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Invited Commentary

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Interpretational issues with the bifactor model: a commentary on 'Defining the *p*-Factor: An Empirical Test of Five Leading Theories' by Southward, Cheavens, and Coccaro

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Abstract

Southward, Cheavens, and Coccaro (2022, *Psychological Medicine*) conducted an ambitious investigation aimed at determining the nature of the general p factor of psychopathology by considering the correlation between the p factor and five candidate constructs. Generally, in this area of research, the bifactor model is preferred to the second order common factor model. In this commentary, we identify several interpretational issues concerning the bifactor model, which are based on a realistic psychometric view of latent variables. These issues may hamper the study of the nature of p factor model using the bifactor model.

Southward, Cheavens, and Coccaro (2022) have conducted an ambitious investigation aimed at determining the nature of the general p factor of psychopathology by considering the correlation between the *p* factor and five candidate constructs. The main result of Southward et al. (2022) is the correlation matrices of the *p* factor and the five candidate constructs in two versions of bifactor models (Caspi et al., 2014; Lahey et al., 2012; Lahey, Moore, Kaczkurki, & Zald, 2021), one with correlated and one with uncorrelated residual group factors (see their online Supplementary Tables S4b and S4c). The reported correlations, which may be interpreted as 'validity coefficients', are useful as they may help to determine the meaning of the p factor. While useful in some contexts, in our view the bifactor model poses problems of interpretation that require more attention in the literature. The aim of the present commentary is to discuss these by comparing the bifactor model and the second-order factor model. The outline of this commentary is as follows. We first present in conceptual terms the bifactor model and the second order factor model. We illustrate these using path diagrams in Figs 1 and 2. The path diagrams do not present the full models; we consider only four indicators of a given common factor, as this is sufficient given our present aims. Following the presentation of these models, we discuss two main issues: the issue of unidimensionality and the analyses of external dependent and independent covariates. We limit this commentary to indicators which are test items or individual symptoms. Sometimes subtests are used as indicators, but this is not relevant to the issues that we discuss. The Figures include parameters to which we refer in our commentary. We consistently mentioned the relevant Figure in referring to the parameters in the model.

Bifactor model and second order factor model

In the bifactor model, the *p* factor (denoted p_1 , see Fig. 1) features as the general factor on which all observed symptoms (or items) load directly (see Fig. 1 below). As such, the factor p_1 accounts for variance that is common to all symptoms, and provides a partial account of the symptom correlations. This is partial, because the sets of symptoms relating to dimensions of psychopathology (e.g. the set consisting of symptoms of depression) are expected to display residual correlations. Specific group factors (denoted s_k , k = 1, 2, ..., K, Fig. 1) are included in the model to account for the residual correlations. In this model, the *p*-factor p_1 is uncorrelated with the specific group factors s_k , but the specific group factors s_k may be correlated (see Southward et al., 2022).

The bifactor model represents one model among a range of candidate models that can be used to analyze the covariance structure of mental health problems (Borsboom et al., 2021; Chen, West, & Sousa, 2006; Markon, 2019; Yung, Thissen, & McLeod, 1999). A well-known alternative model is the second-order factor model, in which the *p* factor features as a second-order factor (denoted p_2 ; see Fig. 2). In this model, the symptoms load on correlated first (or lower) order factors (denoted f_{ko} k = 1,2,...,K, Fig. 2), which in turn load on the



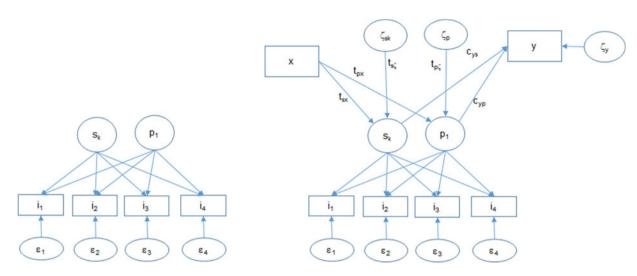


Fig. 1. Left: the bifactor model, with the general first order factor p_1 and the specific group factor s_k . Only one s_k is depicted (in practice there are three or more factors s_k). The observed symptoms or items, i_1 to i_4 , are regressed on p_1 and s_k . The variables ϵ are residuals in the regression of the i_1 to i_4 on f_k . Right: The observed x and y are external predictor and dependent variables, respectively. The variable ζ_p is the residual in the regression of p_1 on x, and ζ_{sk} is the residual in the regression of s_k on x. The variable ζ_y is the residual in the regression of s_k on x. The variable ζ_y is the residual in the regression of p_1 on x, and ζ_{sk} is the residual in the regression of s_k on s_k .

factor p_2 . So, while the observed symptoms are related directly to p_1 in the bifactor model, in the second-order model, the first order factors f_k mediate the relationship between p_2 and the observed symptoms. Thus, in this model the *p*-factor p_2 accounts for the correlations among the factors f_k , and, as such, may represent a common cause of the common factors f_k .

Many articles have been devoted to the interpretative merits of the bifactor model, specifically with respect to regression modeling of a dependent external variable. The merits lie in the fact that the decomposition of the dependent variable variance is simple, because the factors s_k and the factor p_1 are uncorrelated (as mentioned, the specific factors s_k may be correlated, as is the case in online Supplementary Table S4c in Southward et al., 2022). In the second-order factor model, the dependent variance decomposition is complicated by the fact that (1) the factors f_k and p_2 are correlated, and (2) given K + 1 predictors, one regression coefficient has to be fixed (to zero) to achieve model identification. Exactly which coefficient is fixed is generally arbitrary. The same applies with respect to external predictors of the common factors: all common factor in the bifactor model may be regressed on a given predictor, but in the second-order factor model, one regression coefficient has to be fixed.

While the advantage of the bifactor model in regression modeling may be relevant in a predictive context, it does not necessarily translate to an advantage in an explanatory context. In the latter context, the goal is to develop a model that represents the *p*-factor optimally from a theoretical and interpretative point of view. The emphasis on advantage of the bifactor model, coupled

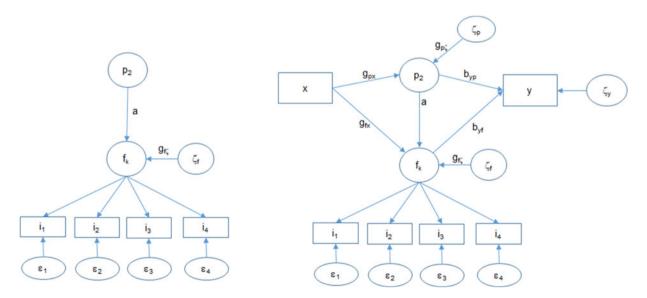


Fig. 2. Left: The second order factor model, with second order factor p_2 , and a single first order common factor f_k (in practice there are three or more factors f_k). The observed symptoms or items, i_1 to i_4 , are indicators of the latent variable f_k . The variables ϵ are residuals in the regression of the indictors in the factor f_k , and the variable ζ_f is the residual in the regression of f_k on p_2 . Right: The observed x and y are external predictor and dependent variables, respectively. The variable ζ_p is the residual in the regression of p_2 on x, ζ_y is the residual in the regression of y on p_2 and f_k , and ζ_f is the residual in the regression of f_k on x and p_2 .

with inconsequential (van Bork, Epskamp, Rhemtulla, Borsboom, & van der Maas, 2017) or potentially misleading (Greene et al., 2019) differences in model fit of the bifactor model and the second-order factor model, has overshadowed the psychometric and interpretative problems of the bifactor model (see also Achenbach, 2021; Pettersson, Larsson, & Lichtenstein, 2021). As mentioned above, we discuss two issues. First, the bifactor model stipulates that symptoms or items associated with a given latent variable (e.g. depression) are bi-dimensional. This inconsistent with the unidimensionality of measurement models, i.e. an important psychometric criterion, which is highly relevant to the interpretation of psychometric test scores (we return to this below). Second, the specification of uncorrelated s_k and p_1 rules out the investigation of role of the p factor as a mediator of the relationship between external dependent and independent covariates and dimensions of psychopathology. This hinders the study of the role of the *p* factor as the general cause of psychopathology.

Unidimensionality v. bi-dimensionality

We adopt the view of symptoms as reflective indicators, which are directly and causally dependent on the latent variable (Borsboom, Mellenbergh, & van Heerden, 2003). Given a single well defined dimension of psychopathology, the set of reflective indicators (i.e. symptoms) making up the psychometric test, designed to measure the dimension, should satisfy unidimensionality. Unidimensionality, which is typically established during the development of the test, e.g. by means of common factor modeling, is important for the interpretation of the test scores as proxies of the dimension of psychopathology, as represented by the common factor f_k . So, we assume that common factor f_k represents an interpretable dimension of psychopathology, which is measured using a unidimensional set of reflective indicators. We emphasize that we do not adhere to the position that psychometric theory is universally normative, in the sense that it requires all constructs to be unidimensional. Whether a construct is better represented in a unidimensional or multidimensional model is a substantive question that will depend on the topic studied. Our point of departure is rather that, if a construct in interpreted as being the common determinant of a set of item responses or subtest scores, unidimensionality is generally accepted as an important criterion. Given this point of departure, an interpretative problem of the bifactor model is that the indicator sets are specified to load on two uncorrelated common factors s_k and p_1 . The implied bi-dimensionality is inconsistent with the unidimensionality of the indicators, as indicators of f_k . This poses a problem of interpretation for both the factors s_k and the factor p_1 . The factors s_k are often interpreted in terms of the indicators on which they load, i.e. the s_k are interpreted simply as f_k . This is problematic for two reasons. First, the common factors f_k and s_k cannot reference the same latent variables, given the presence of the factor p_1 , which is uncorrelated with the specific factors s_k in the bifactor model. Second, because the nature of p_1 is unknown, the nature of s_{k} , the residual common factor in the presence of p_1 , is necessarily unknown as well. In addition, modeling the p-factor as a general first order factor, which accounts for symptom correlations, is problematic, because the explanandum is the correlations among the dimensions of psychopathology. Thus, the p-factor does not explain the correlations between the latent variables that potentially underlie mental disorders, but leaves them unmodeled, and thus essentially treats them as a nuisance. Note that the correlations among the symptoms themselves do not require explanation. The symptoms making up a given test are correlated because they depend on (indicate) the dimension (represented by f_k) that the test was designed to measure. The second order factor model avoids these problems. First, while the nature of the p factor (p_2) unknown, the first order factors (f_k) retain their unidimensionality and their interpretation, and thus are amenable to the standard psychometric analysis. Second, in contrast to the bifactor model, the p_2 factor actually does fulfill a productive explanatory role in the second order factor model, which is consistent with the issue at hand, because it provides an account of the correlations among the distinct and well defined dimensions of psychopathology.

Regression modeling

As mentioned, the interpretational advantage of the second order factor model comes with the limitation with respect to regression modeling of a predictor x and dependent variable y. To ensure identification of the regression model, in which we regress an external dependent variable y on the factors or in which we regress the common factors on an external predictor x (see Fig. 2), requires us to fix one of the K + 1 regression coefficients. Furthermore, the y variance decomposition in the second order factor model is considered hard to interpret in the second order factor model, because the decomposition is not orthogonal due to the correlation between the factors f_k and p_2 . The latter limitation is an issue, only if the cause of the correlation between the two predictors is unknown. In that case, the explained variance includes a component that cannot be attributed unambiguously to either predictor, because it depends in part on this correlation. However, in the second order factor model, the cause of the correlation between the predictors f_k and p_2 is represented explicitly by the parameter a in the regression of f_k on p_2 (Fig. 2), and therefore the variance decomposition does not pose any problem of interpretation. Specifically, in terms of Fig. 2, the explained y variance is due to the direct effect of p_2 on y (path b_{yp}), the indirect effect of p_2 on y (path a^*b_{yf}), and to ζ_f , the first order factor residual ζ_f (path $g_{f\zeta} * b_{vf}$). Note that this model allows us to investigate explicitly the role of the p factor (p_2) as the mediator of the relationship between f_k and the dependent y (i.e. the test of b_{vf} = 0). In the bifactor model, the dependent variable variance is explained by the paths c_{ys} and c_{yp} (Fig. 1). This model precludes a statement concerning the contribution of role of f_k to the explained variance, and cannot address the issue of mediation, where the prediction of y by f_k may be mediated by p_2 (i.e. $f_k \leftarrow p_2 \rightarrow y$).

The same reasoning applies with respect to the predictor x (Fig. 2). Again, the correlation between f_k and p_2 stemming from the parameter a, does not hinder the interpretation of components of explained variance. The explained variance of p_2 is due to the path g_{px} . The explained variance in f_k is decomposed into two interpretable parts: one due to the path $g_{px}^{*}a$, i.e. the part involving p_2 as a mediator, and a part due to g_{fx} . From the point of view of the standard mediation model (e.g. Baron & Kenny, 1986; Maxwell & Cole, 2007) demonstrating that the relations between the predictor x and the factors f_k is fully mediated by p_2 (i.e. the test of $g_{fx} = 0$) would likely advance our understanding of the *p* factor. For instance, genetic pleiotropy is well established in studies of psychopathology (see Grotzinger et al., 2022; Mallard et al., 2022). It would be of interest to determine the extent to which the genetic correlations among the factors f_k are attributable to genetic effects that are mediated by p_2 . In the

bifactor model, the mediation hypothesis, as outlined above in the second order factor model, does not apply, because it does not include the common factors f_k . Similarly, mediation among the predictor x, the factor p_1 and the specific factors s_k does not apply, because the factor p_1 and the specific factors s_k are uncorrelated. One can regress p_1 and s_k on the external predictor X, and view the insignificance of the latter (s_k on x, i.e. parameter $t_{xs} = 0$ in Fig. 1) as evidence of mediation by p_1 of the relation between the predictor x and the items. However, this is a different mediation hypothesis, which involves the items, not the common factors f_k . With respect to external predictor x, we note that the bifactor model appears to be restrictive, as it implicitly rules out the possibility that the external predictor x is associated with both s_k and p_1 . Specifically, a predictor common to s_k and p_1 (i.e. $t_{sx} \neq 0$ and $t_{px} \neq 0$ in Fig. 1) would give rise to a correlation between s_k and p_1 (due to $t_{sx} t_{px}$). This would seem to violate the assumption in the bifactor model that s_k and p_1 are uncorrelated.

Conclusion

In conclusion, the bifactor model poses problems of interpretation, because (1) it implies bi-dimensionality of the indicator set, which does not sit well with the psychometric and substantive ideal of unidimensionality; (2) it operationalizes the *p* factor as a direct source of correlation among items, while its explanandum is arguably the correlations among the factors f_{k} ; (3) the common factors s_k are hard to interpret, and cannot be interpreted as factors f_k ; (4) it does not allow for tests of the mediatory role of the *p* factor vis-à-vis the common factor f_k in the modeling of predictors (x) or dependent variables (y).

We emphasize that our commentary is based on the realistic interpretation of the first order factors (f_k) in the second order factor model (Borsboom et al., 2003). This has considerable explanatory import: the model encodes the scientific hypothesis that a general liability to develop psychopathology makes one more liable to any specific form of psychopathology, and therefore explains why the first order factors (f_k) are correlated. While this is a strong hypothesis that may very well be incorrect, it does represent a theory worthy of further investigation; a theory, moreover, that is consistent with what we would take the 'natural' reading of the p-factor as a general liability. As such, it would be infelicitous if a modeling choice for the bifactor model, which is based on pragmatic concerns regarding prediction and arguably inconsequential differences in model fit, were to be adopted as the default model without further deliberation. We therefore think that the choice between the bifactor model and the second order model deserves more attention than it currently enjoys.

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