

GENERATING STATIONARY RANDOM GRAPHS ON \mathbb{Z} WITH PRESCRIBED INDEPENDENT, IDENTICALLY DISTRIBUTED DEGREES

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Abstract

Let F be a probability distribution with support on the nonnegative integers. We describe two algorithms for generating a stationary random graph, with vertex set \mathbb{Z} , in which the degrees of the vertices are independent, identically distributed random variables with distribution F . Focus is on an algorithm generating a graph in which, initially, a random number of ‘stubs’ with distribution F is attached to each vertex. Each stub is then randomly assigned a direction (left or right) and the edge configuration obtained by pairing stubs pointing to each other, first exhausting all possible connections between nearest neighbors, then linking second-nearest neighbors, and so on. Under the assumption that F has finite mean, it is shown that this algorithm leads to a well-defined configuration, but that the expected length of the shortest edge attached to a given vertex is infinite. It is also shown that any stationary algorithm for pairing stubs with random, independent directions causes the total length of the edges attached to a given vertex to have infinite mean. Connections to the problem of constructing finitary isomorphisms between Bernoulli shifts are discussed.

Keywords: Random graph; degree distribution; stationary algorithm; random walk; finitary isomorphism

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1. Introduction

Recently there has been a lot of interest in the use of random graphs as models of various types of complex network. Several models have been formulated, aiming to capture essential features of the networks in question, such as degree distribution, diameter, and clustering; see, for instance, Dorogovtsev and Mendes (2003) and Bollobás and Riordan (2002) for surveys. Regarding the vertex degree, power-law distributions have been identified in many of the real-world applications, implying that the ordinary Erdős–Rényi graph, introduced in Erdős and Rényi (1959) and giving Poisson-distributed degrees in the limit of large graph size, is not suitable as a model. This has given rise to a number of algorithms for generating graphs with an arbitrary prescribed degree distribution. The most studied one is the so-called configuration model, in which each vertex is assigned a random number of stubs which are then joined pairwise, completely randomly, to form edges. The asymptotic behavior of this model has been studied by Molloy and Reed (1995), (1998), Newman *et al.* (2001), and van der Hofstad *et al.* (2005), among others. Also, Britton *et al.* (2005) considered a modification of the model in

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which multiple edges and self-loops (that is, edges beginning and ending at the same vertex) are forbidden, giving simple graphs as a final result. A different model for generating random graphs with prescribed degrees was studied by Chung and Lu (2002a), (2002b).

A natural generalization of the problem of generating random graphs with prescribed degree distributions is to consider spatial versions of the same problem, where geometric aspects play a role. More precisely, given a probability distribution F with support on the nonnegative integers and a set of vertices with some kind of spatial structure, how should an edge configuration on this vertex set with degree distribution F be generated? The answer to this question clearly depends on the nature of the spatial structure and also on the desired properties of the resulting configuration.

Here we consider the problem of generating a *stationary* random graph with vertex set \mathbb{Z} and with vertex degrees that are independent and identically distributed (i.i.d.) with distribution F . We recall that a graph on \mathbb{Z} is said to be stationary if the distribution of the edge configuration restricted to any finite subset of \mathbb{Z} is translation invariant. Hence, we seek a stationary algorithm producing edges among the vertices of \mathbb{Z} in such a way that the vertex degrees become i.i.d. random variables with distribution F . We have two suggestions for how to do this.

1.1. Stepwise pairing with random directions

Our first suggestion is the one we will spend most of the paper on. The algorithm runs as follows.

1. Independently attach a random number of stubs to each vertex according to the distribution F .
2. Randomly assign a direction (left or right) to each stub, turning it into an arrow.
3. Join stepwise the arrows pointing to each other, first exhausting all possible connections between nearest-neighbor vertices, then looking at second-nearest neighbors, and so on, until all arrows are connected.

This model will be referred to as the stepwise pairing algorithm with random directions (or the SPRD algorithm). In Section 2 we show that, with probability 1, the SPRD algorithm leads to a well-defined configuration, that is, one in which the number of steps required for a right-arrow or a left-arrow to respectively find a left-arrow or a right-arrow to connect to is almost surely finite. However, the shortest edge of a vertex already turns out to have infinite mean; see Section 4. Basically this follows from properties of a random walk structure that arises in the analysis of the model. This analysis is complicated by the fact that the increments of the random walk are not independent, making standard results inapplicable. In Section 5 we prove that if we insist on the directions of the edges of a given vertex being completely random, and also independent of the configuration at all other vertices, then we cannot achieve finite mean for the *total* edge length per vertex in a stationary way. Note, however, that by dropping the requirement that the directions of the edges should be assigned randomly and independently, it is in some cases possible to design algorithms that give finite mean for the edges; see Examples 5.1 and 5.2.

Readers familiar with Bernoulli isomorphisms might have observed that our pairing rule is similar to the pairing rule used by Meshalkin (1959) to construct an isomorphism between two specific Bernoulli shifts with equal entropy. In fact there are connections, which we will briefly explore in Section 3. We also mention the paper by Holroyd and Peres (2005), who dealt with stationary matching in a slightly different setup.

1.2. Annihilating random walk

There is a completely different way of generating a stationary graph with the required properties, by making use of random walks, which we mention for completeness. It can be described in three steps, as follows.

1. Independently attach a random number of stubs to each vertex according to the distribution F .
2. To each stub associate a particle at the same position in \mathbb{Z} and let each particle start a continuous-time random walk on \mathbb{Z} , independent of all the others.
3. Whenever two particles – which start at different locations – meet, draw an edge between the corresponding stubs and remove the particles from the system.

It is not hard to see, and is well known (see, for instance, Arratia (1981)), that this leads to a limiting configuration in which all stubs are connected. However, we will not be concerned with this type of pairing in this paper. See Mattera (2003) for other connections between annihilating random walks and graphs.

2. Definition of the SPRD algorithm

Let us now describe the SPRD algorithm in more detail. First, independently associate to each vertex $i \in \mathbb{Z}$ a random degree, D_i , with distribution F . Think of this as vertex i having D_i ‘stubs’ attached to it. Now turn the stubs into arrows by randomly associating a direction to each of them. More precisely, with probability p a stub is pointed to the right and with probability $1 - p$ it is pointed to the left. Write R_i and L_i for the numbers of right-arrows and left-arrows of vertex i and label the arrows $\{r_{i,j}\}_{j=1}^{R_i}$ and $\{l_{i,j}\}_{j=1}^{L_i}$, respectively. This gives a configuration in which each vertex i has two ordered sets of arrows, $\{r_{i,j}\}_j$ and $\{l_{i,j}\}_j$, associated with it. These arrows will now be matched pairwise – a pair always consisting of one right-arrow and one left-arrow – to create edges between the vertices. The matching is done stepwise, as follows.

1. First consider all pairs of nearest-neighbor vertices, i and $i + 1$, and create $\min\{R_i, L_{i+1}\}$ edges between vertex i and vertex $i + 1$ by joining the arrows $r_{i,j}$ and $l_{i+1,j}$, $j = 1, \dots, \min\{R_i, L_{i+1}\}$.
2. Next consider all pairs of second-nearest-neighbor vertices, i and $i + 2$. If, after step 1, there is at least one unconnected right-arrow at vertex i and at least one unconnected left-arrow at vertex $i + 2$, then we create edge(s) between the vertices i and $i + 2$ by performing all possible connections, always connecting an arrow $r_{i,j}$ or $l_{i+2,j}$ before an arrow $r_{i,j+1}$ or, respectively, $l_{i+2,j+1}$.
- ⋮
- n . In step n , we consider all pairs of vertices, i and $i + n$, at distance n from each other, and connect arrows that remain after the previous steps, never using an arrow $r_{i,j}$ or $l_{i+n,j}$ before an arrow $r_{i,j-1}$ or, respectively, $l_{i+n,j-1}$.
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The above procedure is clearly stationary, but we have yet to show that it leads to a well-defined graph. To this end, define the length of an edge in the resulting configuration to be the distance between its endpoints. In what follows, we will consider only the vertex at the origin.

Write $N_j^{(r)}$ for the length of the edge created by right-arrow number j at the origin, i.e. $r_{0,j}$, and set $N_j^{(r)} = \infty$ if $r_{0,j}$ is never connected. Also, define $N_j^{(l)}$, $j \geq 1$, analogously for the left-arrows. Write

$$N^{(r)} = \max_{1 \leq j \leq R_0} \{N_j^{(r)}\} \quad \text{and} \quad N^{(l)} = \max_{1 \leq j \leq L_0} \{N_j^{(l)}\},$$

and define $N = \max\{N^{(r)}, N^{(l)}\}$.

We first show that the algorithm does not work for $p \neq \frac{1}{2}$. Below, P_p denotes the probability measure associated with the SPRD algorithm when a stub is pointed to the right with probability p .

Proposition 2.1. *If $p \neq \frac{1}{2}$ then $P_p(N = \infty) > 0$.*

Proof. To see that $P_p(N_1^{(r)} = \infty) > 0$ for $p > \frac{1}{2}$, fix $p > \frac{1}{2}$, let $\Delta_i = L_i - R_i$, and define

$$S'_1 = L_1, \quad S'_n = \sum_{i=1}^{n-1} \Delta_i + L_n, \quad n \geq 2.$$

Clearly the first right-arrow at the origin is connected as soon as S'_n takes a positive value; hence, it suffices to show that $P_p(S'_n \leq 0 \text{ for all } n) > 0$. This, however, follows easily by noting that the law of large numbers implies that $S'_n \rightarrow \infty$ almost surely.

Having discarded asymmetric versions of the algorithm, let us move on to the symmetric case, in which the prospects of success should be better. Indeed, the following proposition guarantees that for $p = \frac{1}{2}$ no arrow has to wait infinitely long before it finds another to connect to.

Theorem 2.1. *We have $P_{1/2}(N < \infty) = 1$.*

Proof. For ease of notation, write $P_{1/2} = P$. First note that, by symmetry, it suffices to show that $P(N^{(r)} < \infty) = 1$. Also, by the definition of the algorithm, the arrows $\{r_{0,j}\}$ are used in (stepwise) chronological order, implying that $N_j^{(r)} \leq N_{j+1}^{(r)}$. It follows that $N^{(r)} = N_{R_0}^{(r)}$, and, since $R_0 < \infty$ almost surely, we have finished if we can show that

$$P(N_j^{(r)} < \infty) = 1 \quad \text{for all } j.$$

To do this, we first consider the case $j = 1$ and show that $P(N_1^{(r)} < \infty) = 1$, i.e. that the length of the edge created by the first right-arrow at the origin, $r_{0,1}$, is almost surely finite. This is done by dominating the length of the edge by the time at which a recurrent random walk takes a positive value for the first time. To be more specific, define $\Delta_i = L_i - R_i$ (as above) and write $S_n = \sum_{i=1}^n \Delta_i$. The variables $\{\Delta_i\}$ are i.i.d. and symmetric, implying that $\eta := \inf\{n : S_n > 0\}$ is finite with probability 1. Now note that as soon as $S_n > 0$ we know that the arrow $r_{0,1}$ must have found a left-arrow to connect to. Indeed, if $S_n > 0$ we also have $S_n + R_n > 0$, and the fact that $S_n + R_n > 0$ means that the total number of left-arrows on the vertices $1, \dots, n$ is strictly larger than the total number of right-arrows on the vertices $1, \dots, n - 1$, implying that at some vertex $1, \dots, n$ there must be a left-arrow for $r_{0,1}$ to connect to. It follows that $N_1^{(r)} \leq \eta$, and we have finished.

Now assume in an inductive fashion (on j) that $P(N_j^{(r)} < \infty) = 1$, and suppose that $P(N_{j+1}^{(r)} = \infty) > 0$. Write $\Psi = \{\Psi_i\} = \{(L_i, R_i)\}$ for the random configuration of arrows at the vertices and, for configurations with $N_j^{(r)} < \infty$, introduce a coupled configuration, $\hat{\Psi} = \{\hat{\Psi}_i\}$,

that is identical to Ψ except that the directions of the stubs at the vertex $N_j^{(r)}$ are generated independently. Let $\hat{N}_j^{(r)}$ be the length of the edge formed by $r_{0,j}$ in $\hat{\Psi}$, and define

$$A_j = \{N_{j+1}^{(r)} = \infty\} \cap \{\hat{L}_{N_j^{(r)}} = 0\}.$$

Note that on the event A_j we have $\hat{N}_j^{(r)} = \infty$. Indeed $r_{0,j}$ cannot connect before vertex $N_j^{(r)}$ in $\hat{\Psi}$, since left-arrows have been removed at $N_j^{(r)}$ without any new right-arrows being added. Furthermore, if $r_{0,j}$ were connected to a left-arrow at vertex $m \geq N_j^{(r)}$ in $\hat{\Psi}$, it would imply that $r_{0,j+1}$ were connected at the latest to m in Ψ , which contradicts the fact that $N_{j+1}^{(r)} = \infty$. Hence, we have $P(\hat{N}_j^{(r)} = \infty) \geq P(A_j)$. It follows from the assumption that $P(A_j) > 0$, so since $N_j^{(r)}$ and $\hat{N}_j^{(r)}$ clearly have the same distribution, we have shown that $P(N_j^{(r)} = \infty) > 0$. However, this is a contradiction, and by induction on j we conclude that $P(N_j^{(r)} < \infty) = 1$ for all j , as desired.

3. Connections to Bernoulli isomorphisms

Consider a stochastic process, X , indexed in \mathbb{Z} , with i.i.d. marginals taking values in $1, 2, \dots, s$ with respective probabilities p_1, \dots, p_s . The process X is often called a Bernoulli shift, and is identified with the vector (p_1, \dots, p_s) . Next, consider another such process, Y , taking values in $1, 2, \dots, t$ with respective probabilities q_1, \dots, q_t . We write $S_X = \{1, 2, \dots, s\}^{\mathbb{Z}}$ and $S_Y = \{1, 2, \dots, t\}^{\mathbb{Z}}$. Loosely speaking, X and Y are said to be *isomorphic* if there exists a bijective pairing of almost all realizations of X and Y , such that the pairing commutes with the shift operators on S_X and S_Y .

Meshalkin (1959) was one of the first to explicitly identify such a coding between two particular Bernoulli shifts, namely between $(\frac{1}{4}, \frac{1}{4}, \frac{1}{4}, \frac{1}{4})$ and $(\frac{1}{2}, \frac{1}{8}, \frac{1}{8}, \frac{1}{8}, \frac{1}{8})$. His coding corresponds to our algorithm in the case in which each vertex has at most one edge associated to it as follows. Associate to each edge a random label, a or b , independently and with equal probability. This leads to four equally likely edge labels, namely (l, a) , (l, b) , (r, a) , and (r, b) , where l and r respectively denote an edge pointing to the left and an edge pointing to the right. The coding is now defined so that whenever we see (r, a) or (r, b) we write an r , and whenever we see an l we write (l, x, y) , where x and y are the symbols corresponding to the edge that is formed with the unique stub at that position. It is not hard to see that this codes the original four symbols into five new symbols with respective probabilities $\frac{1}{2}, \frac{1}{8}, \frac{1}{8}, \frac{1}{8},$ and $\frac{1}{8}$, and that this coding is invertible. Furthermore, the coding is *finitary*, that is, we have to look only a (random) finite distance in both directions to see the symbol that should be written in the coding. Indeed, once we have identified the stub that connects to the stub of current interest, we can write down the correct symbol.

This idea can be extended to apply to a general degree distribution, F , with bounded support in our pairing algorithm; that is, every F with bounded support leads to an isomorphism between two particular Bernoulli shifts, as the reader can easily verify. Certain results of coding between Bernoulli shifts then have corollaries for our algorithm. We mention the well-known fact (see, for instance, Parry (1979) or Schmidt (1984)) that in any nontrivial situation, the expected distance that has to be explored in a finitary coding between two Bernoulli shifts with equal entropy has infinite expectation. From this it follows that the *longest* edge at a given vertex in the SPRD algorithm has infinite expected length. Below we strengthen this result to apply to the shortest edge, which does not have an interpretation in the coding setup described here.

4. The mean length of the shortest edge

For the remainder of the paper we consider only the symmetric SPRD algorithm, which, by Theorem 2.1, leads to well-defined configurations. The next task is to look at the expected length of the edges. For distributions with finite mean, we will prove the following theorem.

Theorem 4.1. *If F has finite mean then both $E[N_1^{(r)}]$ and $E[N_1^{(l)}]$ are infinite.*

Define $X_i = L_i - R_{i-1}$, the difference between the numbers of left-arrows and right-arrows at two neighboring vertices. Let $S_n^{(m)} = \sum_{i=m+1}^{m+n} X_i$, and write $\tau_{\uparrow}^{(m,x)}$ to denote the first time the process $S_n^{(m)}$ rises above the level x :

$$\tau_{\uparrow}^{(m,x)} = \min\{n : S_n^{(m)} \geq x\}.$$

Clearly, to prove Theorem 4.1 it suffices to show that $E[N_1^{(r)}] = \infty$. To see why this should be the case, note that if $L_1 = 0$ (which happens with positive probability), then the first right-arrow at the origin is connected when $S_n^{(1)}$ takes a value larger than or equal to 0, that is, when $N_1^{(r)} = \tau_{\uparrow}^{(1,0)} + 1$. Were $S_n^{(1)}$ to have independent increments, it follows from standard random walk theory that $\tau_{\uparrow}^{(1,0)}$ would have infinite mean. However, X_i and X_{i+1} are not independent, since information about the arrow configuration at vertex i is used for both variables.

Let μ denote the mean of F . The following lemma will play a key role in the proof of Theorem 4.1.

Lemma 4.1. *For all $i \in \mathbb{Z}$, we have $E[\tau_{\uparrow}^{(i,2\mu)}] = \infty$.*

Proof. By stationarity, it suffices to show that $E[\tau_{\uparrow}^{(0,2\mu)}] = \infty$. Assume, on the contrary, that $E[\tau_{\uparrow}^{(0,2\mu)}] < \infty$, and define $\tau_{\downarrow}^{(m,x)}$ to be the first time the process $S_n^{(m)}$ falls below the level x :

$$\tau_{\downarrow}^{(m,x)} = \min\{n : S_n^{(m)} \leq x\}.$$

Note that, by symmetry, we have $E[\tau_{\downarrow}^{(0,-2\mu)}] = E[\tau_{\uparrow}^{(0,2\mu)}]$. The idea of the proof is to use the finite-mean assumption to create a linear negative drift for the process $S_n^{(0)}$. By symmetry, $S_n^{(0)}$ must then also have the same positive drift, and to maintain both these drifts it is forced to oscillate more and more vigorously between large positive values and large negative values, something which it will not be able to do in the long run. To turn this heuristic argument into a proof, introduce an i.i.d. sequence, $\{\Delta\tau_j\}$, with mean $E[\tau_{\downarrow}^{(0,-2\mu)}] + 1$ by defining

$$\Delta\tau_0 = 0, \quad \Delta\tau_j = \min\left\{n : S_n^{\left(\sum_{i=0}^{j-1} \Delta\tau_i\right)} \leq -2\mu\right\} + 1, \quad j \geq 1.$$

Here the ‘+1’ is added to ensure independence, noting that X_i and X_k are independent if $|i - k| \geq 2$. This sequence gives rise to a renewal process with time increments $\Delta\tau_j$ and events referred to as *down-transitions* occurring at the time points $\tau_i = \sum_{j=1}^i \Delta\tau_j$. Write M_n for the number of down-transitions in the time interval $[0, n]$ and note that, by the renewal theorem, we have

$$\frac{M_n}{n} \rightarrow \frac{1}{E[\tau^{(0,-2\mu)}] + 1} \quad \text{almost surely, as } n \rightarrow \infty.$$

Hence, defining $2c = (E[\tau^{(0,-2\mu)}] + 1)^{-1}$ and $E_m = \{M_n > nc \text{ for all } n \geq m\}$, it follows that

$$P(E_m) \rightarrow 1 \quad \text{as } m \rightarrow \infty. \tag{4.1}$$

At point τ_d of the d th down-transition we have

$$S_{\tau_d}^{(0)} \leq -2\mu d + \sum_{i=1}^d X_{\tau_i},$$

where $X_{\tau_i} = L_{\tau_i} - R_{\tau_i-1} \leq L_{\tau_i}$. The degree of a vertex $\tau_i - 1$ is atypical, since it is defined as a first passage point for the process $S_n^{(\tau_i-1)}$. However, the vertex τ_i has the unconditional degree distribution F , meaning that $E[L_{\tau_i}] = \mu/2$. Also, since $|\tau_i - \tau_{i-1}| \geq 2$, the variables L_{τ_i} are independent. By combining these facts, we find from the strong law of large numbers that

$$\frac{1}{d} \sum_{i=1}^d X_{\tau_i} \leq \frac{1}{d} \sum_{i=1}^n L_{\tau_i} \rightarrow \frac{\mu}{2} \text{ almost surely, as } n \rightarrow \infty.$$

Defining

$$F_m = \left\{ \sum_{i=1}^{\lfloor nc \rfloor} X_{\tau_i} \leq \lfloor nc \rfloor \mu \text{ for all } n \geq m \right\},$$

where $\lfloor r \rfloor \in \mathbb{R}$ denotes the largest integer smaller than or equal to r , it follows that

$$P(F_m) \rightarrow 1 \text{ as } m \rightarrow \infty. \tag{4.2}$$

Here note that if the sequence $\{\tau_i\}$ were defined in terms of *up-transitions* instead of down-transitions, then estimation of the value of the process after some large number of transitions would require a *lower* bound for the sum of the auxiliary steps X_{τ_i} . This, however, would cause a problem, since, as mentioned above, the ‘negative part’ of each variable X_{τ_i} concerns the arrow configuration at a first passage vertex, which is presumably difficult to control. Hence, we are in the peculiar situation of being able to prove a statement for down-transitions but not (directly) for up-transitions, in an otherwise completely symmetrical situation.

Next, divide \mathbb{Z}^+ into intervals, $\mathcal{I}_k = \{i : kl \leq i < (k+1)l\}$, $k \geq 0$, of length l , and write B_k for the event that the interval \mathcal{I}_k contains a down-transition in the renewal process $\{\tau_i\}$. Clearly, by choosing l to be large, we can ensure that $P(B_k) \geq 0.99$ for all k . Define Y_k to be the sum of the degrees of all vertices in \mathcal{I}_k , that is, $Y_k = \sum_{i=kl}^{(k+1)l-1} D_i$. The distribution of Y_k does not depend on k , and $Y_k < \infty$ almost surely, implying that

$$P(Y_k \geq 2\mu \lfloor klc \rfloor) \rightarrow 0 \text{ as } k \rightarrow \infty. \tag{4.3}$$

By (4.1), (4.2), and (4.3), if we choose k to be sufficiently large, we have

$$P(E_{kl}) \geq 0.99, \tag{4.4}$$

$$P(F_{kl}) \geq 0.99, \tag{4.5}$$

$$P(Y_k \geq 2\mu \lfloor klc \rfloor) \leq 0.5. \tag{4.6}$$

Fix such a k and define

$$D_k^- = \{\text{there exists an } n \in \mathcal{I}_k \text{ such that } S_n^{(0)} \leq -\mu \lfloor klc \rfloor\},$$

$$D_k^+ = \{\text{there exists an } n \in \mathcal{I}_k \text{ such that } S_n^{(0)} \geq \mu \lfloor klc \rfloor\}.$$

Now observe that $B_k \cap E_{kl} \cap F_{kl} \subset D_k^-$; indeed, occurrence of the event E_{kl} implies that $m \geq \lfloor klc \rfloor$ down-transitions have occurred in $[0, kl]$, and occurrence of the event F_{kl} implies that $\sum_{i=1}^{m+1} X_{\tau_i} \leq \mu(m + 1)$. Hence, at the point τ_{m+1} of the next down-transition, we have

$$\begin{aligned} S_{\tau_{m+1}}^{(0)} &\leq -2\mu(m + 1) + \sum_{i=1}^{m+1} X_{\tau_i} \\ &\leq -2\mu \lfloor klc \rfloor + \mu \lfloor klc \rfloor \\ &= -\mu \lfloor klc \rfloor. \end{aligned}$$

However, this means that D_k^- must occur, since, on the event B_k , at least one down-transition takes place in \mathcal{I}_k , that is, $\tau_{m+1} \in \mathcal{I}_k$. It follows that

$$\begin{aligned} P(D_k^-) &\geq P(B_k \cap E_{kl} \cap F_{kl}) \\ &\geq 1 - P(B_k^c) - P(E_{kl}^c) - P(F_{kl}^c) \\ &\geq 0.97. \end{aligned}$$

By symmetry, we have $P(D_k^+) = P(D_k^-)$ and, hence, $P(D_k^+ \cap D_k^-) \geq 0.94$. Now note that, on the event $D_k^+ \cap D_k^-$, both a state above the level $\mu \lfloor klc \rfloor$ and a state below the level $-\mu \lfloor klc \rfloor$ are visited in the interval \mathcal{I}_k and, thus, $Y_k \geq 2\mu \lfloor klc \rfloor$. Therefore,

$$P(Y_k \geq 2\mu \lfloor klc \rfloor) \geq P(D_k^+ \cap D_k^-) \geq 0.94.$$

However, this contradicts (4.6) in the choice of k . Hence, the assumption that $E[\tau_{\uparrow}^{(0,2\mu)}] < \infty$ must fail and the proof of the lemma is complete.

Proof of Theorem 4.1. By symmetry, it suffices to show that $E[N_1^{(r)}] = \infty$. To do this, as before write $\Psi = \{\Psi_i\} = \{(L_i, R_i)\}$ for the random configuration of arrows at the vertices and choose k to be sufficiently large that $P(\sum_{i=1}^k D_i \geq 2\mu) > 0$. Introduce a coupled configuration, $\hat{\Psi} = \{\hat{\Psi}_i\}$, with the same degrees at all vertices and $\hat{\Psi}_i = \Psi_i$ for $i \notin \{1, \dots, k + 1\}$, but in which the directions of the arrows at the vertices $1, \dots, k + 1$ are generated independently. Define

$$A = \left\{ \sum_{i=1}^k D_i \geq 2\mu \right\} \cap \{ \hat{L}_i = 0 \text{ for all } i = 1, \dots, k + 1 \}$$

and let $\hat{N}_1^{(r)}$ be the length of the edge formed by $r_{0,1}$ in $\hat{\Psi}$. We then have

$$E[\hat{N}_1^{(r)}] \geq E[\hat{N}_1^{(r)} \mid A] P(A).$$

Since clearly $\hat{N}_1^{(r)}$ has the same distribution as $N_1^{(r)}$ and $P(A) > 0$, we have finished if we can show that $E[\hat{N}_1^{(r)} \mid A] = \infty$. To this end, let $\hat{S}_n^{(m)}$ be defined in the same way as $S_n^{(m)}$, but based on the coupled configuration $\hat{\Psi}$, and write $\hat{t}_{\uparrow}^{(m,x)} = \inf\{n : \hat{S}_n^{(m)} \geq x\}$. On A , there are in total at least 2μ right-arrows attached to the vertices $1, \dots, k$ and no left-arrows at all on the vertices $1, \dots, k + 1$. Thus, a right-arrow at the origin cannot be connected until the process $\hat{S}_n^{(k+1)}$ takes a value larger than 2μ . It follows that

$$E[\hat{N}_1^{(r)} \mid A] \geq (k + 1) + E[\hat{t}_{\uparrow}^{(k+1,2\mu)} \mid A].$$

The effect that the conditioning on A has on $\hat{\tau}_\uparrow^{(k+1,2\mu)}$ is that the first term in the unconditional sum $\hat{S}_n^{(k+1)}$ is replaced by $L_{k+2} - D_{k+1}$, since on A all D_{k+1} stubs at vertex $k + 1$ point to the right. This means that, conditional on A , the passage time $\hat{\tau}_\uparrow^{(k+1,2\mu)}$ is stochastically larger than in the unconditional case, implying that

$$E[\hat{\tau}_\uparrow^{(k+1,2\mu)} \mid A] \geq E[\hat{\tau}_\uparrow^{(k+1,2\mu)}].$$

Hence,

$$E[\hat{N}_1^{(r)} \mid A] \geq (k + 1) + E[\hat{\tau}_\uparrow^{(k+1,2\mu)}].$$

Since $\hat{\tau}_\uparrow^{(k+1,2\mu)}$ has the same distribution as $\tau_\uparrow^{(k+1,2\mu)}$, it follows from Lemma 4.1 that

$$E[\hat{\tau}_\uparrow^{(k+1,2\mu)}] = \infty,$$

and the proof of the theorem is complete.

5. Finite mean is impossible

We are now at the point of having formulated a stationary algorithm that takes a discrete distribution, F , as input and produces a stationary random edge configuration on \mathbb{Z} with i.i.d. vertex degrees with distribution F . Provided that F has finite mean, all connections are almost surely finite, while the expected length of the connections is infinite. The obvious question is: can we do better? The following simple examples show that if we no longer assign i.i.d. directions to the stubs, then, for certain distributions F , we indeed can.

Example 5.1. Write f_j for the probability that a given vertex has degree j , fix $n \in \mathbb{N}$, and let F be defined as follows:

$$f_j = (n + 1)^{-1}, \quad j \in \{0, 2, 4, \dots, 2n\}, \quad f_j = 0, \quad j \notin \{0, 2, 4, \dots, 2n\}.$$

A configuration with this degree distribution and connections with finite mean is generated by proceeding in the same way as in the SPRD algorithm, except that the directions of the stubs are not assigned randomly but rather according to the deterministic rule that a vertex with degree $2k$ is equipped with exactly k arrows pointing in each direction. To see this, note that, assuming that the origin has degree d , all right-arrows at the origin will be connected as soon as a vertex $i \geq 1$ with degree larger than d is encountered. The expected distance to a vertex with degree exactly d is f_d^{-1} and, removing the conditioning on d , it follows that the expected length of the longest connection to the right is bounded by n . By symmetry, the expected maximal length to the left is also bounded by n .

Example 5.2. Let $F = \delta_1$, that is, every vertex is to have exactly one edge connected to it. To generate such a configuration, attach one stub to each vertex and then imagine that a coin is flipped. If the coin comes up heads then the stubs at the odd vertices are pointed to the right and the stubs at the even vertices are pointed to the left, and vice versa if it comes up tails. The arrows are then connected according to the stepwise pairing algorithm. It is easy to see that with this procedure (which is clearly stationary), all connections will end up having length 1.

Recall that, in the SPRD algorithm, the direction of each stub is chosen randomly to obey (4.4) and independently to obey (4.5). This gives rise to a random walk-type structure which is recurrent but has infinite mean. In Example 5.1 the directions of the edges are not random,

and in Example 5.2 they are not independent. This invalidates the random walk arguments and makes it possible to obtain configurations in which the connections have finite mean. Thus, for some distributions F it is indeed possible to outdo the SPRD algorithm by being clever when assigning the directions of the stubs. However, we conjecture that if the direction is assigned independently to each stub, then it is impossible to formulate a rule for connecting right-arrows to left-arrows in such a way that the expected length of the resulting edges becomes finite. A weaker formulation of this conjecture is proved in Theorem 5.1, below.

Let Ψ be a random configuration of arrows on \mathbb{Z} generated in the first SPRD algorithm: first a random number of stubs with distribution F is attached to each vertex and then each stub is randomly assigned the direction left or right. An algorithm, \mathcal{A} , for connecting the arrows in Ψ will be called a *pairing rule* if with probability 1 each left-arrow is connected to exactly one right-arrow and each right-arrow is connected to exactly one left-arrow. Furthermore, \mathcal{A} is said to be stationary if the resulting joint edge length distributions are translation invariant. For a given pairing rule, \mathcal{A} , write $T_{\mathcal{A}}$ and $N_{\mathcal{A}}$ for the total length of all edges connected to the origin and the length of the longest edge connected to the origin, respectively.

Theorem 5.1. *If F has finite mean, then for all stationary pairing rules \mathcal{A} we have $E[T_{\mathcal{A}}] = \infty$. If, in addition, F has bounded support, then $E[N_{\mathcal{A}}] = \infty$.*

The proof of this theorem is based on a combinatorial lemma involving the concept of *nested* graphs. To define this concept, consider a given edge configuration, $\{(i, j)\}_{i, j \in \mathbb{Z}}$, on \mathbb{Z} . Two edges (i, j) and (i', j') are said to *cross* each other if $i < i' < j < j'$ or $i' < i < j' < j$, and the configuration $\{(i, j)\}_{i, j \in \mathbb{Z}}$ is said to be *nested* if it does not contain any crossed edges. An important observation is that, for a given configuration, ψ , of arrows on \mathbb{Z} , there is a unique nested edge configuration, to be denoted by \mathcal{N}_{ψ} , which is obtained by the stepwise pairing algorithm. Indeed, to avoid crossed edges we are forced to perform all possible connections between vertices at distances $n = 1, 2, \dots$, starting with $n = 1$; conversely, successively performing all possible connections between vertices at distance n , with n increasing, can never in any step create crossed edges, since this would mean that a possible connection in a previous step was missed.

To formulate the aforementioned lemma, write Γ for the set of all arrow configurations ψ on \mathbb{Z} for which all edges in \mathcal{N}_{ψ} are finite. Pick $\psi \in \Gamma$ and, for an edge $e \in \mathcal{N}_{\psi}$, let $\psi_e^{(r)}$ and $\psi_e^{(l)}$ respectively be the sets of right-arrows and left-arrows in ψ that are used to form the edge e and the edges ‘under’ e in \mathcal{N}_{ψ} . More precisely, if e is made up of the arrows $r_{i, j}$ and $l_{i+n, j'}$, then $\psi_e^{(r)}$ consists of the set of arrows $\{r_{i, k}\}_{k=1}^j$ together with all right-arrows at the vertices $i + 1, \dots, i + n - 1$, and $\psi_e^{(l)}$ consists of $\{l_{i+n, k}\}_{k=1}^{j'}$ and all left-arrows at the vertices $i + 1, \dots, i + n - 1$. Write $t_e(\mathcal{N}_{\psi})$ for the total length of all edges ‘under’ e in \mathcal{N}_{ψ} .

Next, let \mathcal{E}_{ψ} be some edge configuration based on the arrow configuration ψ . Call an edge in \mathcal{E}_{ψ} a $\psi_e^{(r)}$ -edge if it contains an arrow belonging to the set $\psi_e^{(r)}$ and let $t_e^{(r)}(\mathcal{E}_{\psi})$ denote the total length of all $\psi_e^{(r)}$ -edges in the configuration \mathcal{E}_{ψ} . Define $t_e^{(l)}(\mathcal{E}_{\psi})$ analogously. The lemma then reads as follows.

Lemma 5.1. *For all $\psi \in \Gamma$, all configurations \mathcal{E}_{ψ} based on ψ , and all $e \in \mathcal{N}_{\psi}$, we have $t_e(\mathcal{N}_{\psi}) \leq t_e^{(r)}(\mathcal{E}_{\psi})$ and $t_e(\mathcal{N}_{\psi}) \leq t_e^{(l)}(\mathcal{E}_{\psi})$.*

Proof. Fix an arrow configuration $\psi \in \Gamma$, an edge $e \in \mathcal{N}_{\psi}$, and an edge configuration, \mathcal{E}_{ψ} , based on ψ . Define $w_k^{(r)}$ to be the number of $\psi_e^{(r)}$ -edges in \mathcal{E}_{ψ} that cross the interval $[k - 1, k]$. More precisely, $w_k^{(r)}$ is the number of edges in \mathcal{E}_{ψ} that have their left-hand endpoints at a vertex $l \leq k - 1$, their right-hand endpoints at $l' \geq k$, and are created by right-arrows that belong to

$\psi_e^{(r)}$. Let $\tilde{w}_k^{(r)}$ be the quantity similarly defined in the nested configuration \mathcal{N}_ψ . We will show that

$$w_k^{(r)} \geq \tilde{w}_k^{(r)} \quad \text{for all } k. \tag{5.1}$$

Since clearly

$$t_e(\mathcal{N}_\psi) = \sum_{k=-\infty}^{\infty} \tilde{w}_k^{(r)} \quad \text{and} \quad t_e^{(r)}(\mathcal{E}_\psi) = \sum_{k=-\infty}^{\infty} w_k^{(r)},$$

this implies that $t_e(\mathcal{N}_\psi) \leq t_e^{(r)}(\mathcal{E}_\psi)$. The inequality $t_e(\mathcal{N}_\psi) \leq t_e^{(l)}(\mathcal{E}_\psi)$ is proved similarly.

To establish (5.1), assume that the edge e connects the vertices i and $i + n$ and is created by right-arrow number j at vertex i . In the nested configuration, all arrows in $\psi_e^{(r)}$ are connected to left-arrows at the vertices $i + 1, \dots, i + n$, meaning that $\tilde{w}_k^{(r)} = 0$ for $k \notin \{i + 1, \dots, i + n\}$ and, hence trivially, $w_k^{(r)} \geq \tilde{w}_k^{(r)}$ for such k . For $k \in \{i + 1, \dots, i + n\}$, note that in any edge configuration based on ψ , at least j $\psi_e^{(r)}$ -edges must cross the interval $[i, i + 1]$, implying that $w_{i+1}^{(r)} \geq j$. Furthermore, the interval $[i + 1, i + 2]$ must be crossed by at least $j + r_{i+1} - l_{i+1}$ $\psi_e^{(r)}$ -edges, where r_{i+1} and l_{i+1} respectively denote the numbers of right-arrows and left-arrows at vertex $i + 1$. Hence, $w_{i+2}^{(r)} \geq j + r_{i+1} - l_{i+1}$. Continuing in the same way, we obtain lower bounds for all $w_k^{(r)}$, $k \in \{i + 1, \dots, i + n\}$. From the construction of the nested configuration \mathcal{N}_ψ , it follows that these bounds hold with equality for the $\tilde{w}_k^{(r)}$, and we obtain (5.1).

Proof of Theorem 5.1. Let \mathcal{A} be a stationary pairing rule for an arrow configuration Ψ generated by attaching a random number of stubs to each vertex and then randomly assigning each stub the direction right or left. If, with positive probability, \mathcal{A} gives rise to configurations with infinitely long connections, then the conclusion of the theorem is immediate. Thus, assume that all edges in a configuration obtained from \mathcal{A} are finite almost surely, and write $T_{\mathcal{A}}^{(r)}$ and $T_{\mathcal{A}}^{(l)}$ for the respective total lengths of the edges created by the right-arrows and the left-arrows at the origin in an edge configuration generated by \mathcal{A} . We will show that $E[T_{\mathcal{A}}^{(r)}]$ and $E[T_{\mathcal{A}}^{(l)}]$ are both infinite.

To prove that $E[T_{\mathcal{A}}^{(r)}] = \infty$, let $T_{\mathcal{N}}^{(r)}$ be the total length of all edges created by the right-arrows at the origin in a nested configuration obtained from the stepwise pairing algorithm, and note that, by Theorem 4.1, we have $E[T_{\mathcal{N}}^{(r)}] = \infty$. If, with probability 1, \mathcal{A} results in a nested configuration, then $T_{\mathcal{A}}^{(r)}$ has the same distribution as $T_{\mathcal{N}}^{(r)}$ and the claim follows. Assume, therefore, that \mathcal{A} produces unnested configurations with positive probability, and let \mathcal{E} be such a configuration with underlying arrow configuration ψ . Write $t_i^{(r)}$ and $\tilde{t}_i^{(r)}$ for the respective total lengths of the edges created by the right-arrows at vertex i in the configurations \mathcal{E} and \mathcal{N}_ψ , and let $\tilde{m}_i^{(r)}$ be the length of the longest edge formed by the right-arrows at vertex i in \mathcal{N}_ψ . It follows from Lemma 5.1 that, for all i , we have

$$\sum_{j=i}^{i+\tilde{m}_i^{(r)}-1} t_j^{(r)} \geq \sum_{j=i}^{i+\tilde{m}_i^{(r)}-1} \tilde{t}_j^{(r)}.$$

By the ergodic theorem $E[T_{\mathcal{N}}^{(r)}]$ is equal to the average of $\tilde{t}_j^{(r)}$ and, hence, for every realization of \mathcal{A} , the average number of edges pointing to the right at a given vertex is bounded from below by $E[T_{\mathcal{N}}^{(r)}]$, proving that $E[T_{\mathcal{A}}^{(r)}] = \infty$. That $E[T_{\mathcal{A}}^{(l)}] = \infty$ is proved analogously, and the first claim of the theorem follows.

The second claim is established by noting that if k is an upper bound for the support of F , then $T_{\mathcal{A}} \leq kM_{\mathcal{A}}$.

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