

## A CONVERGENCE PROPERTY OF DUBINS' REPRESENTATION OF DISTRIBUTIONS

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Let  $(X_n)_{n \geq 1}$  be a sequence of random variables with zero means and uniformly bounded variances. Let  $\tau_n$  be the stopping time defined on a given Brownian motion  $(B_t)_{t \geq 0}$ ,  $B_0 = 0$ , by Dubins' method such that  $B(\tau_n)$  has the same distribution as  $X_n$ . We prove that  $X_n$  converging to 0 in distribution implies that  $\tau_n$  converges to 0 in probability. Examples are presented to illustrate the result is the best possible.

### 1. INTRODUCTION

This is an extension of my paper (Sheu [7]) which dealt with problems related to the Skorohod representation (Skorohod [8]). More specifically, one is seeking a probability space supporting a Brownian motion  $(B_t)_{t \geq 0}$  starting at the origin and a stopping time  $\tau$  so that  $B_\tau$  will be distributed according to a given distribution (or random variable). Special interest is focussed on whether such  $\tau$  can be required to depend only on Brownian paths without further randomisation (see Root [6], Dubins [4], Chacon and Walsh [3], Azéma and Yor [1], Bass [2], Vallois [9], etcetera).

For convenience, we use the notation  $B_\tau \sim X$  to mean  $B_\tau$  and  $X$  have the same distribution and say  $B_\tau$  represents  $X$ . In an investigation of Root's method of representation, Loynes [5] posed a problem of convergence. He asked whether the stopping time  $\tau$  depends on the distribution of  $X$  continuously in a certain sense. In fact, he showed that if  $(X_n)$  is a sequence of random variables with zero means and uniformly bounded variances and if  $X_n$  converges in distribution, then the stopping time  $\tau_n$  constructed by Root's method,  $B_{\tau_n} \sim X_n$ , converges in probability. Motivated by his results, we shall answer the same question for the Dubins method of representation in this paper.

### 2. MAIN RESULTS

The Dubins method can be described as follows (Dubins [4]). Let  $\mu$  be a distribution on the real line with finite expectation  $E\mu$ , let  $\mu^+$  and  $\mu^-$  denote the conditional distribution of  $\mu$  given  $[E\mu, \infty)$  and  $(-\infty, E\mu)$ , respectively. If  $\mu$  is degenerate, set  $\mu^+ = \mu^- = \mu$ . For any set  $K$  of  $n$ -tuples, let  $(m; K)$  be the set of  $(n+1)$ -tuples of the form  $(m, x)$ ,  $x \in K$ . Now introduce for each  $\mu$  and each  $n \geq 1$ , a finite set of

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$n$ -tuples of real numbers  $H_n(\mu)$  as follows:  $H_1(\mu) = \{E\mu\}$ ,  $K_n(\mu) = H_n(\mu^+) \cup H_n(\mu^-)$ ,  $H_{n+1}(\mu) = (E\mu; K_n(\mu))$ . Let  $(B_t)_{t \geq 0}$  be a given Brownian motion,  $B_0 = 0$ , and  $\mu$  a distribution with finite expectation  $E\mu$ . The Dubins stopping time  $\tau$  is defined to be the least  $t$  such that for all  $n \geq 1$ , there is an  $n$ -tuple  $t_1 \leq t_2 \leq \dots \leq t_n \leq t$  for which  $(B_{t_1}, B_{t_2}, \dots, B_{t_n}) \in H_n(\mu)$ . Dubins showed that  $B_\tau$  is distributed according to  $\mu$  and if  $E\mu = 0$ , then  $E(\tau) = \int_{-\infty}^{\infty} x^2 d\mu$ . Turning to the question of convergence, we have

**THEOREM 1.** *Let  $(X_n)_{n \geq 1}$  be a sequence of random variables with zero means and uniformly bounded variances. Let  $\tau_n$  be the Dubins stopping time for representing  $X_n$ ,  $B_{\tau_n} \sim X_n$ ,  $n \geq 1$ . Then  $X_n$  converging to 0 in distribution implies  $\tau_n$  converges to 0 in probability.*

**PROOF:** Let  $\mu_n$  be the distribution of  $X_n$ ,  $n \geq 1$ , and let

$$T_a = \inf\{t \geq 0: B_t = a\}.$$

Since  $X_n$  converges to 0 in distribution and  $E(X_n^2)$  is uniformly bounded, we conclude  $E|X_n| \rightarrow 0$ ; that is

$$E(X_n^+) \rightarrow 0 \text{ and } E(X_n^-) \rightarrow 0.$$

Given  $\epsilon > 0$ , choose  $\delta > 0$  so that  $P(T_\delta \geq \frac{\epsilon}{2}) \leq \frac{\epsilon}{2}$ . Then there exists an  $N = N(\epsilon)$  such that, if  $n \geq N$ , either

$$E\mu_n^+ = E(X_n^+)/P(0 \leq X_n < \infty) \leq \delta$$

or

$$E\mu_n^- = E(X_n^-)/P(-\infty < X_n < 0) \leq \delta.$$

By Dubins' construction, we see

$$\tau_n \leq \inf\{t \geq 0: B_t = 0, B_s = \delta, \text{ some } s \leq t\}$$

or

$$\tau_n \leq \inf\{t \geq 0: B_t = 0, B_s = -\delta, \text{ some } s \leq t\}.$$

By the strong Markov property of Brownian motion, we have  $P(\tau_n \geq \epsilon) \leq 2P(T_\epsilon \geq \frac{\epsilon}{2}) \leq \epsilon$ . That is,  $\tau_n$  converges to 0 in probability. ■

The following examples indicate the result obtained is the best possible.

**Example 1.** Let  $P(X_n = 0) = 1 - \frac{2}{n}$ ,  $P(X_n = n) = P(X_n = -n) = \frac{1}{n}$ . Then  $E(X_n) = 0$ ,  $E(X_n^2) = 2n$ . By computation,  $E\mu_n^+ = \frac{n}{n-1}$ ,  $E\mu_n^- = -n$ . Clearly,  $X_n$  converges to 0 in distribution but  $\tau_n$  does not. This example shows that the assumption of uniformly bounded variances can not be omitted.

**Example 2.** Let  $P(X_n = \frac{3n+1}{n}) = \frac{2}{10}$ ,  $P(X_n = -\frac{1}{n}) = \frac{1}{2}$ ,  $P(X_n = \frac{-2n+1}{n}) = \frac{3}{10}$  and let  $P(X = 3) = \frac{2}{10}$ ,  $P(X = 0) = \frac{1}{2}$ ,  $P(X = -2) = \frac{3}{10}$ . Then  $E(X) = E(X_n) = 0$ ,  $E(X_n^2) \leq \frac{32}{5}$ . By computation,  $\mu_n^+$ ,  $\mu_n^{-+}$ ,  $\mu_n^{--}$  are degenerate and  $E\mu_n^+ = \frac{3n+1}{n} \rightarrow 3$ ,  $E\mu_n^- = \frac{-3n-1}{4n} \rightarrow -\frac{3}{4}$ ,  $E\mu_n^{-+} = -\frac{1}{n} \rightarrow 0$ ,  $E\mu_n^{--} = \frac{-2n+1}{n} \rightarrow -2$ . Also,  $\mu^-$ ,  $\mu^{++}$ ,  $\mu^{+-}$  are degenerate and  $E\mu^+ = \frac{6}{7}$ ,  $E\mu^- = -2$ ,  $E\mu^{++} = 3$ ,  $E\mu^{--} = 0$ . Therefore,  $X_n$  converges to  $X$  in distribution but the Dubins stopping time,  $\tau_n$ ,  $B_{\tau_n} \sim X_n$ , does not converge to  $\tau$  in probability, where  $B_\tau \sim X$ . This example shows that the assumption that the limiting distribution of  $X_n$  is degenerate can not be omitted.

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