf_{(y_1, \dots, y_m) \in \mathbb{R}^m, \sum_{i=1}^m y_i = y} \left\{ \text{VaR}_{1-\sum_{i=1}^m \overline{\Lambda}_i(y_i)}^+(X) \vee y \right\} \\ = \inf_{y \in \mathbb{R}} \left\{ \text{VaR}_{\Lambda^*(y)}^+(X) \vee y \right\} \end{aligned} \tag{4.12}$$

$$= \text{VaR}_{\Lambda^*}^+(X), \tag{4.13}$$

where (4.13) is due to Lemma 3, and (4.12) follows since

$$\inf \left\{ \text{VaR}_{1-\sum_{i=1}^m \overline{\Lambda}_i(y_i)}^+(X) : (y_1, \dots, y_m) \in \mathbb{R}^m, \sum_{i=1}^m y_i = y \right\} = \text{VaR}_{\Lambda^*(y)}^+(X). \tag{4.14}$$

More detail is given on (4.14) as follows. Denote by LHS the left-hand side of (4.14). Obviously, $\text{LHS} \geq \text{VaR}_{\Lambda^*(y)}^+(X)$ since VaR_λ^+ is increasing in λ and $1 - \sum_{i=1}^m \overline{\Lambda}_i(y_i) \geq \Lambda^*(y)$. On the other hand, note that VaR_λ^+ is right-continuous in λ . By (3.2), there exists a sequence $\{(y_{1n}, \dots, y_{mn})\}_{n \in \mathbb{N}}$ satisfying that $y = \sum_{k=1}^m y_{kn}$ and $1 - \sum_{i=1}^m \overline{\Lambda}_i(y_{in}) \searrow \Lambda^*(y)$ as $n \rightarrow \infty$. Thus, the lower bound $\text{VaR}_{\Lambda^*(y)}^+(X)$ is attainable by LHS. Therefore, (4.14) is true. \square

It should be pointed out that (4.14) does not hold for VaR^- because VaR^- is left-continuous but not right-continuous in λ . It is the reason the equality in (4.1) cannot be expected without additional conditions. In Theorem 2, it is required that $\Lambda_i < 1$ for all $i \in [m]$. If $\Lambda_{i_0} \equiv 1$ and $\Lambda_{j_0} < 1$ for some $i_0, j_0 \in [m]$, then $\square_{i=1}^m \text{VaR}_{\Lambda_i(y_i)}^+(X) = +\infty > \text{VaR}_{1-\sum_{i=1}^m \overline{\Lambda}_i(y_i)}^+(X)$ for $X \in L^0$, violating (4.11), and hence (4.7) does not hold.

An explicit formula of the inf-convolution is also obtained in Theorem 3 for the case of a mixed collection of left and right Lambda VaRs. Its proof can be found in Appendix C.

Theorem 3. Let $\Lambda_i : \mathbb{R} \rightarrow (0, 1]$ be decreasing and $\kappa_i \in \{-, +\}$ for $i \in [m]$, and let Λ^* be defined by (3.2), with $\Lambda^*(-\infty) > 0$. If $\kappa_i = +$ for at least one i , and $\Lambda_j < 1$ whenever $\kappa_j = +$, then

$$\square_{i=1}^m \text{VaR}_{\Lambda_i}^{\kappa_i}(X) = \text{VaR}_{\Lambda^*}^+(X), \quad X \in L^0. \tag{4.15}$$

As a special consequence of Theorems 1, 2 and 3, we get the following VaR inf-convolution formulas of ordinary VaRs:

Corollary 1. (1) [12, Corollary 2] For $\lambda_1, \lambda_2 \in [0, 1]$ such that $\lambda = \lambda_1 + \lambda_2 - 1 > 0$, we have

$$\text{VaR}_{\lambda_1}^- \square \text{VaR}_{\lambda_2}^-(X) = \text{VaR}_\lambda^-(X), \quad X \in L^0.$$

(2) [21, Theorem 1] For $\lambda_1, \lambda_2 \in [0, 1)$ such that $\lambda = \lambda_1 + \lambda_2 - 1 \geq 0$, we have

$$\text{VaR}_{\lambda_1}^+ \square \text{VaR}_{\lambda_2}^+(X) = \text{VaR}_{\lambda}^+(X), \quad X \in L^0.$$

(3) [21, Theorem 1] For $\lambda_1 \in [0, 1], \lambda_2 \in [0, 1)$ such that $\lambda = \lambda_1 + \lambda_2 - 1 \geq 0$, we have

$$\text{VaR}_{\lambda_1}^- \square \text{VaR}_{\lambda_2}^+(X) = \text{VaR}_{\lambda}^+(X), \quad X \in L^0.$$

Theorems 1, 2 and 3 are established under the assumption $\Lambda^*(-\infty) > 0$. In the end of this section, we present other results of inf-convolution of Lambda VaRs under the assumption $\Lambda^*(-\infty) \leq 0$. All proofs are postponed to Appendix C. The proofs of Propositions 11, 12 and 13 are based on Lemmas 6 and 7.

Proposition 11. Let $\Lambda_i : \mathbb{R} \rightarrow [0, 1]$ be decreasing for each $i \in [m]$, and let Λ^* be defined by (3.2) with $\Lambda^*(-\infty) \leq 0$.

- (1) If $\Lambda^*(-\infty) < 0$, then $\square_{i=1}^m \text{VaR}_{\Lambda_i}^-(X) = -\infty$ for $X \in L^0$.
- (2) If $\Lambda^*(-\infty) = 0$, then $\square_{i=1}^m \text{VaR}_{\Lambda_i}^-(X) = \min\{\sup L, \text{ess-inf}(X)\}$, where

$$L := \left\{ x \in \mathbb{R} : \Lambda^*(x) = 0, \exists \{x_i\} \text{ such that } \sum_{i=1}^m x_i = x \text{ and } \sum_{i=1}^m \overline{\Lambda}_i(x_i) = 1 \right\}.$$

Proposition 12. Let $\Lambda_i : \mathbb{R} \rightarrow [0, 1)$ be decreasing for each $i \in [m]$, and let Λ^* be defined by (3.2) with $\Lambda^*(-\infty) \leq 0$.

- (1) If $\Lambda^*(-\infty) < 0$, then $\square_{i=1}^m \text{VaR}_{\Lambda_i}^+(X) = -\infty$.
- (2) If $\Lambda^*(-\infty) = 0$, then $\square_{i=1}^m \text{VaR}_{\Lambda_i}^+(X) = \min\{\sup T, \text{ess-inf}(X)\}$, where $T = \{x \in \mathbb{R} : \Lambda^*(x) = 0\}$.

Proposition 13. Let $\Lambda_i : \mathbb{R} \rightarrow [0, 1]$ be decreasing and $\kappa_i \in \{-, +\}$ for $i \in [m]$, and let Λ^* be defined by (3.2) with $\Lambda^*(-\infty) \leq 0$. Assume that $\kappa_i = +$ for at least one i , and $\Lambda_i < 1$ whenever $\kappa_i = +$.

- (1) If $\Lambda^*(-\infty) < 0$, then $\square_{i=1}^m \text{VaR}_{\Lambda_i}^{\kappa_i}(X) = -\infty$.
- (2) If $\Lambda^*(-\infty) = 0$, then $\square_{i=1}^m \text{VaR}_{\Lambda_i}^{\kappa_i}(X) = \min\{\sup T, \text{ess-inf}(X)\}$, where $T = \{x \in \mathbb{R} : \Lambda^*(x) = 0\}$.

Lemma 6. Let $\Lambda_i : \mathbb{R} \rightarrow [0, 1]$ be decreasing for each $i \in [m]$, and let $X \in L^0$.

- (1) If $X \geq x_0$, a.s., with $\Lambda^*(x_0) = 0$, and if there does exist $(x_1, \dots, x_m) \in \mathbb{R}^m$ such that $\sum_{i=1}^m x_i = x_0$ and $\sum_{i=1}^m \overline{\Lambda}_i(x_i) = 1$, then $\square_{i=1}^m \text{VaR}_{\Lambda_i}^-(X) \geq x_0$.
- (2) If $\Lambda^*(-\infty) \leq 0$, then $\square_{i=1}^m \text{VaR}_{\Lambda_i}^-(X) \leq \text{ess-inf}(X)$.

Lemma 7. Let $\Lambda_i : \mathbb{R} \rightarrow [0, 1)$ be decreasing for $i \in [m]$, and let $X \in L^0$.

- (1) If $X \geq x_0 \in \mathbb{R}$, a.s., and $\Lambda^*(x_0) = 0$, then $\square_{i=1}^m \text{VaR}_{\Lambda_i}^+(X) \geq x_0$.
- (2) If $\Lambda^*(-\infty) \leq 0$, then $\square_{i=1}^m \text{VaR}_{\Lambda_i}^+(X) \leq \text{ess-inf}(X)$.

5. Optimal risk sharing for Lambda VaRs

From the proof of Theorem 1 and Example 1, it is known that whether Λ^* satisfies property (P₁) in Proposition 7 plays an important role in establishing the equality of (4.1). In this section, we consider the case $\Lambda^*(-\infty) > 0$, and study optimal allocations of inf-convolution for several Lambda VaRs according to whether Λ^* satisfies (P₁) or (P₂) in Proposition 7.

5.1. Left Lambda VaRs

Theorem 4. *Let $\Lambda_i : \mathbb{R} \rightarrow (0, 1]$ be decreasing for $i \in [m]$ with $\Lambda^*(-\infty) > 0$. For any $X \in L^0$, denote $x_0 = \text{VaR}_{\Lambda^*}^-(X)$. If there exists $(x_1, \dots, x_m) \in \mathbb{R}^m$ such that $\sum_{i=1}^m x_i = x_0$ and $\overline{\Lambda^*}(x_0) = \sum_{i=1}^m \overline{\Lambda_i}(x_i)$, then*

$$\square \sum_{i=1}^m \text{VaR}_{\Lambda_i}^-(X) = \text{VaR}_{\Lambda^*}^-(X) \tag{5.1}$$

Moreover, there exists an optimal allocation $(X_1, \dots, X_m) \in \mathbb{A}_m(X)$ satisfying $x_i = \text{VaR}_{\Lambda_i}^-(X_i)$ for $i \in [m]$.

Proof. Note that $x_0 = \text{VaR}_{\Lambda^*}^-(X) \in \mathbb{R}$ since $\Lambda^*(-\infty) > 0$. First, assume that $\overline{\Lambda^*}(x_0) = 0$. By the definition of Λ^* , $\Lambda_i \equiv 1$ for all $i \in [m]$ and, hence, $\Lambda^* \equiv 1$. In view of Theorem 1, we conclude that (5.1) holds and $(X, 0, \dots, 0)$ is an optimal allocation of X .

Next, consider the case $\Lambda^*(x_0) < 1$. We will construct an optimal allocation of X directly. Note that $\{X < x_0\}$, $\{X = x_0\}$ and $\{X > x_0\}$ constitute a partition of Ω . Construct an allocation $\mathbf{X} \in \mathbb{A}_m(X)$ as follows: On the set $\{X < x_0\}$, define $X_k = x_k$ for $k \in [m - 1]$ and $X_m = X - \sum_{i=1}^{m-1} x_i$. On the set $\{X = x_0\}$, define $X_j = x_j$ for $j \in [m]$. On the set $\{X > x_0\}$, let $\{C_1, \dots, C_m\}$ be a partition of $\{X > x_0\}$, satisfying that

$$\mathbb{P}(C_j) = \mathbb{P}(X > x_0) \cdot \frac{\overline{\Lambda_j}(x_j)}{\sum_{i=1}^m \overline{\Lambda_i}(x_i)}, \quad j \in [m].$$

Then, define $X_j = X - x_0 + x_j$ on C_j and $X_j = x_j$ on $\{X > x\} \setminus C_j$ for $j \in [m]$. Therefore, $\mathbf{X} \in \mathbb{A}_m(X)$ has the following representation:

$$\begin{aligned} X_k &= x_k + (X - x_0) 1_{C_k}, \quad k \in [m - 1], \\ X_m &= X - \sum_{i=1}^{m-1} X_i = x_m + (X - x_0) 1_{C_m} + (X - x_0) 1_{\{X \leq x_0\}}. \end{aligned} \tag{5.2}$$

By Lemma 1 and Proposition 8, $x_0 = \text{VaR}_{\Lambda^*}^-(X)$ implies that $\mathbb{P}(X > x_0) \leq \Lambda^*(x_0)$. Also, by construction, it follows that

$$\mathbb{P}(X_j > x_j) = \mathbb{P}(C_j) = \overline{\Lambda_j}(x_j) \cdot \frac{\mathbb{P}(X > x_0)}{\Lambda^*(x_0)} \leq \overline{\Lambda_j}(x_j), \quad j \in [m], \tag{5.3}$$

implying $\text{VaR}_{\Lambda_j}^-(X_j) \leq x_j$. Thus, $\sum_{i=1}^m \text{VaR}_{\Lambda_i}^-(X_i) \leq x_0 = \text{VaR}_{\Lambda^*}^-(X)$. By Theorem 1, we conclude that (5.1) holds and \mathbf{X} is an optimal allocation of X .

Theorem 4 states that Property (P₁) in Proposition 7 is a sufficient condition for (5.1). In Theorem 5, we show that, under Property (P₂) in Proposition 7, $\text{VaR}_{\Lambda^*}^-(X) = \text{VaR}_{\Lambda^*}^+(X)$ is a necessary and sufficient condition for (5.1) to hold. We will consider the inf-convolution of left Lambda VaRs in Theorem 10 when $\text{VaR}_{\Lambda^*}^-(X) < \text{VaR}_{\Lambda^*}^+(X)$.

Theorem 5. Let $\Lambda_i : \mathbb{R} \rightarrow (0, 1]$ be decreasing for $i \in [m]$ with $\Lambda^*(-\infty) > 0$. For any $X \in L^0$, denote $x_0 = \text{VaR}_{\Lambda^*}^-(X)$. It there does not exist $(x_1, \dots, x_m) \in \mathbb{R}^m$ such that $\sum_{i=1}^m x_i = x_0$ and $\overline{\Lambda^*}(x_0) = \sum_{i=1}^m \overline{\Lambda}_i(x_i)$, then (5.1) holds if and only if $\text{VaR}_{\Lambda^*}^-(X) = \text{VaR}_{\Lambda^*}^+(X)$. Furthermore, under the condition $\text{VaR}_{\Lambda^*}^-(X) = \text{VaR}_{\Lambda^*}^+(X)$, if $\mathbb{P}(X > x_0) < \overline{\Lambda^*}(x_0)$, an optimal allocation of X exists; if $\mathbb{P}(X > x_0) = \overline{\Lambda^*}(x_0)$, no optimal allocation of X exists, while there exists a sequence of asymptotically optimal allocations.

Proof. Necessity: We prove it by contradiction. Assume on the contrary that $\text{VaR}_{\Lambda^*}^-(X) < \text{VaR}_{\Lambda^*}^+(X)$. Under this assumption, from (5.1) it follows that there exists $(y_1, \dots, y_m) \in \mathbb{R}^m$ satisfying $\sum_{i=1}^m y_i = y < \text{VaR}_{\Lambda^*}^+(X)$, where $y_j = \text{VaR}_{\Lambda_j}^-(X_j)$ for $j \in [m]$. Note that

$$\{X \geq \text{VaR}_{\Lambda^*}^+(X)\} \subset \left\{X > \sum_{i=1}^m y_i\right\} \subset \bigcup_{i=1}^m \{X_i > y_i\}.$$

By Lemma 1, we have $\mathbb{P}(X_j > y_j) \leq \overline{\Lambda}_j(y_j)$ for $j \in [m]$. Hence we obtain that

$$\mathbb{P}(X \geq \text{VaR}_{\Lambda^*}^+(X)) \leq \sum_{i=1}^m \mathbb{P}(X_i > y_i) \leq \sum_{i=1}^m \overline{\Lambda}_i(y_i). \tag{5.4}$$

By Proposition 7, we have $\Lambda^*(y) = \Lambda^*(x_0) = \Lambda^*(-\infty)$ and $\sum_{i=1}^m \overline{\Lambda}_i(y_i) < \overline{\Lambda^*}(x_0)$. On the other hand, note that $\mathbb{P}(X > x_0) \geq \overline{\Lambda^*}(x_0)$ (Otherwise, if $\mathbb{P}(X > x_0) < \overline{\Lambda^*}(x_0)$, then $\text{VaR}_{\Lambda^*}^+(X) \leq x_0$, a contradiction.) Therefore,

$$\mathbb{P}(X \geq \text{VaR}_{\Lambda^*}^+(X)) = \mathbb{P}(X > \text{VaR}_{\Lambda^*}^-(X)) = \mathbb{P}(X > x_0) \geq \overline{\Lambda^*}(x_0) > \sum_{i=1}^m \overline{\Lambda}_i(y_i),$$

which contradicts (5.4). This proves the necessity.

Sufficiency: Suppose that $\text{VaR}_{\Lambda^*}^-(X) = \text{VaR}_{\Lambda^*}^+(X)$. First, we consider the case $\mathbb{P}(X > x_0) < \overline{\Lambda^*}(x_0)$. In this case, there exists $(x_1, \dots, x_m) \in \mathbb{R}^m$ such that $\sum_{i=1}^m x_i = x_0$ and $\sum_{i=1}^m \overline{\Lambda}_i(x_i) \in (\mathbb{P}(X > x_0), \overline{\Lambda^*}(x_0))$. Let $X \in \mathbb{A}_m(X)$ be as defined by (5.2). Then, $\mathbb{P}(X_j > x_j) = \mathbb{P}(C_j) < \overline{\Lambda}_j(x_j)$ for $j \in [m]$, implying $\text{VaR}_{\Lambda_j}^+(X_j) \leq x_j$ for $j \in [m]$. Thus,

$$\square \sum_{i=1}^m \text{VaR}_{\Lambda_i}^-(X) \leq \sum_{i=1}^m \text{VaR}_{\Lambda_i}^-(X_i) \leq \sum_{i=1}^m x_i = \text{VaR}_{\Lambda^*}^-(X).$$

This, together with Theorem 1, implies our desired statement (5.1). Moreover, $\text{VaR}_{\Lambda_j}^-(X_j) = x_j$ for $k \in [m]$, and X is an optimal allocation of X .

Next, consider the case $\mathbb{P}(X > x_0) = \overline{\Lambda^*}(x_0)$. In this case we show that no optimal allocation exists, but that there exists a sequence of allocation $(X_{1n}, \dots, X_{mn}) \in \mathbb{A}_m(X)$, $n \in \mathbb{N}$, such that $\sum_{i=1}^m \text{VaR}_{\Lambda_i}^-(X_{i,n}) \rightarrow x_0$.

Assume on the contrary that there exists an optimal allocation of X , say, $X = (X_1, \dots, X_m)$. Denote $x_j := \text{VaR}_{\Lambda_j}^-(X_j)$ for $j \in [m]$. Then we have $\sum_{i=1}^m x_i = x_0$ and

$$\overline{\Lambda^*}(x_0) = \mathbb{P}(X > x_0) \leq \sum_{i=1}^m \mathbb{P}(X_i > x_i) \leq \sum_{i=1}^m \overline{\Lambda}_i(x_i). \tag{5.5}$$

However, by Proposition 7, $\sum_{i=1}^m \bar{\Lambda}_i(x_i) < \bar{\Lambda}^*(x_0)$, which contradicts (5.5). Therefore, no optimal allocation exists. In order to find a sequence of admissible allocations of X approaching the lower bound of the inf-convolution, we consider the following two cases.

Case 1: Suppose that $\mathbb{P}(X > x_0 + \epsilon) < \mathbb{P}(X > x_0)$ for any $\epsilon > 0$. Denote $\delta_n = \mathbb{P}(X > x_0 + 1/n)$. There exists a sequence $\{(x_{1n}, \dots, x_{mn})\}_{n \in \mathbb{N}}$ such that $\sum_{i=1}^m x_{in} = x_0$ and $\sum_{i=1}^m \bar{\Lambda}_i(x_{in}) \in (\delta_n, \bar{\Lambda}^*(x_0))$. In an atomless probability space, let $\{C_{1n}, \dots, C_{mn}\}$ be a partition of $\{X > x_0 + 1/n\}$, satisfying

$$\mathbb{P}(C_{jn}) = \delta_n \frac{\bar{\Lambda}_j(x_{jn})}{\sum_{i=1}^m \bar{\Lambda}_i(x_{in})}, \quad j \in [m].$$

Define

$$X_{k,n} = x_{kn} + \frac{1}{nm} + \left(X - x_0 - \frac{1}{n}\right) 1_{C_{kn}}, \quad k \in [m-1], \tag{5.6}$$

and $X_{mn} = X - \sum_{i=1}^{m-1} X_{in}$. Then,

$$\mathbb{P}\left(X_{jn} > x_{jn} + \frac{1}{mn}\right) = \mathbb{P}(C_{jn}) = \delta_n \cdot \frac{\bar{\Lambda}_j(x_{jn})}{\sum_{i=1}^m \bar{\Lambda}_i(x_{in})} < \bar{\Lambda}_j(x_{jn}),$$

implying $\text{VaR}_{\bar{\Lambda}_j}^-(X_{jn}) \leq x_{jn} + 1/(mn)$. Thus,

$$\sum_{i=1}^m \text{VaR}_{\bar{\Lambda}_i}^-(X_{in}) \leq \sum_{i=1}^m x_{in} + \frac{1}{n} = x_0 + \frac{1}{n}. \tag{5.7}$$

Case 2: Suppose that $\mathbb{P}(X > x_0 + \epsilon_0) = \mathbb{P}(X > x_0)$ for some $\epsilon_0 > 0$. In this case, from $\mathbb{P}(X > x_0) = \bar{\Lambda}^*(x_0)$, it follows that $\bar{\Lambda}^*(x_0) > \bar{\Lambda}^*(x_0 + \epsilon)$ for any $\epsilon > 0$. By Proposition 7, there exists a sequence $\{(x_{1n}, \dots, x_{mn})\}_{n \in \mathbb{N}}$ such that $\sum_{i=1}^m x_{in} = x_0 + 1/n$ and $\sum_{i=1}^m \bar{\Lambda}_i(x_{in}) = \bar{\Lambda}^*(x_0 + 1/n)$. Then,

$$\bar{\Lambda}_1\left(x_{1n} - \frac{1}{n}\right) + \sum_{i=2}^m \bar{\Lambda}_i(x_{in}) \leq \bar{\Lambda}^*(x_0).$$

In an atomless probability space, let (C_{1n}, \dots, C_{mn}) be a partition of the set $\{X > x_0\}$, satisfying

$$\mathbb{P}(C_{1n}) = \bar{\Lambda}_1(x_{1n}) - \frac{1}{2} \left[\sum_{i=1}^m \bar{\Lambda}_i(x_{in}) - \bar{\Lambda}^*(x_0) \right] = \bar{\Lambda}_1(x_{1n}) - \frac{1}{2} \left[\bar{\Lambda}^*\left(x_0 + \frac{1}{n}\right) - \bar{\Lambda}^*(x_0) \right],$$

and

$$\mathbb{P}(C_{kn}) = (\bar{\Lambda}^*(x_0) - \mathbb{P}(C_{1n})) \frac{\bar{\Lambda}_k(x_{kn})}{\sum_{i=2}^m \bar{\Lambda}_i(x_{in})}, \quad k = 2, \dots, m.$$

It is easy to see that $\mathbb{P}(C_{1n}) + \sum_{i=2}^m \bar{\Lambda}_i(x_{in}) > \bar{\Lambda}^*(x_0)$, which implies that

$$\mathbb{P}(C_{1n}) \in \left(\bar{\Lambda}_1\left(x_{1n} - \frac{1}{n}\right), \bar{\Lambda}_1(x_{1n}) \right)$$

and $\mathbb{P}(C_{kn}) < \overline{\Lambda}_k(x_{kn})$ for $k \in [m]$. Construct a sequence of admissible allocations $(X_{1n}, \dots, X_{mn}) \in \mathbb{A}_m(X)$ as follows:

$$\begin{aligned} X_{1n} &= x_{1n} - \frac{1}{n} + (X - x_0) 1_{C_{1n}}, \\ X_{jn} &= x_{jn} + (X - x_0) 1_{C_{jn}}, \quad j = 2, \dots, m - 1, \\ X_{mn} &= X - \sum_{i=1}^{m-1} X_{in}. \end{aligned} \tag{5.8}$$

Note that $\mathbb{P}(X_{1n} > x_{1n} - 1/n) = \mathbb{P}(C_{1n}) < \overline{\Lambda}_1(x_{1n})$ and $\mathbb{P}(X_{kn} > x_{kn}) = \mathbb{P}(C_{kn}) < \overline{\Lambda}_k(x_{kn})$ for $k \geq 2$. Hence, $\text{VaR}_{\Lambda_i}^-(X_{in}) \leq x_{in}$ for $i \in [m]$. Therefore,

$$\sum_{i=1}^m \text{VaR}_{\Lambda_i}^-(X_{in}) \leq \sum_{i=1}^m x_{in} = x_0 + \frac{1}{n}. \tag{5.9}$$

In view of Theorem 1, we conclude our desired statement from (5.7) and (5.9).

Corollary 2 in Embrechts *et al.* [12] is a special case of Theorem 4 with $\Lambda_i \equiv \lambda_i$ for $i \in [m]$ satisfying $\Lambda^* \equiv \sum_{i=1}^m \lambda_i - m + 1 > 0$. Also, Proposition 9 gives a sufficient condition on the Λ_i under which property (P₁) of Proposition 7 holds. An immediate consequence of Theorem 4 is the following corollary.

Corollary 2. *Let $\Lambda_i : \mathbb{R} \rightarrow (0, 1)$ be decreasing for $i \in [m]$, with $\Lambda^*(-\infty) > 0$. If for any $j \in [m]$ there exists $x_j \in \mathbb{R}$ such that $\Lambda_j(x_j) = \Lambda_j(+\infty)$, then $\square_{i=1}^m \text{VaR}_{\Lambda_i}^-(X) = \text{VaR}_{\Lambda^*}^-(X)$, for which an optimal allocation exists.*

5.2. Right Lambda VaRs

Theorem 6. *Let $\Lambda_i : \mathbb{R} \rightarrow (0, 1)$ be decreasing for $i \in [m]$, with $\Lambda^*(-\infty) > 0$. For any $X \in L^0$, denote $x_0 := \text{VaR}_{\Lambda^*}^+(X)$. If there exists $(x_1, \dots, x_m) \in \mathbb{R}^m$ such that $\sum_{i=1}^m x_i = x_0$ and $\overline{\Lambda^*}(x_0) = \sum_{i=1}^m \overline{\Lambda}_i(x_i)$, then*

$$\square_{i=1}^m \text{VaR}_{\Lambda_i}^+(X) = \text{VaR}_{\Lambda^*}^+(X). \tag{5.10}$$

Furthermore,

- (1) *If $\mathbb{P}(X > x_0) < \overline{\Lambda^*}(x_0)$, then an optimal allocation exists.*
- (2) *If $\mathbb{P}(X > x_0) = \overline{\Lambda^*}(x_0)$, and $\mathbb{P}(X > x_0 + \epsilon) < \mathbb{P}(X > x_0)$ for any $\epsilon > 0$, then an optimal allocation exists.*
- (3) *Suppose that $\mathbb{P}(X > x_0) = \overline{\Lambda^*}(x_0)$ and $\mathbb{P}(X > x_0 + \epsilon_0) = \mathbb{P}(X > x_0)$ for some $\epsilon_0 > 0$ and that $\Lambda^*(x_0 + \epsilon) < \Lambda^*(x_0)$ for any $\epsilon > 0$.*
 - *If $\Lambda_j(x_j + \epsilon) < \Lambda_j(x_j)$ for any $\epsilon > 0$ and $j \in [m]$, then an optimal allocation exists.*
 - *If, for any $(y_1, \dots, y_m) \in \mathbb{R}^m$ satisfying $\sum_{i=1}^m y_i = x_0$ and $\sum_{i=1}^m \overline{\Lambda}_i(y_i) = \overline{\Lambda^*}(x_0)$, there always exists some $\tau_0 > 0$ such that $\Lambda_k(y_k) = \Lambda_k(y_k + \tau_0)$ for some $k \in [m]$, then no optimal allocation exists.*

Moreover, if an optimal allocation exists, then there exists $(X_1, \dots, X_m) \in \mathbb{A}_m(X)$ such that $\text{VaR}_{\Lambda_i}^+(X_i) = x_i$ for $i \in [m]$. If no optimal allocation exists, then there exists a sequence of allocations $(X_{1n}, \dots, X_{mn}) \in \mathbb{A}_m(X)$ such that $\text{VaR}_{\Lambda_j}^+(X_{jn}) \rightarrow x_j$ as $n \rightarrow \infty$ for $j \in [m]$, and $\sum_{i=1}^m \text{VaR}_{\Lambda_i}^+(X_{in}) \rightarrow x_0$.

Theorem 7. Let $\Lambda_i : \mathbb{R} \rightarrow (0, 1)$ be decreasing for $i \in [m]$, with $\Lambda^*(-\infty) > 0$. For any $X \in L^0$, denote $x_0 := \text{VaR}_{\Lambda^*}^+(X)$. If there does not exist $(x_1, \dots, x_m) \in \mathbb{R}^m$ such that $\sum_{i=1}^m x_i = x_0$ and $\overline{\Lambda^*}(x_0) = \sum_{i=1}^m \overline{\Lambda}_i(x_i)$, then (5.10) holds. Furthermore,

- (1) If $\mathbb{P}(X > x_0) < \overline{\Lambda^*}(x_0)$, an optimal allocation exists.
- (2) If $\mathbb{P}(X > x_0) = \overline{\Lambda^*}(x_0)$, no optimal allocation exists, while there exists a sequence of asymptotically optimal allocations.

In Theorems 6 and 7, the range of the Λ_i cannot be weakened from $(0, 1)$ to be $(0, 1]$ as shown by the following counterexample.

Example 2. Let $\Lambda_1(x) = 1_{\{x < 2\}} + (4/5) * 1_{\{x \geq 2\}}$ and $\Lambda_2(x) = (4/5) * 1_{\{x < 0\}} + (1/2) * 1_{\{x \geq 0\}}$. From (3.2), it follows that $\Lambda^*(x) = (1/2) * 1_{\{x < 2\}} + (3/10) * 1_{\{x \geq 2\}}$. Let X be a $(0, 1)$ -uniformly distributed random variable. Then $\text{VaR}_{\Lambda^*}^+(X) = x_0 = 1/2$, $\mathbb{P}(X > x_0) = \overline{\Lambda^*}(x_0)$, and $\overline{\Lambda^*}(1/2) = \overline{\Lambda}_1(1/2) + \overline{\Lambda}_2(0)$. If Theorem 6 holds, then there exist $(X_1, X_2) \in \mathbb{A}_2(X)$ and $\text{VaR}_{\Lambda_1}^+(X_1) = 1/2$. However, from the definition of $\text{VaR}_{\Lambda_1}^+$, it follows that $\text{VaR}_{\Lambda_1}^+(Y) \geq 2$ for any random variable $Y \in L^0$. This is a contradiction. Thus, Theorem 6 does not hold in this case.

5.3. Mixed Lambda VaRs

Theorem 8. Let $\Lambda_i : \mathbb{R} \rightarrow (0, 1]$ be decreasing for $i \in [m]$, with $\Lambda^*(-\infty) > 0$, and let $\kappa_i \in \{-, +\}$ for $i \in [m]$ such that $K := \{j : \kappa_j = +, j \in [m]\} \neq \emptyset$. Assume that $\Lambda_j < 1$ for $j \in K$. For any $X \in L^0$, denote $x_0 = \text{VaR}_{\Lambda^*}^+(X)$. If there exists $(x_1, \dots, x_m) \in \mathbb{R}^m$ such that $\sum_{i=1}^m x_i = x_0$ and $\overline{\Lambda^*}(x_0) = \sum_{i=1}^m \overline{\Lambda}_i(x_i)$, then

$$\square_{i=1}^m \text{VaR}_{\Lambda_i}^{\kappa_i}(X) = \text{VaR}_{\Lambda^*}^+(X). \tag{5.11}$$

Furthermore,

- (1) If $\mathbb{P}(X > x_0) < \overline{\Lambda^*}(x_0)$, then an optimal allocation exists.
- (2) If $\mathbb{P}(X > x_0) = \overline{\Lambda^*}(x_0)$, and $\mathbb{P}(X > x_0 + \epsilon) < \mathbb{P}(X > x_0)$ for any $\epsilon > 0$, then an optimal allocation exists.
- (3) Suppose that $\mathbb{P}(X > x_0) = \overline{\Lambda^*}(x_0)$ and $\mathbb{P}(X > x_0 + \epsilon_0) = \mathbb{P}(X > x_0)$ for some $\epsilon_0 > 0$, and that $\Lambda^*(x_0 + \epsilon) < \Lambda^*(x_0)$ for any $\epsilon > 0$.
 - If $\Lambda_j(x_j + \epsilon) < \Lambda_j(x_j)$ for any $\epsilon > 0$ and $j \in K$, then an optimal allocation exists.
 - If, for any $(y_1, \dots, y_m) \in \mathbb{R}^m$ satisfying $\sum_{i=1}^m y_i = x_0$ and $\sum_{i=1}^m \overline{\Lambda}_i(y_i) = \overline{\Lambda^*}(x_0)$, there always exists some $\tau_0 > 0$ such that $\Lambda_k(y_k) = \Lambda_k(y_k + \tau_0)$ for some $k \in [m]$, then no optimal allocation exists.

Moreover, if an optimal allocation exists, then there exists $(X_1, \dots, X_m) \in \mathbb{A}_m(X)$ such that $\text{VaR}_{\Lambda_i}^{\kappa_i}(X_i) = x_i$ for $i \in [m]$. If no optimal allocation exists, then there exists a sequence of

allocations $(X_{1n}, \dots, X_{mn}) \in \mathbb{A}_m(X)$ such that $\text{VaR}_{\Lambda_j}^{k_j}(X_{jn}) \rightarrow x_j$ as $n \rightarrow \infty$ for $j \in [m]$ and $\sum_{i=1}^m \text{VaR}_{\Lambda_i}^{k_i}(X_{in}) \rightarrow x_0$ as $n \rightarrow \infty$.

Theorem 9. *Let the Λ_i be the same as those in Theorem 8. For any $X \in L^0$, denote $x_0 = \text{VaR}_{\Lambda^*}^+(X)$. If there does not exist $(x_1, \dots, x_m) \in \mathbb{R}^m$ such that $\sum_{i=1}^m x_i = x_0$ and $\overline{\Lambda^*}(x_0) = \sum_{i=1}^m \overline{\Lambda_i}(x_i)$, then (5.11) holds. Furthermore,*

- (1) *If $\mathbb{P}(X > x_0) < \overline{\Lambda^*}(x_0)$, an optimal allocation exists.*
- (2) *If $\mathbb{P}(X > x_0) = \overline{\Lambda^*}(x_0)$, no optimal allocation exists, while there exists a sequence of asymptotically optimal allocations.*

In Theorem 5, we establish (5.1) under the assumption $\text{VaR}_{\Lambda^*}^-(X) = \text{VaR}_{\Lambda^*}^+(X)$ and (P_2) in Proposition 7. How about the explicit formula of $\square_{i=1}^m \text{VaR}_{\Lambda_i}^-(X)$ under the assumption $\text{VaR}_{\Lambda^*}^-(X) < \text{VaR}_{\Lambda^*}^+(X)$ and (P_2) ? By a similar argument to that in the proof of Theorems 7, we have the next result.

Theorem 10. *Let $\Lambda_i : \mathbb{R} \rightarrow (0, 1]$ be decreasing for $i \in [m]$ with $\Lambda^*(-\infty) > 0$. For any $X \in L^0$, denote $x_0 = \text{VaR}_{\Lambda^*}^-(X)$. Assume that there does not exist $(x_1, \dots, x_m) \in \mathbb{R}^m$ such that $\sum_{i=1}^m x_i = x_0$ and $\overline{\Lambda^*}(x_0) = \sum_{i=1}^m \overline{\Lambda_i}(x_i)$. If $\text{VaR}_{\Lambda^*}^-(X) < \text{VaR}_{\Lambda^*}^+(X)$, then*

$$\square_{i=1}^m \text{VaR}_{\Lambda_i}^-(X) = \text{VaR}_{\Lambda^*}^+(X). \tag{5.12}$$

Furthermore,

- (1) *If $\mathbb{P}(X > x_0) < \overline{\Lambda^*}(x_0)$, an optimal allocation exists.*
- (2) *If $\mathbb{P}(X > x_0) = \overline{\Lambda^*}(x_0)$, no optimal allocation exists, while there exists a sequence of asymptotically optimal allocations.*

6. Comonotonic inf-convolution of Lambda VaRs

In this section, we consider inf-convolution of Lambda VaRs constrained to comonotonic allocations, that is, allocations are constrained to be comonotonic. Comonotonicity, an extremal form of positive dependence, was introduced and has been widely used in economics, financial mathematics and actuarial science over the last two decades. The formal definition and its characterization can be found in Dhaene *et al.* [10, 11]. Random variables X_1, \dots, X_m are said to be comonotonic if there exist a random variable Z and increasing functions g_1, \dots, g_m such that $X_i = g_i(Z)$ almost surely for $i \in [m]$. Comonotonicity of more than two random variables is equivalent to pair-wise comonotonicity. In the sequel, when X_1, \dots, X_m are comonotonic, we denote by $X_i // \sum_{k=1}^m X_k$ for $i \in [m]$.

It is well-known that ordinary VaRs possess comonotonic additivity on L^0 , that is, the VaR of a sum of comonotonic random variables is simply the sum of the VaRs of the marginal distributions [10, Theorem 5]. However, this property is not true for Lambda VaRs. In the next proposition, we prove that Lambda VaRs possess comonotonic subadditivity on L^0_+ . The property of comonotonic subadditivity was first proposed by Song and Yan [24] and further investigated by Song and Yan [25].

Proposition 14. Let $\Lambda : \mathbb{R}_+ \rightarrow [0, 1]$ be decreasing, and let X_1 and X_2 be nonnegative comonotonic random variables. Then

$$\text{VaR}_\Lambda^-(X_1 + X_2) \leq \text{VaR}_\Lambda^-(X_1) + \text{VaR}_\Lambda^-(X_2) \tag{6.1}$$

and

$$\text{VaR}_\Lambda^+(X_1 + X_2) \leq \text{VaR}_\Lambda^+(X_1) + \text{VaR}_\Lambda^+(X_2). \tag{6.2}$$

Proof. Denote by $x = \text{VaR}_\Lambda^\kappa(X_1 + X_2)$ for $\kappa \in \{-, +\}$ and set $X = X_1 + X_2$. Without loss of generality, assume that $0 < x < \infty$. First note that $x > \text{ess-sup}(X)$ occurs only for VaR_Λ^+ and $x = \sup\{t : \Lambda(t) = 1\} > \text{ess-sup}(X)$. Thus, $\text{VaR}_\Lambda^+(X_i) = x$ for $i = 1, 2$ and, hence, $\text{VaR}_\Lambda^+(X_1) + \text{VaR}_\Lambda^+(X_2) = 2x > x = \text{VaR}_\Lambda^+(X_1 + X_2)$. That is, (6.2) holds when $x > \text{ess-sup}(X)$.

Next, assume that $x \leq \text{ess-sup}(X)$. Then there exists $\lambda \in [\Lambda(x+), \Lambda(x-)]$ and $\alpha \in [0, 1]$ such that $\text{VaR}_\lambda^\alpha(X_1 + X_2) = x$, where $\text{VaR}_\lambda^\alpha = (1 - \alpha) \text{VaR}_\lambda^- + \alpha \text{VaR}_\lambda^+$. Since X_1 and X_2 are comonotonic, it follows that $\text{VaR}_\lambda^\alpha(X_1 + X_2) = \text{VaR}_\lambda^\alpha(X_1) + \text{VaR}_\lambda^\alpha(X_2)$. Denote by $x_1 = \text{VaR}_\lambda^\alpha(X_1)$ and $x_2 = \text{VaR}_\lambda^\alpha(X_2)$. Now consider two cases.

Case 1. Consider the right Lambda VaR, and assume that $x_i < x$ for $i = 1, 2$ (otherwise, (6.2) is trivial). Then, for any $\epsilon > 0$, $\mathbb{P}(X_1 \leq x_1 - \epsilon) \leq \lambda \leq \Lambda(x-) \leq \Lambda(x_1 - \epsilon)$, implying $\text{VaR}_\Lambda^+(X_1) \geq x_1 - \epsilon$. Since ϵ is arbitrary, we have $\text{VaR}_\Lambda^+(X_1) \geq x_1$. Similarly, $\text{VaR}_\Lambda^+(X_2) \geq x_2$. So we get (6.2) when $x_i < x$ for $i = 1, 2$.

Case 2. Consider the left Lambda VaR, and assume that $x_i < x$ for $i = 1, 2$.

- (i) If $\Lambda(x-) > \lambda$, then $\mathbb{P}(X_1 \leq x_1 - \epsilon) \leq \lambda < \Lambda(x-) \leq \Lambda(x_1 - \epsilon)$ for any $\epsilon > 0$, implying $\text{VaR}_\Lambda^-(X_1) \geq x_1 - \epsilon$ and, hence, $\text{VaR}_\Lambda^-(X_1) \geq x_1$. Similarly, we have $\text{VaR}_\Lambda^-(X_2) \geq x_2$. Therefore, we conclude (6.1) when $\Lambda(x-) > \lambda$.
- (ii) If $\Lambda(x-) = \lambda$ and $\Lambda(x - \epsilon) > \Lambda(x-)$ for any $\epsilon > 0$, then $\mathbb{P}(X_1 \leq x_1 - \epsilon) \leq \lambda \leq \Lambda(x-) < \Lambda(x_1 - \epsilon)$, implying $\text{VaR}_\Lambda^-(X_1) \geq x_1$. Similarly, we have $\text{VaR}_\Lambda^-(X_2) \geq x_2$. We also obtain (6.1) in subcase (ii).
- (iii) If $\Lambda(x-) = \lambda$ and $\Lambda(x - \epsilon_0) = \Lambda(x-)$ for some $\epsilon_0 > 0$, it follows from the definition of VaR_Λ^- that $\mathbb{P}(X_1 + X_2 \leq x - \epsilon) < \lambda$ for any $\epsilon > 0$. This implies $\alpha = 0$, i.e. $x = \text{VaR}_\lambda^-(X_1 + X_2) = x_1 + x_2$. Also, since $\mathbb{P}(X_1 \leq x_1 - \epsilon) < \lambda \leq \Lambda(x_1 - \epsilon)$, we have $\text{VaR}_\Lambda^-(X_1) \geq x_1$. Similarly, $\text{VaR}_\Lambda^-(X_2) \geq x_2$. Again, we conclude (6.1) in subcase (iii).

This completes the proof of the proposition. □

Remark 2. Proposition 14 cannot be true without the assumption $X \in L_+^0$. Counterexamples are as follows.

- (1) Let $\Lambda(x) = (1/2) \cdot 1_{(-\infty, a)}(x)$, and $X = Y \sim U(a - 1, a + 1)$ with $a < 0$. Then $\text{VaR}_\Lambda^-(X + Y) = \text{VaR}_\Lambda^-(X) = \text{VaR}_\Lambda^-(Y) = a$. So we get that $\text{VaR}_\Lambda^-(X + Y) > \text{VaR}_\Lambda^-(X) + \text{VaR}_\Lambda^-(Y)$, violating (6.1).
- (2) Let $\Lambda(x) = 1_{(-\infty, c)}(x)$, and $X = Y \sim U(a, b)$ with $a < b < c < 0$. Then $\text{VaR}_\Lambda^+(X) = \text{VaR}_\Lambda^+(Y) = c$, and $\text{VaR}_\Lambda^+(X + Y) = c$. So we get that $\text{VaR}_\Lambda^+(X + Y) > \text{VaR}_\Lambda^+(X) + \text{VaR}_\Lambda^+(Y)$, violating (6.2).

In view of Remark 2, we restrict ourselves to considering nonnegative random variables in L_+^0 . For $X \in L_+^0$, we define the set of comonotonic allocations as

$$\mathbb{A}_m^{c+}(X) = \{(X_1, \dots, X_m) \in \mathbb{A}_m^+(X) : X_i \in L_+^0, X_i // X, i \in [m]\}.$$

The constrained (comonotonic) inf-convolution of risk measures ρ_1, \dots, ρ_m is defined as

$$\boxplus_{i=1}^m \rho_i(X) = \inf \left\{ \sum_{i=1}^m \rho_i(X_i) : (X_1, \dots, X_m) \in \mathbb{A}_m^{c+}(X) \right\}.$$

An m -tuple $(X_1, \dots, X_m) \in \mathbb{A}_m^{c+}$ is said to be an optimal constrained allocation of X for (ρ_1, \dots, ρ_m) if $\sum_{i=1}^m \rho_i(X_i) = \boxplus_{i=1}^m \rho_i(X)$.

Theorem 11. Let $\Lambda_i : \mathbb{R}_+ \rightarrow [0, 1]$ be decreasing for $i \in [m]$, and set $\Lambda = \min_{1 \leq i \leq m} \Lambda_i$, with $m \geq 2$. Then

$$\boxplus_{i=1}^m \text{VaR}_{\Lambda_i}^-(X) = \text{VaR}_{\Lambda}^-(X), \quad X \in L_+^0. \tag{6.3}$$

If, in addition, $\Lambda_i < 1$ for each i , then

$$\boxplus_{i=1}^m \text{VaR}_{\Lambda_i}^+(X) = \text{VaR}_{\Lambda}^+(X), \quad X \in L_+^0. \tag{6.4}$$

Proof. First, consider the case of the left Lambda VaR. Choose any $X \in \mathbb{A}_m^{c+}(X)$. Since $\Lambda \leq \Lambda_i$ for each i , and VaR_{Λ}^- is increasing in Λ , it follows that $\text{VaR}_{\Lambda}^-(X) \geq \text{VaR}_{\Lambda_i}^-(X)$ for each i . Then $\sum_{i=1}^m \text{VaR}_{\Lambda_i}^-(X_i) \geq \sum_{i=1}^m \text{VaR}_{\Lambda}^-(X_i) \geq \text{VaR}_{\Lambda}^-(X)$, where the second inequality follows from Proposition 14. Thus, we have

$$\boxplus_{i=1}^m \text{VaR}_{\Lambda_i}^-(X) \geq \text{VaR}_{\Lambda}^-(X). \tag{6.5}$$

To prove the reverse inequality of (6.5), we set $k = \text{argmin}_{1 \leq i \leq m} \text{VaR}_{\Lambda_i}^-(X)$, that is, $\inf\{x \in \mathbb{R}_+ : F_X(x) \geq \Lambda_k(x)\} \leq \inf\{x \in \mathbb{R}_+ : F_X(x) \geq \Lambda_i(x)\}$ for $i \neq k$. Then

$$\begin{aligned} \text{VaR}_{\Lambda_k}^-(X) &= \inf\{x \in \mathbb{R}_+ : F_X(x) \geq \Lambda_k(x)\} \\ &= \inf\{x \in \mathbb{R}_+ : F_X(x) \geq \Lambda_j(x) \text{ for some } j \in [m]\} \\ &= \inf \left\{ x \in \mathbb{R}_+ : F_X(x) \geq \min_{1 \leq j \leq m} \Lambda_j(x) \right\} = \text{VaR}_{\Lambda}^-(X). \end{aligned}$$

Now choose $X_k = X$ and $X_i = 0$ for all $i \neq k$. Obviously, $\text{VaR}_{\Lambda_i}^-(X_i) = 0$ for $i \geq 2$. So we have $\sum_{i=1}^m \text{VaR}_{\Lambda_i}^-(X_i) = \text{VaR}_{\Lambda_k}^-(X) = \text{VaR}_{\Lambda}^-(X)$, implying that $\boxplus_{i=1}^m \text{VaR}_{\Lambda_i}^-(X) \leq \text{VaR}_{\Lambda}^-(X)$. This proves (6.3).

The proof of the right Lambda VaR is similar by observing that $\Lambda_i < 1$ implies $\text{VaR}_{\Lambda_i}^+(0) = 0$.

Remark 3. One may wonder whether $\text{VaR}_{\Lambda_1}^- \boxplus (\text{VaR}_{\Lambda_2}^+(X)) = \text{VaR}_{\Lambda}^+(X)$ for $X \in L_+^0$, with $\Lambda = \min\{\Lambda_1, \Lambda_2\}$. However, this is not true, as shown by the following counterexample. Let $X_1 = X_2 \sim B(1, 1/2)$, $X = X_1 + X_2$, and define $\Lambda_1 = \Lambda_2 \equiv 1/2$. Then $\text{VaR}_{1/2}^+(X) = 2$, $\text{VaR}_{1/2}^-(X_1) = 0$, and $\text{VaR}_{1/2}^+(X_2) = 1$. Therefore, $\text{VaR}_{\Lambda_1}^- \boxplus \text{VaR}_{\Lambda_2}^+(X) < \text{VaR}_{\Lambda}^+(X)$.

For optimal comonotonic allocations, see Jouini *et al.* [19] for law-determined convex risk measures, and Cui *et al.* [8] for general distortion risk measures in the context of designing

optimal reinsurance contracts. Embrechts *et al.* [12] obtained explicit formulas for comonotonic inf-convolutions under distortion risk measures including RVaR and ES. Wang and Zitikis [28] provided analytical solutions to inf-convolution for VaRs under the weak comonotonicity constraints on the dependence structure of admissible allocations. Weak comonotonicity ranges from strong comonotonicity to no dependence structure. Liu *et al.* [21] considered comonotonic inf-convolution of tail risk measures.

The next proposition gives a connection between comonotonic inf-convolution and unrestricted inf-convolution of Lambda VaRs.

Proposition 15. *Let $\Lambda_i : \mathbb{R}_+ \rightarrow (0, 1]$ be decreasing for each $i \in [m]$.*

(1) *At least $m - 1$ of the Λ_i are equal to 1 if and only if*

$$\boxplus_{i=1}^m \text{VaR}_{\Lambda_i}^-(X) = \boxminus_{i=1}^m \text{VaR}_{\Lambda_i}^-(X) \text{ for all } X \in L_+^0. \tag{6.6}$$

(2) *If $\Lambda_i \equiv 1$ for some $i \in [m]$, then*

$$\boxplus_{i=1}^m \text{VaR}_{\Lambda_i}^+(X) = \boxminus_{i=1}^m \text{VaR}_{\Lambda_i}^+(X) \text{ for all } X \in L_+^0. \tag{6.7}$$

Proof. (1) The necessity is trivial. For the sufficiency, by Theorem 11, Eq. (6.6) holds if and only if $\text{VaR}_{\Lambda^*}^-(X) = \text{VaR}_{\min\{\Lambda_1, \dots, \Lambda_m\}}^-(X)$ for all $X \in L_+^0$ or, equivalently, $\Lambda^* = \min_{1 \leq i \leq m} \Lambda_i$ almost everywhere with respect to the Lebesgue measure. So we have $\overline{\Lambda^*}(+\infty) = \sum_{i=1}^m \overline{\Lambda_i}(+\infty) = \max_{1 \leq i \leq m} \overline{\Lambda_i}(+\infty)$, implying that at least $m - 1$ of the Λ_i are equal to 1.

(2) The proof follows immediately since both sides of (6.7) are infinity when $\Lambda_i \equiv 1$ for some i .

7. Concluding remarks

This paper is based on a PhD thesis [29]. In this paper, we give a thorough discussion of a risk sharing problem, in which there are m agents equipped with respective risk measures $\text{VaR}_{\Lambda_1}^{\kappa_1}, \text{VaR}_{\Lambda_2}^{\kappa_2}, \dots, \text{VaR}_{\Lambda_m}^{\kappa_m}$, where $\kappa_i \in \{-, +\}$ and $\Lambda_i : \mathbb{R} \rightarrow [0, 1]$ is decreasing and right-continuous for each i . We obtain the explicit formulas of inf-convolution and optimal allocations of a random variable in L^0 under different assumptions.

During the revision of this paper, we note that Liu [22] also studied the risk sharing problem among multiple agents using $\text{VaR}_{\Lambda_1}^-, \dots, \text{VaR}_{\Lambda_m}^-$ as their preferences when the Λ_i are all decreasing and right-continuous, or increasing and left-continuous. However, when the Λ_i are all decreasing and right-continuous, the explicit formulas of inf-convolution with respect to left Lambda VaRs and the construction of optimal allocations in two papers are different. Our approach is based on the inf-convolution of decreasing functions, introduced and investigated in Section 3. Moreover, Liu [22] investigated the inf-convolution of two risk measures: (i) VaR_{Λ}^- and one law-invariant monotone risk measure without cash-additivity; (ii) $\widetilde{\text{VaR}}_{\Lambda}^-$ and one risk measure that is consistent with the second-order stochastic dominance. In Cases (i) and (ii), no assumption on monotonicity of Λ is imposed.

Appendix A. Proofs of results in Section 2

Proof of Proposition 1. It suffices to prove (2.1) since the other inequality is equivalent to (2.1). For $\lambda \in (0, 1)$, we have

$$\begin{aligned} \text{VaR}_\Lambda^+(\lambda X) &= \inf\left\{t : F_X\left(\frac{t}{\lambda}\right) > \Lambda(t)\right\} = \lambda \inf\{x : F_X(x) > \Lambda(\lambda x)\} \\ &\geq \lambda \inf\{x : F_X(x) > \Lambda(x)\} = \lambda \text{VaR}_\Lambda^+(X), \end{aligned}$$

where the inequality follows since Λ is decreasing. The proof for VaR_Λ^- is similar and, hence, omitted. □

Proof of Proposition 2. We prove only the necessity. Assume on the contrary that Λ is not a constant. Then there exist $x_1 < x_2$ such that $\Lambda(x_1) > \Lambda(x_2)$. Now choose $x_0 < x_1$, and let $X \in L^0$ with distribution function F_X such that $F_X(x_0) = [\Lambda(x_1) + \Lambda(x_2)]/2$. Since F_X is right-continuous at point x_0 , there exists $\epsilon > 0$ such that $x_0 + \epsilon < x_1$ and $F(x_0 + \epsilon) < \Lambda(x_1) \leq \Lambda(x_0 + \epsilon)$, which implies $\text{VaR}_\Lambda^\kappa(X) \geq x_0 + \epsilon > x_0$, where $\kappa \in \{-, +\}$. Note that $\mathbb{P}(X + x_2 - x_0 \leq x_2) = F_X(x_0) > \Lambda(x_2)$. Thus, $\text{VaR}_\Lambda^\kappa(X + x_2 - x_0) \leq x_2$. Moreover, by the assumption of translation invariance of $\text{VaR}_\Lambda^\kappa$, we have $\text{VaR}_\Lambda^\kappa(X) \leq x_0$. This contradicts with $\text{VaR}_\Lambda^\kappa(X) > x_0$. □

Proof of Proposition 3. The proof consists of the following three steps:

- (1) Suppose that there exist $0 < x_1 < x_2$ such that $\Lambda(x_1) > \Lambda(x_2)$. Choose $0 < x_0 < x_1$, and let $X \in L^0$, with distribution function F_X satisfying $F_X(x_0) = [\Lambda(x_1) + \Lambda(x_2)]/2$. From the proof of Proposition 2, it follows that $\text{VaR}_\Lambda^\kappa(X) > x_0$, where $\kappa \in \{-, +\}$. On the other hand, $\mathbb{P}((x_2/x_0)X \leq x_2) = \mathbb{P}(X \leq x_0) > \Lambda(x_2)$, which implies $\text{VaR}_\Lambda^\kappa((x_2/x_0)X) \leq x_2$. Moreover, by positive homogeneity of $\text{VaR}_\Lambda^\kappa$, we obtain that $(x_2/x_0) \text{VaR}_\Lambda^\kappa(X) \leq x_2$ and, hence, $\text{VaR}_\Lambda^\kappa(X) \leq x_0$, a contradiction. This means that Λ is constant on $(0, \infty)$.
- (2) Suppose that there exist $x_1 < x_2 < 0$ such that $\Lambda(x_1) > \Lambda(x_2)$. Choose $x_0 \in (x_2, 0)$, and let $X \in L^0$ with distribution function F_X satisfying $F_X(x_0) = [\Lambda(x_1) + \Lambda(x_2)]/2$. Obviously, we have $\text{VaR}_\Lambda^\kappa(X) \leq x_0$ by the definition of Lambda VaRs. Since F_X is right-continuous at point x_0 , there exists $\epsilon > 0$ such that $x_0 + \epsilon < 0$ and $\mathbb{P}\left(\frac{x_1}{x_0 + \epsilon} X \leq x_1\right) = \mathbb{P}(X \leq x_0 + \epsilon) < \Lambda(x_1)$, which implies $\text{VaR}_\Lambda^\kappa\left(\frac{x_1}{x_0 + \epsilon} X\right) = \frac{x_1}{x_0 + \epsilon} \text{VaR}_\Lambda^\kappa(X) \geq x_1$. Thus, $\text{VaR}_\Lambda^\kappa(X) \geq x_0 + \epsilon$, leading to a contradiction. Therefore, Λ is also constant on $(-\infty, 0)$.
- (3) From the previous discussions, it follows that Λ has the representation (2.2) with $1 \geq \alpha_1 \geq \alpha_2 \geq \alpha_3 \geq 0$. If Λ is of the form (2.2), it can be checked that for any $X \in L^0$,

$$\text{VaR}_\Lambda^\kappa(X) = \begin{cases} 0, & F_X(0) \in (\alpha_3, \alpha_1), \\ \max\{0, \text{VaR}_{\alpha_3}^\kappa(X)\}, & F_X(0) \leq \alpha_3 < \alpha_1, \\ \min\{0, \text{VaR}_{\alpha_1}^\kappa(X)\}, & F_X(0) \geq \alpha_1 > \alpha_3, \\ \text{VaR}_{\alpha_1}^\kappa(X), & \alpha_1 = \alpha_3. \end{cases}$$

Thus, $\text{VaR}_\Lambda^\kappa$ is positive homogeneous on L^0 .

This completes the proof of the proposition. □

Proof of Lemma 1. We prove only (2.3) and (2.6); the proofs of (2.4) and (2.5) are similar.

Necessity of (2.3): If $\mathbb{P}(X > x) \leq \bar{\Lambda}(x+)$, then $\mathbb{P}(X \leq x) \geq \Lambda(x+)$, which implies $\mathbb{P}(X \leq x + \epsilon) \geq \Lambda(x + \epsilon)$ for any $\epsilon > 0$ since $\Lambda(t)$ is decreasing in t . Thus, $\text{VaR}_{\bar{\Lambda}}^-(X) \leq x + \epsilon$. Setting $\epsilon \rightarrow 0$, we get $\text{VaR}_{\bar{\Lambda}}^-(X) \leq x$.

Sufficiency of (2.3): If $\text{VaR}_{\bar{\Lambda}}^-(X) \leq x$, then $\mathbb{P}(X \leq x + \epsilon) \geq \Lambda(x + \epsilon)$ for any $\epsilon > 0$. Setting $\epsilon \rightarrow 0$, it follows that $\mathbb{P}(X \leq x) \geq \Lambda(x+)$.

Necessity of (2.6): If $\mathbb{P}(X \geq x) \geq \bar{\Lambda}(x-)$, then $\mathbb{P}(X < x) \leq \Lambda(x-)$. Thus, $\mathbb{P}(X \leq x - \epsilon) \leq \Lambda(x - \epsilon)$ for any $\epsilon > 0$, which implies $\text{VaR}_{\bar{\Lambda}}^+(X) \geq x - \epsilon$. Setting $\epsilon \rightarrow 0$ yields $\text{VaR}_{\bar{\Lambda}}^+(X) \geq x$.

Sufficiency of (2.6): If $\text{VaR}_{\bar{\Lambda}}^+(X) \geq x$, then $\mathbb{P}(X \leq x - \epsilon) \leq \Lambda(x - \epsilon)$ for any $\epsilon > 0$. Setting $\epsilon \rightarrow 0$, it follows that $\mathbb{P}(X < x) \leq \Lambda(x-)$, that is, $\mathbb{P}(X \geq x) \geq \bar{\Lambda}(x-)$. \square

Proof of Lemma 3. The proof is similar to that of Proposition 3.1 in Han *et al.* [16]. If $\Lambda \equiv 1$, then both sides of (2.7) are infinite and, thus, (2.7) holds trivially. Next, we consider the case $\Lambda \neq 1$. Since $\mathbb{P}(X \leq x) > \lambda$ for $x \in \mathbb{R}$ and $\lambda \in [0, 1)$ implies $\text{VaR}_{\lambda}^+(X) \leq x$, it follows that

$$\begin{aligned} \text{VaR}_{\Lambda}^+(X) &= \inf\{x \in \mathbb{R} : \mathbb{P}(X \leq x) > \Lambda(x)\} \geq \inf\{x \in \mathbb{R} : \text{VaR}_{\Lambda(x)}^+(X) \leq x\} \\ &\geq \inf\{\text{VaR}_{\Lambda(x)}^+(X) \vee x : \text{VaR}_{\Lambda(x)}^+(X) \leq x\} \geq \inf_{x \in \mathbb{R}} \left\{ \text{VaR}_{\Lambda(x)}^+(X) \vee x \right\}. \end{aligned}$$

To prove (2.7), it suffices to prove that for any $x \in \mathbb{R}$,

$$\text{VaR}_{\Lambda}^+(X) \leq \text{VaR}_{\Lambda(x)}^+(X) \vee x. \tag{A.1}$$

It is trivial for $x \geq \text{VaR}_{\Lambda}^+(X)$. If $x < \text{VaR}_{\Lambda}^+(X)$, then $\mathbb{P}(X \leq x) \leq \Lambda(x)$. Thus, for any $a \in (x, \text{VaR}_{\Lambda}^+(X))$, we have $\mathbb{P}(X \leq a) \leq \Lambda(a) \leq \Lambda(x)$ since $\Lambda(t)$ is decreasing, which implies $\text{VaR}_{\Lambda(x)}^+(X) \geq a$. Therefore, (A.1) follows since a is chosen arbitrarily. This completes the proof of the lemma. \square

Appendix B. Proofs of results in Section 3

Proof of Proposition 5. It suffices to prove that $\overline{\bigcirc_{i=1}^3 \Lambda_i} = \overline{(\Lambda_1 \bigcirc \Lambda_2) \bigcirc \Lambda_3}$. Denote $\Lambda_{12} = \Lambda_1 \bigcirc \Lambda_2$. First, for any given $y \in \mathbb{R}$, in view of (3.2), there exists a sequence $\{(y_{1n}, y_{2n}, y_{3n})\}_{n \in \mathbb{N}}$ such that $\sum_{i=1}^3 y_{in} = y$ and $\overline{\Lambda_1}(y_{1n}) + \overline{\Lambda_2}(y_{2n}) + \overline{\Lambda_3}(y_{3n}) \rightarrow \overline{\bigcirc_{i=1}^3 \Lambda_i}(y)$ as $n \rightarrow \infty$. Also, it follows from (3.2) that

$$\overline{\Lambda_1}(y_{1n}) + \overline{\Lambda_2}(y_{2n}) + \overline{\Lambda_3}(y_{3n}) \leq \overline{\Lambda_{12}}(y_{1n} + y_{2n}) + \overline{\Lambda_3}(y_{3n}) \leq \overline{\Lambda_{12} \bigcirc \Lambda_3}(y).$$

Letting $n \rightarrow \infty$ yields that

$$\overline{\bigcirc_{i=1}^3 \Lambda_i}(y) \leq \overline{\Lambda_{12} \bigcirc \Lambda_3}(y). \tag{B.1}$$

Next, we prove that the reverse inequality of (B.1) also holds. For fixed $y \in \mathbb{R}$, there exists a sequence $\{(z_n, z_{3n})\}_{n \in \mathbb{N}}$ such that $z_n + z_{3n} = y$ and $\overline{\Lambda_{12}}(z_n) + \overline{\Lambda_3}(z_{3n}) \rightarrow \overline{\Lambda_{12} \bigcirc \Lambda_3}(y)$ as $n \rightarrow \infty$. Also, for each n , there exists a sequence $\{(z_{1,n_j}, z_{2,n_j})\}_{j \in \mathbb{N}}$ such that $z_{1,n_j} + z_{2,n_j} = z_n$ and $\overline{\Lambda_1}(z_{1,n_j}) + \overline{\Lambda_2}(z_{2,n_j}) \rightarrow \overline{\Lambda_{12}}(z_n)$ as $j \rightarrow \infty$. Note that

$$\overline{\Lambda_1}(z_{1,n_j}) + \overline{\Lambda_2}(z_{2,n_j}) + \overline{\Lambda_3}(z_{3,n}) \leq \overline{\bigcirc_{i=1}^3 \Lambda_i}(y).$$

First letting $j \rightarrow \infty$ and then $n \rightarrow \infty$, we have $\overline{\Lambda_{12} \bigcirc \Lambda_3}(y) \leq \overline{\bigcirc_{i=1}^3 \Lambda_i}(y)$. Combining with (B.1), we conclude $\overline{\bigcirc_{i=1}^3 \Lambda_i} = \overline{(\Lambda_1 \bigcirc \Lambda_2) \bigcirc \Lambda_3}$. \square

Proof of Proposition 6. (1) For any sequence $\{(y_{1n}, \dots, y_{mn})\}_{n \in \mathbb{N}}$ such that $\sum_{i=1}^m y_{in} = y_n$ and $y_n \rightarrow -\infty$ as $n \rightarrow \infty$, there exist a subsequence $\{n_k\}$ and some $i \in [m]$ such that $y_{i,n_k} \rightarrow -\infty$ as $k \rightarrow \infty$. Setting $k \rightarrow \infty$ in the following inequality

$$\overline{\Lambda}_i(y_{i,n_k}) + \sum_{j \neq i} \overline{\Lambda}_j(y_{j,n_k}) \leq \overline{\Lambda}_i(y_{i,n_k}) + \sum_{j \neq i} \overline{\Lambda}_j(+\infty),$$

we have

$$\overline{\Lambda}^*(-\infty) \leq \max_{1 \leq i \leq m} \left(\overline{\Lambda}_i(-\infty) + \sum_{j \neq i} \overline{\Lambda}_j(+\infty) \right). \tag{B.2}$$

On the other hand, to prove the reverse inequality of (B.2), assume without loss of generality that

$$\max_{1 \leq i \leq m} \left(\overline{\Lambda}_i(-\infty) + \sum_{j \neq i} \overline{\Lambda}_j(+\infty) \right) = \overline{\Lambda}_1(-\infty) + \sum_{j=2}^m \overline{\Lambda}_j(+\infty).$$

Choose $(x_{1n}, \dots, x_{mn}) = (-nm, n, \dots, n)$. Then $\overline{\Lambda}^*(-n) \geq \overline{\Lambda}_1(-nm) + \sum_{j=2}^m \overline{\Lambda}_j(n)$. Thus, the reverse inequality of (B.2) follows by setting $n \rightarrow +\infty$.

(2) For any $(y_1, \dots, y_m) \in \mathbb{R}^m$, we have $\sum_{i=1}^m \overline{\Lambda}_i(+\infty) \geq \sum_{i=1}^m \overline{\Lambda}_i(y_i)$, implying that $\overline{\Lambda}^*(y) \leq \sum_{i=1}^m \overline{\Lambda}_i(+\infty)$. Thus, $\overline{\Lambda}^*(+\infty) \leq \sum_{i=1}^m \overline{\Lambda}_i(+\infty)$. On the other hand, choosing $y_1 = \dots = y_m = y/m$, we get that $\overline{\Lambda}^*(+\infty) \geq \sum_{i=1}^m \overline{\Lambda}_i(+\infty)$ by setting $y \rightarrow \infty$. The desired result follows. \square

To prove Propositions 7 and 8, we need the following lemma.

Lemma 8. For $m = 2$ and any $y \in \mathbb{R}$, $\Lambda^*(y)$ has one of the following representations:

- (1) there exists $(x_1, x_2) \in \mathbb{R}^2$ such that $x_1 + x_2 = y$ and $\overline{\Lambda}^*(y) = \overline{\Lambda}_1(x_1-) + \overline{\Lambda}_2(x_2+)$;
- (2) there exists $(x_1, x_2) \in \mathbb{R}^2$ such that $x_1 + x_2 = y$ and $\overline{\Lambda}^*(y) = \overline{\Lambda}_1(x_1+) + \overline{\Lambda}_2(x_2-)$;
- (3) there exists $(x_1, x_2) \in \mathbb{R}^2$ such that $x_1 + x_2 = y$ and $\overline{\Lambda}^*(y) = \overline{\Lambda}_1(x_1) + \overline{\Lambda}_2(x_2)$;
- (4) $\overline{\Lambda}^*(y) = (\overline{\Lambda}_1(+\infty) + \overline{\Lambda}_2(-\infty)) \vee (\overline{\Lambda}_1(-\infty) + \overline{\Lambda}_2(+\infty)) = \overline{\Lambda}^*(-\infty)$.

Furthermore, $\Lambda^*(y)$ has one of the former three representations when $\Lambda^*(y) < \Lambda^*(-\infty)$. Additionally, if both Λ_1 and Λ_2 are right-continuous, then only Case (3) occurs.

Proof. For given $y \in \mathbb{R}$, there exists a sequence $\{(y_{1n}, y_{2n})\}_{n \in \mathbb{N}}$ such that $y_{1n} + y_{2n} = y$ and $\overline{\Lambda}_1(y_{1n}) + \overline{\Lambda}_2(y_{2n}) \rightarrow \overline{\Lambda}^*(y)$ as $n \rightarrow \infty$. Two cases arise: First, if $\{(y_{1n}, y_{2n})\}_{n \in \mathbb{N}}$ has a converging subsequence with finite limiting point, that is, there exists $\{n_j\}$ such that $y_{1,n_j} \rightarrow x_1$ and $y_{2,n_j} \rightarrow x_2$ as $j \rightarrow \infty$. This leads to $\overline{\Lambda}^*(y) = \overline{\Lambda}_1(x_1-) + \overline{\Lambda}_2(x_2+)$ or $\overline{\Lambda}_1(x_1+) + \overline{\Lambda}_2(x_2-)$ or $\overline{\Lambda}_1(x_1) + \overline{\Lambda}_2(x_2)$. If Λ_1 and Λ_2 are right-continuous, then $\overline{\Lambda}^*(y) = \overline{\Lambda}_1(x_1) + \overline{\Lambda}_2(x_2)$.

Second, if $\{(y_{1n}, y_{2n})\}_{n \in \mathbb{N}}$ does not have a cluster point, then there exists $\{n_j\}$ such that $y_{1,n_j} \rightarrow +\infty$ ($-\infty$) and $y_{2,n_j} \rightarrow -\infty$ ($+\infty$) as $j \rightarrow \infty$. This leads to Case (4). This proves the lemma. \square

Proof of Proposition 7. If Λ^* does not satisfy (P₁), then there exists a sequence $\{(y_{1n}, \dots, y_{mn})\}_{n \in \mathbb{N}}$ such that $\sum_{i=1}^m y_{in} = y$ and $\sum_{i=1}^m \overline{\Lambda}_i(y_{in}) \rightarrow \overline{\Lambda}^*(y)$ as $n \rightarrow \infty$. Moreover, $\{y_{in}\}_{n \in \mathbb{N}}$ does not have a cluster point in for some i . Without loss of generality, assume that

$y_{1,n_j} \rightarrow -\infty$ as $j \rightarrow \infty$. Note that $\overline{\Lambda}_i(y_{in}) \leq \overline{\Lambda}_i(+\infty)$ for $i \geq 2$. This implies that $\overline{\Lambda}^*(y) \leq \overline{\Lambda}_1(-\infty) + \sum_{i=2}^m \overline{\Lambda}_i(+\infty)$. On the other hand, by Proposition 6, we have

$$\overline{\Lambda}^*(y) \geq \overline{\Lambda}^*(-\infty) = \max_{1 \leq i \leq m} \left(\overline{\Lambda}_i(-\infty) + \sum_{j \neq i} \overline{\Lambda}_j(+\infty) \right).$$

Therefore, $\overline{\Lambda}^*(y) = \overline{\Lambda}^*(-\infty)$.

Now suppose that $\Lambda^*(y_0)$ has Property (P₂). Assume on the contrary that there exists $x_0 < y_0$ and $\Lambda^*(x_0)$ satisfies (P₁), that is, there exists $(y_1, \dots, y_m) \in \mathbb{R}^m$ such that $\sum_{i=1}^m y_i = x_0$ and $\overline{\Lambda}^*(x_0) = \sum_{i=1}^m \overline{\Lambda}_i(y_i)$. Then

$$\overline{\Lambda}^*(y_0) \geq \overline{\Lambda}_1(y_1 + y_0 - x_0) + \sum_{i=2}^m \overline{\Lambda}_i(y_i) \geq \overline{\Lambda}^*(x_0).$$

However, $\overline{\Lambda}^*(y_0) = \overline{\Lambda}^*(x_0) = \overline{\Lambda}^*(-\infty)$. So we get $\overline{\Lambda}^*(y_0) = \overline{\Lambda}_1(y_1 + y_0 - x_0) + \sum_{i=2}^m \overline{\Lambda}_i(y_i)$. This means that $\Lambda^*(y_0)$ does not satisfy (P₂), which is a contradiction. The desired result follows. □

Proof of Proposition 8. We give the proof only for $m = 2$ since, in view of Proposition 5, the proof of the general case $m \geq 3$ follows by induction.

(1) For $m = 2$, assume without loss of generality that Λ_1 is continuous. By Lemma 6, for any $x \in \mathbb{R}$, either one of the following two cases holds for $\Lambda(x)$:

Case 1.1. There exists $(x_1, x_2) \in \mathbb{R}^2$ such that $x_1 + x_2 = x$ and $\overline{\Lambda}^*(x) = \overline{\Lambda}_1(x_1) + \overline{\Lambda}_2(x_2)$.

Case 1.2. $\overline{\Lambda}^*(x) = (\overline{\Lambda}_1(+\infty) + \overline{\Lambda}_2(-\infty)) \vee (\overline{\Lambda}_1(-\infty) + \overline{\Lambda}_2(+\infty)) = \overline{\Lambda}^*(-\infty)$.

If Λ^* is not continuous, there exists some $x_0 \in \mathbb{R}$ such that $\Lambda^*(x_0) > \Lambda^*(x_0+)$ or $\Lambda^*(x_0) < \Lambda^*(x_0-)$.

First, we prove that Λ^* is right-continuous by contradiction. Assume on the contrary that there exists some $x_0 \in \mathbb{R}$ such that $\Lambda^*(x_0) > \Lambda^*(x_0+)$. Choose $\{x_n\}_{n \in \mathbb{N}}$ such that $x_n \rightarrow x_0+$. By Lemma 6, $\Lambda^*(x_n) < \Lambda^*(-\infty)$ and Case 1.1 applies, i.e. there exists a sequence $\{(x_{1n}, x_{2n})\}_{n \in \mathbb{N}}$ such that $x_{1n} + x_{2n} = x_n$ and $\overline{\Lambda}_1(x_{1n}) + \overline{\Lambda}_2(x_{2n}) = \overline{\Lambda}^*(x_n)$. If $\{x_{1n}\}_{n \in \mathbb{N}}$ has a cluster point, there exists $\{n_k\}$ such that $x_{1,n_k} \rightarrow y_1$ and $x_{2,n_k} \rightarrow y_2$ as $k \rightarrow \infty$, where $y_1 + y_2 = x_0$. Thus,

$$\overline{\Lambda}^*(x_0+) = \lim_{k \rightarrow \infty} \overline{\Lambda}_1(x_{1,n_k}) + \overline{\Lambda}_2(x_{2,n_k}) = \overline{\Lambda}_1(y_1) + \overline{\Lambda}_2(y_2).$$

By (3.2), it follows that $\overline{\Lambda}_1(y_1) + \overline{\Lambda}_2(y_2) \leq \overline{\Lambda}^*(x_0)$. This contradicts with $\Lambda^*(x_0) > \Lambda^*(x_0+)$. If $\{x_{1n}\}_{n \in \mathbb{N}}$ does not have a cluster point, then there exists $\{n_k\}$ such that $x_{1,n_k} \rightarrow +\infty$ and $x_{2,n_k} \rightarrow -\infty$ as $k \rightarrow \infty$. Thus,

$$\begin{aligned} \overline{\Lambda}^*(x_0+) &= \lim_{j \rightarrow \infty} \overline{\Lambda}_1(x_{1,n_j}) + \overline{\Lambda}_2(x_{2,n_j}) \\ &= (\overline{\Lambda}_1(+\infty) + \overline{\Lambda}_2(-\infty)) \vee (\overline{\Lambda}_1(-\infty) + \overline{\Lambda}_2(+\infty)) = \overline{\Lambda}^*(-\infty), \end{aligned}$$

implying that $\Lambda^*(x_0) = \Lambda^*(x_0+)$ by the monotonicity of Λ^* . This leads to a contradiction and proves right-continuity of Λ^* .

Next, we prove that Λ^* is left-continuous, also by contradiction. Assume that there exists $x_0 \in \mathbb{R}$ such that $\Lambda^*(x_0-) > \Lambda^*(x_0)$. Denote $\epsilon = \Lambda^*(x_0-) - \Lambda^*(x_0)$. By Lemma 6, there

exists $(x_1, x_2) \in \mathbb{R}^2$ satisfying that $x_1 + x_2 = x_0$ and $\overline{\Lambda_1}(x_1) + \overline{\Lambda_2}(x_2+) = \overline{\Lambda^*}(x_0)$. Since Λ_1 is continuous, it follows that $\overline{\Lambda_1}(x) > \overline{\Lambda_1}(x_1) - \epsilon/2$ for $x \in (x_1 - \delta, x_1)$, where $\delta > 0$ is small enough. Then

$$\begin{aligned} \overline{\Lambda^*}\left(x_0 - \frac{\delta}{2}\right) &\geq \overline{\Lambda_1}\left(x_1 - \frac{\delta}{2}\right) + \overline{\Lambda_2}(x_2+) \\ &> \overline{\Lambda_1}(x_1) - \frac{\epsilon}{2} + \overline{\Lambda_2}(x_2+) = \overline{\Lambda^*}(x_0) - \frac{\epsilon}{2}, \end{aligned}$$

which implies that $\epsilon = \overline{\Lambda^*}(x_0) - \overline{\Lambda^*}(x_0 - \delta/2) > \epsilon/2 > \overline{\Lambda^*}(x_0) - \overline{\Lambda^*}(x_0 - \delta/2)$. This contradicts with $\Lambda^*(x_0 - \delta/2) \leq \Lambda^*(x_0 - \delta/2)$ and proves Part (1).

(2) Assume on the contrary that Λ^* is not right-continuous, i.e. $\Lambda^*(x_0) > \Lambda^*(x_0+)$ for some $x_0 \in \mathbb{R}$. Choose $x_n \rightarrow x_0+$. By Lemma 6 and the right-continuity of Λ_1 and Λ_2 , it follows that $\Lambda^*(x_n) < \Lambda^*(-\infty)$ and that there exists $(x_{1n}, x_{2n}) \in \mathbb{R}^2$ such that $x_{1n} + x_{2n} = x_n$ and $\overline{\Lambda_1}(x_{1n}) + \overline{\Lambda_2}(x_{2n}) = \overline{\Lambda^*}(x_n)$.

If $\{x_{1n}\}_{n \in \mathbb{N}}$ has a cluster point, then there exists $\{n_k\}$ such that $x_{1,n_k} \rightarrow y_1$ and $x_{2,n_k} \rightarrow y_2$ as $k \rightarrow \infty$, where $y_1 + y_2 = x_0$. Now, two cases arise.

Case 2.1 Suppose that $x_{1,n_k} \rightarrow y_1+$ and $x_{2,n_k} \rightarrow y_2+$ as $k \rightarrow \infty$. In this case,

$$\begin{aligned} \overline{\Lambda^*}(x_0+) &= \lim_{k \rightarrow \infty} \overline{\Lambda_1}(x_{1,n_k}) + \overline{\Lambda_2}(x_{2,n_k}) \\ &= \overline{\Lambda_1}(y_1+) + \overline{\Lambda_2}(y_2+) = \overline{\Lambda_1}(y_1) + \overline{\Lambda_2}(y_2), \end{aligned}$$

where the last equality follows from the right-continuity of Λ_1 and Λ_2 . Moreover, from the definition of Λ^* , we have $\overline{\Lambda_1}(y_1) + \overline{\Lambda_2}(y_2) \leq \overline{\Lambda^*}(x_0)$. Thus, $\overline{\Lambda^*}(x_0+) \leq \overline{\Lambda^*}(x_0)$. This contradicts with $\Lambda^*(x_0) > \Lambda^*(x_0+)$.

Case 2.2 Suppose that $x_{1,n_k} \rightarrow y_1+$ and $x_{2,n_k} \rightarrow y_2-$ as $k \rightarrow \infty$. In this case,

$$\begin{aligned} \overline{\Lambda^*}(x_0+) &= \lim_{k \rightarrow \infty} \overline{\Lambda_1}(x_{1,n_k}) + \overline{\Lambda_2}(x_{2,n_k}) \\ &= \overline{\Lambda_1}(y_1+) + \overline{\Lambda_2}(y_2-) = \overline{\Lambda_1}(y_1) + \overline{\Lambda_2}(y_2-), \end{aligned}$$

and $\overline{\Lambda_1}(y_1) + \overline{\Lambda_2}(y_2-) \leq \overline{\Lambda_1}(y_1) + \overline{\Lambda_2}(y_2) \leq \overline{\Lambda^*}(x_0)$. Thus, $\overline{\Lambda^*}(x_0+) \leq \overline{\Lambda^*}(x_0)$. This is also a contradiction.

If $\{x_{1n}\}_{n \in \mathbb{N}}$ does not have a cluster point, a similar argument to that of Part (1) yields the desired result. □

Proof of Proposition 9. If $\Lambda^*(x) < \Lambda^*(-\infty)$, the desired result follows from Proposition 7. Now assume that $\Lambda^*(x) = \Lambda^*(-\infty)$. Note that for any $k \in [m]$,

$$\overline{\Lambda^*}(x) \geq \overline{\Lambda_k}\left(x - \sum_{i \neq k} y_i\right) + \sum_{i \neq k}^m \overline{\Lambda_i}(y_i) \geq \overline{\Lambda_k}(-\infty) + \sum_{i \neq k} \overline{\Lambda_i}(+\infty). \tag{B.3}$$

By Proposition 6, it follows from (B.3) that there exists some k_0 such that $\overline{\Lambda_{k_0}}\left(x - \sum_{i \neq k_0} y_i\right) = \overline{\Lambda_{k_0}}(-\infty)$, and the equality in (B.3) holds for $k = k_0$. Therefore, $\overline{\Lambda^*}(x) = \sum_{i=1}^m \overline{\Lambda_i}(x_i)$ holds by choosing $x_{k_0} = x - \sum_{i \neq k_0} y_i$ and $x_i = y_i$ for $i \neq k_0$ with $\sum_{i=1}^m x_i = x$. □

Proof of Proposition 10. (1) Necessity. Since Λ^* is constant, $\Lambda^*(-\infty) = \Lambda^*(+\infty)$. By Proposition 6, there exists $k_0 \in [m]$ such that

$$\Lambda^*(-\infty) = 1 - \overline{\Lambda}_{k_0}(-\infty) - \sum_{j \neq k_0} \overline{\Lambda}_j(+\infty), \quad \Lambda^*(+\infty) = 1 - \sum_{i=1}^m \overline{\Lambda}_i(+\infty),$$

implying that $\Lambda_{k_0}(-\infty) = \Lambda_{k_0}(+\infty)$, i.e. Λ_{k_0} is constant.

Sufficiency. Suppose that $\Lambda_{i_0} \equiv c$. By Proposition 6, we have $\overline{\Lambda}^*(+\infty) = 1 - a + \sum_{j \neq i_0} \overline{\Lambda}_j(+\infty)$. Also, $\overline{\Lambda}^*(-\infty) \geq 1 - a + \sum_{j \neq i_0} \overline{\Lambda}_j(+\infty) = \overline{\Lambda}^*(+\infty)$. Therefore, $\overline{\Lambda}^*(x) \equiv 1 - a + \sum_{j \neq i_0} \overline{\Lambda}_j(+\infty)$ for all $x \in \mathbb{R}$.

(2) Since Λ^* is constant, by part (1) there exists k_0 such that Λ_{k_0} is constant. Without loss of generality, assume $k_0 = 1$.

Sufficiency. Suppose that $\Lambda_{i_0}(y) > \Lambda_{i_0}(+\infty)$ for some i_0 and all $y \in \mathbb{R}$. Then, for any $(x_1, \dots, x_m) \in \mathbb{R}^m$ with $x = \sum_{i=1}^m x_i$, $\sum_{i=1}^m \overline{\Lambda}_i(x_i) < \sum_{i=1}^m \overline{\Lambda}_i(+\infty) = \overline{\Lambda}^*(+\infty)$. Thus, $\overline{\Lambda}^*(x) = \overline{\Lambda}^*(+\infty) > \sum_{i=1}^m \overline{\Lambda}_i(x_i)$.

Necessity. Assume on the contrary that for all Λ_i , there exists $x_i \in \mathbb{R}$ such that $\Lambda_i(x_i) = \Lambda_i(+\infty)$. Since Λ_1 is constant, it follows that for any $x \in \mathbb{R}$,

$$\overline{\Lambda}_1 \left(x - \sum_{i=2}^m x_i \right) + \sum_{i=2}^m \overline{\Lambda}_i(x_i) = \sum_{i=1}^m \overline{\Lambda}_i(+\infty) = \overline{\Lambda}^*(+\infty) = \overline{\Lambda}^*(x).$$

This contradicts with the assumption that $\overline{\Lambda}^*(x) > \sum_{i=1}^m \overline{\Lambda}_i(y_i)$ with $y_1 = x - \sum_{i=2}^m x_i$ and $y_k = x_k$ for $k \geq 2$. This proves the desired result. \square

Appendix C. Proofs of results in Section 4

Proof of Lemma 4. If $\lambda_{i_0} = 0$ for some $i_0 \in [m]$, then $\text{VaR}_{\lambda_i}^-(Y) = -\infty$ for $Y \in L^0$. Thus, the right-hand side (RHS) of (4.2) is $\sum_{i=1}^m y_i$. Note that the left-hand side (LHS) of (4.2) is larger than or equal to $\sum_{i=1}^m y_i$ and that the lower bound can be attained by choosing $X_{i_0} = X - \sum_{j \neq i_0} y_j$ and $X_j = y_j$ for $j \neq i_0$. Thus, (4.2) holds for this special case. So we assume that $\lambda_i \in (0, 1]$ for $i \in [m]$.

If $\square_{i=1}^m \text{VaR}_{\lambda_i}^-(X) \leq \sum_{i=1}^m y_i$, by cash invariance of VaR and Corollary 2 in [12], there exists optimal allocation $(X_1, \dots, X_m) \in \mathbb{A}_m$ for $\square_{i=1}^m \text{VaR}_{\lambda_i}^-(X)$ such that $\text{VaR}_{\lambda_i}^-(X_i) \leq y_i$ for $i \in [m]$. Thus, $\sum_{i=1}^m \text{VaR}_{\lambda_i}^-(X_i) \vee y_i = \sum_{i=1}^m y_i$, attaining the lower bound of LHS. So, (4.2) holds for this case.

If $\square_{i=1}^m \text{VaR}_{\lambda_i}^-(X) > \sum_{i=1}^m y_i$, also by cash invariance of VaR and Corollary 2 in [12], there exists optimal allocation $(X_1, \dots, X_m) \in \mathbb{A}_m$ for $\square_{i=1}^m \text{VaR}_{\lambda_i}^-(X)$ such that $\text{VaR}_{\lambda_i}^-(X_i) > y_i$ for $i \in [m]$. Thus, we have

$$\sum_{i=1}^m \text{VaR}_{\lambda_i}^-(X_i) \vee y_i = \sum_{i=1}^m \text{VaR}_{\lambda_i}^-(X_i) = \square_{i=1}^m \text{VaR}_{\lambda_i}^-(X),$$

implying that

$$\text{LHS} \leq \square_{i=1}^m \text{VaR}_{\lambda_i}^-(X). \tag{C.1}$$

Next, we show that the reverse inequality of (C.1) is also true. To this end, for any other allocation $(Y_1, \dots, Y_m) \in \mathbb{A}_m(X)$, denote $K = \{k : \text{VaR}_{\lambda_k}^-(Y_k) < y_i, k \in [m]\}$. Then,

$$\begin{aligned} \sum_{i=1}^m \text{VaR}_{\lambda_i}^-(Y_i) \vee y_i &= \sum_{i \in K} y_i + \sum_{i \in [m] \setminus K} \text{VaR}_{\lambda_i}^-(Y_i) \\ &\geq \sum_{i \in K} y_i + \square_{i=1}^m \text{VaR}_{\lambda_i}^-(X) - \sum_{i \in K} \text{VaR}_{\lambda_i}^-(Y_i) \geq \square_{i=1}^m \text{VaR}_{\lambda_i}^-(X), \end{aligned}$$

where the first inequality follows from the fact that $\sum_{i=1}^m \text{VaR}_{\lambda_i}^-(Y_i) \geq \square_{i=1}^m \text{VaR}_{\lambda_i}^-(X)$. Thus, the reverse inequality of (C.1) holds. This completes the proof. \square

Proof of Theorem 3. The proof is similar to those of Theorems 1 and 2. By Lemmas 2, 3 and 5, we have

$$\begin{aligned} \square_{i=1}^m \text{VaR}_{\Lambda_i}^{\kappa_i}(X) &= \inf_{(X_1, \dots, X_m) \in \mathbb{A}_m(X)} \sum_{i=1}^m \text{VaR}_{\Lambda_i}^{\kappa_i}(X_i) \\ &= \inf_{(X_1, \dots, X_m) \in \mathbb{A}_m(X)} \sum_{i=1}^m \inf_{y_i \in \mathbb{R}} \left\{ \text{VaR}_{\Lambda_i(y_i)}^{\kappa_i}(X_i) \vee y_i \right\} \\ &= \inf_{y_1, \dots, y_m \in \mathbb{R}} \inf_{(X_1, \dots, X_m) \in \mathbb{A}_m(X)} \left\{ \sum_{i=1}^m \text{VaR}_{\Lambda_i(y_i)}^{\kappa_i}(X_i) \vee y_i \right\} \\ &= \inf_{y_1, \dots, y_m \in \mathbb{R}} \left\{ \square_{i=1}^m \text{VaR}_{\Lambda_i(y_i)}^{\kappa_i}(X) \vee \sum_{i=1}^m y_i \right\} \\ &= \inf_{y_1, \dots, y_m \in \mathbb{R}} \left\{ \text{VaR}_{1 - \sum_{i=1}^m \overline{\Lambda}_i(y_i)}^+(X) \vee \sum_{i=1}^m y_i \right\} \tag{C.2} \\ &= \text{VaR}_{\Lambda^*}^+(X), \end{aligned}$$

where (C.2) follows from Theorem 1 of [21]. The rest of the proof is the same as that of Theorem 2 and, hence, omitted. \square

Proof of Theorem 11. (1) Suppose that $\Lambda^*(-\infty) < 0$. Then, for any $x < \text{ess-inf}(X)$, there exists $(x_1, \dots, x_m) \in \mathbb{R}^m$ such that $\sum_{i=1}^m x_i = x$ and $\sum_{i=1}^m \overline{\Lambda}_i(x_i) > 1$. Let $\{A_1, \dots, A_m\}$ be a partition of Ω , satisfying

$$\mathbb{P}(A_i) = \frac{\overline{\Lambda}_i(x_i)}{\sum_{j=1}^m \overline{\Lambda}_j(x_j)}, \quad i \in [m].$$

Define $X_j = (X - x)1_{A_j} + x_j$ for $j \in [m - 1]$, and $X_m = X - \sum_{j=1}^{m-1} X_j$. Then $\mathbb{P}(X_i > x_i) \leq \mathbb{P}(A_i) < \overline{\Lambda}_i(x_i)$, implying $\text{VaR}_{\Lambda_i}^+(X_i) \leq x_i$ for $i \in [m]$. Thus, $\sum_{i=1}^m \text{VaR}_{\Lambda_i}^+(X_i) \leq \sum_{i=1}^m x_i = x$. This proves part (1) by letting $x \searrow -\infty$.

(2) First, consider the case $L = \emptyset$. For any $x < \text{ess-inf}(X)$, we have $\overline{\Lambda}^*(x) > 1$ or $\overline{\Lambda}^*(x) = 1$. If $\overline{\Lambda}^*(x) = 1$, then $x \notin L$ and, hence, there exists $(x_1, \dots, x_m) \in \mathbb{R}^m$ such that $\sum_{i=1}^m x_i = x$ and

$\sum_{i=1}^m \overline{\Lambda}_i(x_i) = 1$. Therefore, we can always choose $(x_1, \dots, x_m) \in \mathbb{R}^m$ such that $\sum_{i=1}^m x_i = x$ and $\sum_{i=1}^m \overline{\Lambda}_i(x_i) \geq 1$ whenever $\overline{\Lambda}^*(x) > 1$ or $\overline{\Lambda}^*(x) = 1$. Construct (X_1, \dots, X_m) as in part (1). Then, $\mathbb{P}(X_j > x_j) \leq \mathbb{P}(A_j) \leq \overline{\Lambda}_j(x_j)$ for $j \in [m]$, implying $\text{VaR}_{\overline{\Lambda}_j}^-(X_j) \leq x_j$. Therefore, $\square_{i=1}^m \text{VaR}_{\overline{\Lambda}_i}^-(X) = -\infty$.

Next, consider $L \neq \emptyset$. Two subcases arise.

Subcase 1: Suppose $\sup L < \text{ess-inf}(X)$. By Lemma 6, we get $\square_{i=1}^m \text{VaR}_{\overline{\Lambda}_i}^-(X) \geq \sup L$. For any $x \in (\sup L, \text{ess-inf}(X))$, there exists $(x_1, \dots, x_m) \in \mathbb{R}^m$ such that $\sum_{i=1}^m x_i = x$ and $\sum_{i=1}^m \overline{\Lambda}_i(x_i) \geq 1$. Construct $X \in \mathbb{A}_m(X)$ as in part (1). Similarly, we have $\text{VaR}_{\overline{\Lambda}_j}^-(X_j) \leq x_j$ for $j \in [m]$. Thus, $\sum_{i=1}^m \text{VaR}_{\overline{\Lambda}_i}^-(X_i) \leq x$, yielding $\square_{i=1}^m \text{VaR}_{\overline{\Lambda}_i}^-(X) \leq \sup L$ by setting $x \searrow \sup L$. Therefore, $\square_{i=1}^m \text{VaR}_{\overline{\Lambda}_i}^-(X) = \sup L$.

Subcase 2: Suppose $\sup L \geq \text{ess-inf}(X)$. By part (1) of Lemma 6, we get that $\square_{i=1}^m \text{VaR}_{\overline{\Lambda}_i}^-(X) \geq \text{ess-inf}(X)$. Also, by part (2) of Lemma 6, $\square_{i=1}^m \text{VaR}_{\overline{\Lambda}_i}^-(X) \leq \text{ess-inf}(X)$. Thus,

$$\square_{i=1}^m \text{VaR}_{\overline{\Lambda}_i}^-(X) = \text{ess-inf}(X) = \min \{ \sup L, \text{ess-inf}(X) \}.$$

This completes the proof of the proposition. □

Proof of Proposition 12. (1) The proof is similar to that of part (1) of Proposition 11.

(2) First, consider $T = \emptyset$. Then, for any $x < \text{ess-inf}(X)$, we have $\Lambda^*(x) < 0$. Thus, there exists $(x_1, \dots, x_m) \in \mathbb{R}^m$ such that $x = \sum_{i=1}^m x_i$ and $\sum_{i=1}^m \overline{\Lambda}_i(x_i) > 1$. Construct X as in the proof of part (1) of Proposition 11. Then $\mathbb{P}(X_j > x_j) \leq \mathbb{P}(A_j) < \overline{\Lambda}_j(x_j)$ for $j \in [m]$, implying $\text{VaR}_{\overline{\Lambda}_j}^+(X_j) \leq x_j$. Therefore, $\square_{i=1}^m \text{VaR}_{\overline{\Lambda}_i}^+(X) = -\infty$.

Next, consider $T \neq \emptyset$. Two subcases arise.

Subcase 1: Suppose $\sup T < \text{ess-inf}(X)$. For any $x \in (\sup T, \text{ess-inf}(X))$, there exists $(x_1, \dots, x_m) \in \mathbb{R}^m$ such that $\sum_{i=1}^m x_i = x$ and $\sum_{i=1}^m \overline{\Lambda}_i(x_i) > 1$. Construct $X \in \mathbb{A}_m(X)$ as in part (1). Similarly, we have $\text{VaR}_{\overline{\Lambda}_j}^+(X_j) \leq x_j$ for $j \in [m]$. Thus, $\sum_{i=1}^m \text{VaR}_{\overline{\Lambda}_i}^+(X_i) \leq x$, yielding $\square_{i=1}^m \text{VaR}_{\overline{\Lambda}_i}^+(X) \leq \sup L$ by setting $x \searrow \sup T$. On the other hand, by Lemma 7, for any $x \in T$, we have $\square_{i=1}^m \text{VaR}_{\overline{\Lambda}_i}^+(X) \geq x$. Thus, $\square_{i=1}^m \text{VaR}_{\overline{\Lambda}_i}^+(X) \geq \sup L$. This proves part (2) in Subcase 1.

Subcase 2: Suppose $\sup T \geq \text{ess-inf}(X)$. By part (1) of Lemma 7, we get that $\square_{i=1}^m \text{VaR}_{\overline{\Lambda}_i}^+(X) \geq \text{ess-inf}(X)$. Also, by part (2) of Lemma 7, $\square_{i=1}^m \text{VaR}_{\overline{\Lambda}_i}^+(X) \leq \text{ess-inf}(X)$. Thus, $\square_{i=1}^m \text{VaR}_{\overline{\Lambda}_i}^+(X) = \text{ess-inf}(X) = \min \{ \sup T, \text{ess-inf}(X) \}$. This completes the proof of the proposition. □

Proof of Proposition 13. The proof is similar to that of Proposition 12. □

Proof of Lemma 6. (1) Assume on the contrary that $\square_{i=1}^m \text{VaR}_{\overline{\Lambda}_i}^-(X) < x_0$. Then there exists $(X_1, \dots, X_m) \in \mathbb{A}_m(X)$ such that $\sum_{i=1}^m \text{VaR}_{\overline{\Lambda}_i}^-(X_i) < x_0$. Denote $x_i = \text{VaR}_{\overline{\Lambda}_i}^-(X_i)$ for $i \in [m]$. Since Λ_i is right-continuous, by Lemma 1 it follows that $\mathbb{P}(X_i > x_i) \leq \overline{\Lambda}_i(x_i)$. Thus, $\sum_{i=1}^m \mathbb{P}(X_i > x_i) \leq \sum_{i=1}^m \overline{\Lambda}_i(x_i) \leq \overline{\Lambda}^*(x_0) = 1$. However, $1 = \mathbb{P}(X \geq x_0) \leq \sum_{i=1}^m \mathbb{P}(X_i > x_i)$. This leads a contradiction.

(2) For any $x > \text{ess-inf}(X)$, $\mathbb{P}(X > x) < 1$. From the definition of Λ^* and its monotonicity, it follows that

$$\overline{\Lambda}^*(x) = \sup_{y_1 + \dots + y_m = x} \sum_{i=1}^m \overline{\Lambda}_i(y_i) \geq \overline{\Lambda}^*(-\infty) \geq 1,$$

implying that there exists $(x_1, \dots, x_m) \in \mathbb{R}^m$ such that $\sum_{i=1}^m x_i = x$ and $\sum_{i=1}^m \overline{\Lambda}_i(x_i) > \mathbb{P}(X > x)$. Let $\{A_1, \dots, A_m\}$ be a partition of the set $\{X > x\}$, satisfying that

$$\mathbb{P}(A_i) = \mathbb{P}(X > x) \cdot \frac{\overline{\Lambda}_i(x_i)}{\sum_{j=1}^m \overline{\Lambda}_j(x_j)}, \quad i \in [m],$$

and define $X_m = X - \sum_{j=1}^{m-1} X_j$, where $X_j = (X - x) 1_{A_j} + x_j$ for $j \in [m - 1]$. Obviously, $(X_1, \dots, X_m) \in \mathbb{A}_m(X)$ and $\mathbb{P}(X_i > x_i) \leq \mathbb{P}(A_i) < \overline{\Lambda}_i(x_i)$ for $i \in [m]$, implying $\text{VaR}_{\Lambda_i}^-(X_i) \leq x_i$ for $i \in [m]$. So we get $\sum_{i=1}^m \text{VaR}_{\Lambda_i}^-(X_i) \leq x$. The desired result now follows by setting $x \downarrow \text{ess-inf}(X)$. \square

Proof of Lemma 7. (1) Assume on the contrary that $\square_{i=1}^m \text{VaR}_{\Lambda_i}^+(X) < x_0$. Then there exists $(X_1, \dots, X_m) \in \mathbb{A}_m(X)$ such that $\sum_{i=1}^m \text{VaR}_{\Lambda_i}^+(X_i) < x_0$. Denote $x_i = \text{VaR}_{\Lambda_i}^+(X_i)$ for $i \in [m]$, and set $\epsilon = x_0 - \sum_{i=1}^m x_i$. Since $\Lambda^*(x_0) = 0$, it follows that $\sum_{i=1}^m \overline{\Lambda}_i(x_i) \leq 1$. Since $x_1 + \epsilon/2 + \sum_{i=2}^m x_i < x_0$, we have

$$\overline{\Lambda}_1\left(x_1 + \frac{\epsilon}{2}\right) + \sum_{i=2}^m \overline{\Lambda}_i(x_i) \leq \overline{\Lambda}^*(x_0) = 1.$$

By Lemma 1 and right-continuity of Λ_i , it follows that $\mathbb{P}(X_i > x_i) \leq \overline{\Lambda}_i(x_i)$ and $\mathbb{P}(X_1 > x_1 + \epsilon/2) < \overline{\Lambda}_1(x_1 + \epsilon/2)$. Thus,

$$\begin{aligned} 1 &= \mathbb{P}\left(X > \sum_{i=1}^m x_i + \frac{\epsilon}{2}\right) \leq \mathbb{P}\left(X_1 > x_1 + \frac{\epsilon}{2}\right) + \sum_{i=2}^m \mathbb{P}(X_i > x_i) \\ &< \overline{\Lambda}_1\left(x_1 + \frac{\epsilon}{2}\right) + \sum_{i=2}^m \overline{\Lambda}_i(x_i) \leq 1, \end{aligned}$$

which is a contradiction. This proves part (1).

(2) The proof is the same as that of part (2) of Lemma 6. \square

Appendix D. Proofs of results in Section 5

Proof of Theorem 6. Eq. (5.10) follows from Theorem 2. We focus on constructing its optimal and asymptotically optimal allocations according to different situations.

(1) Suppose that $\mathbb{P}(X > x_0) < \overline{\Lambda}^*(x_0)$. Let $\mathbf{X} \in \mathbb{A}_m(X)$ be constructed as in the proof of Theorem 4. It is easy to see that $\mathbb{P}(X_j > x_j) < \overline{\Lambda}_j(x_j)$, implying $\text{VaR}_{\Lambda_j}^+(X_j) \leq x_j$ for $j \in [m]$. By (2.6) in Lemma 1, we have

$$\mathbb{P}(X_j < x_j) \leq \mathbb{P}(X < x_0) \leq \Lambda(x_0-) \leq \Lambda_j(x_j-),$$

implying $\text{VaR}_{\Lambda_j}^+(X_j) \geq x_j$ for $j \in [m]$, where the last inequality follows from the fact that $\Lambda(y) \leq \Lambda_j(y)$ for any $y \in \mathbb{R}$. Thus, $\text{VaR}_{\Lambda_j}^+(X_j) = x_j$ for $j \in [m]$, i.e. \mathbf{X} is an optimal allocation of X .

(2) Suppose that $\mathbb{P}(X > x_0) = \overline{\Lambda}^*(x_0)$, and $\mathbb{P}(X > x_0 + \epsilon) < \mathbb{P}(X > x_0)$ for any $\epsilon > 0$. Let $\{B_n\}_{n \in \mathbb{N}}$ be a partition of $\{X > x_0\}$, defined by $B_1 = \{X > x_0 + 1\}$ and

$$B_n = \left\{x_0 + \frac{1}{n} < X \leq x_0 + \frac{1}{n-1}\right\}, \quad n \geq 2.$$

For $k \geq 1$, let $\{B_{k1}, \dots, B_{km}\}$ be a partition of B_k , satisfying

$$\mathbb{P}(B_{kj}) = \mathbb{P}(B_k) \cdot \frac{\overline{\Lambda}_j(x_j)}{\sum_{i=1}^m \overline{\Lambda}_i(x_i)}, \quad j \in [m].$$

Denote $C_j = \bigcup_{k \geq 1} B_{kj}$ for $j \in [m]$. Thus, $\{C_1, \dots, C_m\}$ constitutes a partition of $\{X > x_0\}$. Construct an allocation of X as follows:

$$X_i = x_i + (X - x_0) 1_{C_i}, \quad i \in [m - 1]; \quad X_m = X - \sum_{j=1}^{m-1} X_j. \tag{D.1}$$

Note that $\mathbb{P}(X_j > x_j) = \mathbb{P}(C_j) = \overline{\Lambda}_j(x_j)$ for $j \in [m]$ and that $\mathbb{P}(X_j > x_j + \epsilon) < \mathbb{P}(C_j) = \overline{\Lambda}_j(x_j)$ for any $\epsilon > 0$. Thus, $\text{VaR}_{\Lambda_j}^+(X_j) = x_j$ for $j \in [m]$. This proves part (2).

(3) First, suppose that $\Lambda_j(x_j + \epsilon) < \Lambda_j(x_j)$ for any $\epsilon > 0$. Let $X \in \mathbb{A}_m(X)$ be defined as in (D.1). Then $\mathbb{P}(X_j > x_j) = \overline{\Lambda}_j(x_j) < \overline{\Lambda}_j(x_j + \epsilon)$, implying $\text{VaR}_{\Lambda_j}^+(X_j) = x_j$ for $j \in [m]$. Thus, X is an optimal allocation of X .

Next, consider the second half of part (3). We prove that no optimal allocation exists by contradiction. Assume on the contrary that there exists an optimal allocation $X \in \mathbb{A}_m(X)$. Denote $y_j = \text{VaR}_{\Lambda_j}^+(X_j)$ for $j \in [m]$, satisfying $\sum_{i=1}^m y_i = x_0$. By the assumption of part (3), there exists k , say, $k = 1$, such that $\Lambda_k(y_k) = \Lambda_k(y_k + \tau_0)$. Denote $\epsilon_1 = \min\{\epsilon_0, \tau_0\} / 2$. Then $\mathbb{P}(X_1 > y_1 + \epsilon_1) < \overline{\Lambda}_1(y_1 + \epsilon_1) = \overline{\Lambda}_1(y_1)$, and $\mathbb{P}(X_k > y_k) \leq \overline{\Lambda}_k(y_k)$ for $k \geq 2$. Hence,

$$\begin{aligned} \mathbb{P}(X > x_0) &= \mathbb{P}(X > x_0 + \epsilon_1) \leq \mathbb{P}(X_1 > y_1 + \epsilon_1) + \sum_{i=2}^m \mathbb{P}(X_i > y_i) \\ &< \sum_{i=1}^m \overline{\Lambda}_i(y_i) = \overline{\Lambda}^*(x_0), \end{aligned}$$

which contradicts with the assumption $\mathbb{P}(X > x_0) = \overline{\Lambda}^*(x_0)$.

Let $(X_{1n}, \dots, X_{mn}) \in \mathbb{A}_m(X)$ be as defined by (5.8). By a similar argument to that of part (2) in Theorem 5, we have

$$\mathbb{P}\left(X_{1n} > x_{1n} - \frac{1}{n}\right) < \overline{\Lambda}_1(x_{1n}), \quad \mathbb{P}(X_{kn} > x_{kn}) < \overline{\Lambda}_k(x_{kn}), \quad k \geq 2,$$

implying that $\text{VaR}_{\Lambda_i}^+(X_{in}) \leq x_{in}$ for $i \in [m]$. Hence, $\sum_{i=1}^m \text{VaR}_{\Lambda_i}^+(X_{in}) \leq \sum_{i=1}^m x_{in} = x_0 + 1/n$. By Theorem 2, the desired statement follows by letting $n \rightarrow +\infty$. \square

Proof of Theorem 7. (1) Suppose that $\mathbb{P}(X > x_0) < \overline{\Lambda}^*(x_0)$. In this case, there exists $(x_1, \dots, x_m) \in \mathbb{R}^m$ such that $x_0 = \sum_{i=1}^m x_i$ and $\sum_{i=1}^m \overline{\Lambda}_i(x_i) \in (\mathbb{P}(X > x_0), \overline{\Lambda}^*(x_0))$. Let $X \in \mathbb{A}(X)$ be as defined by (5.2). Then $\mathbb{P}(X_j > x_j) = \mathbb{P}(C_j) < \overline{\Lambda}_j(x_j)$, implying that $\text{VaR}_{\Lambda_j}^+(X_j) \leq x_j$ for $j \in [m]$ and

$$\square \text{VaR}_{\Lambda_i}^+(X) \leq \sum_{i=1}^m \text{VaR}_{\Lambda_i}^+(X_i) \leq \sum_{i=1}^m x_i = x_0.$$

In view of Theorem 2, we conclude that X is an optimal allocation of X with $\text{VaR}_{\Lambda_j}^+(X_j) = x_j$ for $j \in [m]$.

(2) First, we show that no optimal allocation exists by contradiction. Assume on the contrary that there exists an optimal allocation $X \in \mathbb{A}_m(X)$. Denote $x_j = \text{VaR}_{\Lambda_j}^+(X_j)$ for $j \in [m]$. Then $\sum_{i=1}^m x_i = x_0$ and $\overline{\Lambda}^*(x_0) = \mathbb{P}(X > x_0) \leq \sum_{i=1}^m \mathbb{P}(X_i > x_i) \leq \sum_{i=1}^m \overline{\Lambda}_i(x_i)$. However, by Proposition 7, it follows that $\sum_{i=1}^m \overline{\Lambda}_i(x_i) < \overline{\Lambda}^*(x_0)$, a contradiction.

Next, we turn to constructing a sequence of asymptotically optimal allocations. We consider two subcases.

Subcase 1: Suppose that $\mathbb{P}(X > x_0 + \epsilon) < \mathbb{P}(X > x_0)$ for any $\epsilon > 0$. Let $(X_{1n}, \dots, X_{mn}) \in \mathbb{A}_m(X)$ be defined by (5.6). Then

$$\mathbb{P}\left(X_{jn} > x_{jn} + \frac{1}{mn}\right) < \overline{\Lambda}_j(x_{jn}), \quad j \in [m],$$

implying $\text{VaR}_{\Lambda_j}^+(X_{jn}) \leq x_{jn} + 1/(mn)$ for $j \in [m]$. Thus, $\sum_{i=1}^m \text{VaR}_{\Lambda_i}^+(X_{in}) \leq \sum_{i=1}^m x_{in} + 1/n = x_0 + 1/n$.

Subcase 2: Suppose that $\mathbb{P}(X > x_0 + \epsilon_0) = \mathbb{P}(X > x_0)$ for some $\epsilon_0 > 0$. Then $\overline{\Lambda}^*(x_0 + \epsilon) < \overline{\Lambda}^*(x_0)$ for any $\epsilon > 0$. Construct $(X_{1n}, \dots, X_{mn}) \in \mathbb{A}_m(X)$ as in (5.8). Similarly, we have $\sum_{i=1}^m \text{VaR}_{\Lambda_i}^+(X_{i,n}) \leq x_0 + 1/n$.

By Theorem 2, the desired statement follows by letting $n \rightarrow +\infty$. □

Proof of Theorem 8. We prove only part (3); the proofs of parts (1) and (2) are the same as those of Theorem 6. We consider two subcases.

Subcase 1: Suppose that $\overline{\Lambda}_i(x_i + \epsilon) < \overline{\Lambda}_i(x_i)$ for any $\epsilon > 0$ and $i \in K$. Let $(X_1, \dots, X_m) \in \mathbb{A}_m(X)$ be defined by (D.1). Then, $\mathbb{P}(X_i > x_i) = \overline{\Lambda}_i(x_i) < \overline{\Lambda}_i(x_i + \epsilon)$ for $i \in K$, and $\mathbb{P}(X_i > x_i) = \overline{\Lambda}_i(x_i)$ for $i \notin K$. Thus, $\text{VaR}_{\Lambda_j}^{k_j}(X_j) \leq x_j$ for all $j \in [m]$, implying $\sum_{i=1}^m \text{VaR}_{\Lambda_i}^{k_i}(X_i) = \text{VaR}_{\overline{\Lambda}^*}^+(X)$.

Subcase 2: Suppose that for any $(y_1, \dots, y_m) \in \mathbb{R}^m$ satisfying $\sum_{i=1}^m y_i = x_0$ and $\sum_{i=1}^m \overline{\Lambda}_i(y_i) = \overline{\Lambda}^*(x_0)$, there always exists some $\tau_0 > 0$ such that $\overline{\Lambda}_k(y_k) = \overline{\Lambda}_k(y_k + \tau_0)$ for some $k \in [m]$. By a similar argument to that in the proof of Theorem 6 (3), we can construct a sequence of allocations $(X_{1n}, \dots, X_{mn}) \in \mathbb{A}_m(X)$, satisfying $\sum_{i=1}^m \text{VaR}_{\Lambda_i}^{k_i}(X_{in}) \rightarrow x_0$. □

Proof of Theorem 9. The proof is similar to that of Theorem 7. □

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