

#### ARTICLE

# Behaviour of the minimum degree throughout the *d*-process\*

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#### Abstract

The d-process generates a graph at random by starting with an empty graph with n vertices, then adding edges one at a time uniformly at random among all pairs of vertices which have degrees at most d-1 and are not mutually joined. We show that, in the evolution of a random graph with n vertices under the d-process with d fixed, with high probability, for each  $j \in \{0, 1, \ldots, d-2\}$ , the minimum degree jumps from j to j+1 when the number of steps left is on the order of  $\ln(n)^{d-j-1}$ . This answers a question of Ruciński and Wormald. More specifically, we show that, when the last vertex of degree j disappears, the number of steps left divided by  $\ln(n)^{d-j-1}$  converges in distribution to the exponential random variable of mean  $\frac{j!}{2(d-1)!}$ ; furthermore, these d-1 distributions are independent.

**Keywords:** Random graph; d-process **2020 MSC Codes:** Primary: 05C80

#### 1. Introduction

There are numerous models that generate different types of sparse random graphs. Among them is the *d-process*, defined in the following way: start with *n* vertices and 0 edges, and at each time step, choose a pair  $\{u,v\}$  uniformly at random over all pairs consisting of vertices which have degree less than *d* and are not joined to each other by an edge. *d* could be allowed to change with *n*, but for the rest of this paper *d* is always a fixed constant (this is also assumed in all relevant citations). Ruciński and Wormald showed that with high probability, abbreviated "w.h.p." (i.e. with probability converging to 1 as  $n \to \infty$ ) the *d*-process ends with  $\lfloor dn/2 \rfloor$  edges [11]. There is still much that is unknown about the *d*-process; for example, it is not known whether the *d*-process is *contiguous* with the *d*-uniform random graph model for any  $d \ge 2$ ; i.e. if any event that happens with high probability in one happens with high probability in the other. A recent paper by Molloy, Surya, and Warnke [8] disproves this relation if there is enough "non-uniformity" of the vertex degrees (with an appropriate modification of the *d*-process for non-regular graphs); it also contains a good history of the *d*-process. See ref. [7, Section 9.6] for more on contiguity.

A couple of notable results have been given for the case where d = 2: the expected numbers of cycles of constant sizes were studied by Ruciński and Wormald in ref. [10], and in ref. [13], Telcs, Wormald, and Zhou calculated the probability that the 2-process ends with a Hamiltonian cycle.

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In these works, the authors establish estimates on certain graph parameters, such as the number of vertices of a certain degree, that hold throughout the process. This is done with the so-called "differential equations method" for random graph processes, which uses martingale inequalities to give variable bounds; in ref. [14] Wormald gives a thorough description of this method.

More recently, Ruciński and Wormald announced a new analysis of the d-process that hinges on a coupling with a balls-in-bins process. This simple argument gives a precise estimate of the probability that the d-process ends with  $\lfloor dn/2 \rfloor$  edges (i.e. the probability that the d-process reaches saturation). This argument includes estimates for the number of vertices of each degree near the end of the process. This work was presented by Ruciński at the 2023 *Random Structures and Algorithms* conference. Ruciński's presentation included the following problem (which was open at the time): when do we expect the last vertex of degree j (for any j from 0 to d-2) to disappear? The question was also stated earlier for d=2 and j=0 by Ruciński and Wormald [10, Question 3]. In November of 2023, after the first release of the pre-print of this paper, Ruciński and Wormald released a pre-print of their balls-and-bins argument which also included an answer to Ruciński's question [12]. Our main result uses the differential equations method (as described in the previous paragraph) and gives a slightly stronger answer:

**Theorem 1.** Consider the d-process on a vertex set of size n, and for each  $\ell \in \{0\} \cup [d-2]$ , let the random variable  $T_{\ell}$  be the step at which the number of vertices of degree at most  $\ell$  becomes 0. Then the sequence (over n) of random d-1-tuples consisting of the variables

$$V_n^{(\ell)} = \frac{(d-1)!(dn-2T_\ell)}{\ell!(\ln{(n)})^{d-1-\ell}}$$

converges in distribution to the product of d-1 independent exponential random variables of mean 1.

In this paper we use the differential equations method with increasingly precise estimates of certain random variables; these estimates are known as *self-correcting*. Previous results that use self-correcting estimates include [13], [6], [3], [4], [5], and [9]. There have been various approaches to achieving self-correcting estimates; the approach in this paper uses *critical intervals*, regions of possible values of a random variable in which we expect subsequent variables to increase/decrease over time. Critical intervals used in this fashion first appeared in a result of Bohman and Picollelli [6]. For an introduction to and discussion of the method see Bohman, Frieze, and Lubetzky [3].

The proof of Theorem 1 is divided into four sections. In Section 2, we introduce random variables of the form  $S_i^{(j)}$  which count the number of vertices of degree at most j after i steps, define approximating functions  $s_j(t)$  with the eventual goal of showing that  $S_i^{(j)} \approx ns_j(i/n)$  for most of the process, and derive useful properties of these functions. One such property is that, when there are at most  $n^c$  steps left for some constant c < 1,

$$\frac{s_j(i/n)}{s_{j-1}(i/n)} = \Theta(\ln(n))$$

for each j; this hierarchy of functions helps us to focus on each variable  $S_i^{(j)}$  independently of the others when it is near 0, which motivates the form of the limiting exponential random variables in Theorem 1. At the end of Section 2 we introduce two martingale inequalities used by Bohman [2] and make a slight modification to them to use later in the paper. In Section 3, we work with a more 'standard martingale method' (without the use of critical intervals) to show that  $S_i^{(j)} \approx ns_j(i/n)$  until there are  $n^{1-1/(100d)}$  steps left. Here we allow the error bounds to increase over time. In Section 4, we use a more refined martingale method (including the use of critical intervals) to show that  $S_i^{(j)} \approx ns_j(i/n)$  continues to hold until there are  $\ln(n)^{d-0.8-j}$  steps left; here the error bounds

decrease over time, and so are self-correcting. In Section 5, we complete the proof of Theorem 1 by using a pairing argument to show that, in the last steps of the d-process, the behaviour of the random variable in question can be well-approximated by a certain uniform distribution of time steps. Sections 4 and 5 are both parts of a proof by induction over a series of intervals of time steps, though we give each part its own section as the methods used in each are very different.

## 2. Preliminaries

First, two technical notes: we use the standard notation of symbols o, O,  $\Theta$ ,  $\omega$ ,  $\Omega$ ,  $\ll$ ,  $\gg$ , and  $\sim$  to compare functions asymptotically (e.g. see pages 9-10 of [7]). We also note that, throughout the paper, we assume n to be arbitrarily large.

In this section we set up sequences of random variables, describe how the evolution of the d-process depends on these, and deduce properties of certain *approximating functions*; such functions are used to estimate the number of vertices of given degrees throughout the process (much of this is also described in ref. [13] with similar notation; the one major difference is that we use i instead of t for the number of time steps, and t instead of x for the corresponding time variable). Consider a sequence of graphs  $G_0, G_1, \ldots, G_{\lfloor dn/2 \rfloor}$ , where  $G_0$  is the empty graph of n vertices, and for  $i \in [n]$ , let  $G_i$  be formed by adding an edge uniformly at random to  $G_{i-1}$  so that the maximum degree of  $G_i$  is at most d (in the unlikely event that there are no valid edges to add after s steps for some  $s < \lfloor dn/2 \rfloor$ , let  $G_i = G_s$  for all i > s). Next, we define several sequences of random variables: For all  $i, j, j_1, j_2$  such that  $0 \le i \le \lfloor \frac{dn}{2} \rfloor$ ,  $0 \le j \le d$ , and  $0 \le j_1 \le j_2 \le d - 1$ , define:

 $Y_i^{(j)}$  := the number of vertices in  $G_i$  with degree j  $S_i^{(j)}$  := the number of vertices in  $G_i$  with degree at most j

 $Z_i^{(j_1,j_2)} :=$ the number of edges  $\{v_1, v_2\}$  in  $G_i$  for which

 $\min\{\deg(v_1), \deg(v_2)\} = j_1 \text{ and } \max\{\deg(v_1), \deg(v_2)\} = j_2$ 

$$Z_i := \sum_{0 \le j_1 \le j_2 \le d-1} Z_i^{(j_1, j_2)}.$$

By definition and by edge-counting we have

$$\sum_{j=0}^{d} Y_i^{(j)} = n$$
 and  $\sum_{j=1}^{d} j Y_i^{(j)} = 2i$ .

Combining the equations above gives us

$$\sum_{i=0}^{d-1} S_i^{(j)} = \sum_{i=0}^{d-1} (d-j) Y_i^{(j)} = dn - 2i.$$
 (1)

One can also quickly verify from (1) and by definition of  $S_i^{(d-1)}$ ,  $Z_i^{(j_1,j_2)}$ , and  $Z_i$ , that

$$dS_i^{(d-1)} \ge \max\{2Z_i, dn - 2i\} \quad \text{and} \quad jY_i^{(j)} \ge \sum_{k \le j} Z_i^{(k,j)} + \sum_{k \ge j} Z_i^{(j,k)}. \tag{2}$$

Throughout the process we will keep track of the variables  $S^{(j)}$  using martingale arguments. This is sufficient for us, as the Y variables can be derived from the S variables, and because none of the Z variables will have any significant effect in any of our calculations, as we will see later.

Our next step is to estimate the expected on-step change of  $S_i^{(j)}$ , known as the "trend hypothesis" in ref. [14]. Note that  $S_i^{(j)} - S_{i+1}^{(j)}$  equals the number of vertices of degree j that are picked at the i+1 time step; hence, for all  $j \in \{0\} \cup [d-1]$ :

$$\mathbb{E}\left[S_{i+1}^{(j)} - S_{i}^{(j)} \mid G_{i}\right] = \frac{-Y_{i}^{(j)}\left(S_{i}^{(d-1)} - 1\right) + \sum_{k \leq j} Z_{i}^{(k,j)} + \sum_{k \geq j} Z_{i}^{(j,k)}}{\binom{S_{i}^{(d-1)}}{2} - Z_{i}}$$

$$= \frac{-2Y_{i}^{(j)}}{S_{i}^{(d-1)}} \left(1 + O\left(\frac{1}{dn - 2i}\right)\right) \quad \text{by (2)}$$

$$= \frac{-2Y_{i}^{(j)}}{S_{i}^{(d-1)}} + O\left(\frac{1}{dn - 2i}\right). \tag{4}$$

For  $j \in \{0\} \cup [d-1]$  we define approximating functions  $y_j : [0, d/2) \to \mathbb{R}$  and  $s_j : [0, d/2) \to \mathbb{R}$ : let  $y_j(t)$ ,  $s_j(t)$  be functions such that  $s_j = \sum_{k=0}^j y_k$ ,  $y_0(0) = 1$  and  $y_k(0) = 0$  for all  $k \in [d-1]$  (equivalent to  $s_j(0) = 1$  for all j), and (assuming the "dummy functions"  $y_{-1}(t) = s_{-1}(t) = 0$ ):

$$\frac{ds_j}{dt} = \frac{-2y_j}{s_{d-1}} = \frac{2(s_{j-1} - s_j)}{s_{d-1}} \qquad \frac{dy_j}{dt} = \frac{2(y_{j-1} - y_j)}{s_{d-1}}.$$
 (5)

By (5) and the chain rule, for all  $j \in [d-1]$ :

$$\frac{dy_j}{dy_0} - \frac{y_j}{y_0} = -\frac{y_{j-1}}{y_0}.$$

Since the above equation is first-order linear, we have, for some constant  $C_j$ :

$$y_j = y_0 \left( -\int \frac{y_{j-1}}{y_0^2} dy_0 + C_j \right).$$

Using the above recursively with initial conditions, we have, for all  $j \in [d-1]$ :

$$y_j = \frac{y_0(-\ln(y_0))^j}{j!}.$$
 (6)

To solve for an explicit formula relating  $y_0$  and t, note that, by (5):

$$\frac{d}{dt} \left( \sum_{j=0}^{d-1} (d-j) y_j \right) = \frac{-2 \sum_{j=0}^{d-1} y_j}{s_{d-1}} = -2,$$

hence, using initial conditions (note the resemblance to (1)):

$$\sum_{j=0}^{d-1} s_j = \sum_{j=0}^{d-1} (d-j)y_j = d-2t.$$
 (7)

Equations (6) and (7) together give us a complete description of the functions  $y_j$  and  $s_j$ . We will now prove some useful properties of these functions. To start, we can combine (6) and (7) to get

$$\sum_{j=0}^{d-1} \frac{y_0(-\ln(y_0))^j(d-j)}{j!} = d - 2t.$$
 (8)

Note that, by continuity of  $y_0$  and by (8),  $y_0 > 0$  over its domain. Next, by summing up (6) over  $j \in \{0\} \cup [d-1]$  ((6) holds for j=0 also) one can see that  $s_{d-1}$  is positive if  $y_0 \le 1$ . This, combined with  $\frac{dy_0}{dt} = \frac{-2y_0}{s_{d-1}}$ , tells us that  $y_0$  is decreasing and  $s_{d-1}$  is positive. It follows from  $y_0 \in (0,1]$  and (6) that each  $y_j$  is positive for  $t \ne 0$ . In turn, this implies that  $0 \le y_j \le s_j \le s_{d-1}$  for each j. From this it follows that  $ds_{d-1}$  is at least the left expression of (7), so  $s_{d-1} \ge 1 - \frac{2t}{d}$ . We make a special note of the last couple of properties mentioned:

$$0 \le y_j \le s_j \le s_{d-1} \text{ for all } j \quad \text{and} \quad s_{d-1}(i/n) \ge 1 - \frac{2i}{dn}.$$
 (9)

Next, we want to understand the behaviour of each function when t is close to  $\frac{d}{2}$ , as this is the most critical point of the process. Consider (8) again. As  $t \to \frac{d}{2}$ ,  $y_0 \to 0$ , so  $\frac{y_0(-\ln(y_0))^{d-1}}{(d-1)!}$  will be the most dominant term on the left; hence,

$$t \to \frac{d}{2} \Longrightarrow y_0 \sim \frac{(d-1)!(d-2t)}{(-\ln{(d-2t)})^{d-1}}.$$

This, combined with (6) gives us, for all  $j \in \{0\} \cup [d-1]$ :

$$t \to \frac{d}{2} \Longrightarrow y_j(t) \sim s_j(t) \sim \frac{(d-1)!(d-2t)}{j!(-\ln(d-2t))^{d-1-j}}.$$
 (10)

For large enough t (and hence, for a large enough step i), we can approximate the above expression:

$$i \ge \frac{dn}{2} - n^{1-1/(100d)} \implies ny_j(i/n) \sim ns_j(i/n) = \Theta\left(\ln\left(n\right)^{-d+1+j} \left(\frac{dn}{2} - i\right)\right). \tag{11}$$

One can now see that, near the end of the process,  $s_j/s_{j-1} = \Theta(\ln(n))$ , as mentioned in the introduction.

Finally, we introduce two martingale inequalities from a result of Bohman [2] which will be used in Section 4 in a slightly modified form. The original inequalities are as follows:

**Lemma 2** (Lemma 6 from [2]). Suppose  $a, \eta$ , and N are positive,  $\eta \le N/2$ , and  $a < \eta m$ . If  $0 = A_0, A_1, \ldots, A_m$  is a submartingale such that  $-\eta \le A_{i+1} - A_i \le N$  for all i, then

$$\mathbb{P}[A_m < -a] < e^{-\frac{a^2}{3\eta Nm}}.$$

**Lemma 3** (Lemma 7 from [2]). Suppose a,  $\eta$ , and N are positive,  $\eta \le N/10$ , and  $a < \eta m$ . If  $0 = A_0, A_1, \ldots, A_m$  is a supermartingale such that  $-\eta \le A_{i+1} - A_i \le N$  for all i, then

$$\mathbb{P}[A_m \ge a] \le e^{-\frac{a^2}{3\eta Nm}}.$$

We present the following modification, which removes the requirement  $a < \eta m$  and modifies one of the inequalities slightly:

**Corollary 4.** Suppose a,  $\eta$ , and N are positive, and  $\eta \le N/2$ . If  $0 = A_0, A_1, \ldots, A_m$  is a submartingale such that  $-\eta \le A_{i+1} - A_i \le N$  for all i, then

$$\mathbb{P}[A_m \le -a] \le e^{-\frac{a^2}{3\eta Nm}}.$$

**Corollary 5.** Suppose  $a, \eta$ , and N are positive, and  $\eta \le N/10$ . If  $0 = A_0, A_1, \ldots, A_m$  is a supermartingale such that  $-\eta \le A_{i+1} - A_i \le N$  for all i, then

$$\mathbb{P}[A_m \ge a] \le e^{-\frac{a^2}{3\eta Nm}} + e^{-\frac{a}{6N}}.$$

Corollary 4 is nearly immediate from Lemma 2: first, one can extend the result to include  $a = \eta m$  by using left-continuity (with respect to a) of both sides of the inequality; we hence assume  $a > \eta m$ . Since  $A_{i+1} - A_i \ge -\eta > -a/m$ , then  $A_m = A_m - A_0 > -a$ . We now derive Corollary 5 from Lemma 3: assume  $a \ge \eta m$ , and let  $m' \in \mathbb{Z}^+$  such that  $a < \eta m' \le 2a$ . Extend the martingale by adding variables  $A_{m+1}, \ldots, A_{m'}$  which are all equal to  $A_m$ . Apply Lemma 3 with m replaced with m', and use  $\eta m' \le 2a$  to get

$$\mathbb{P}[A_m \ge a] = \mathbb{P}[A_{m'} \ge a] \le e^{-\frac{a}{6N}}.$$

Combining the case  $a < \eta m$  from Lemma 3 and the case  $a \ge \eta m$  above gives Corollary 5.

# 3. First phase

Let  $i_{trans} = \lfloor \frac{dn}{2} - n^{1-1/(100d)} \rfloor$ . The objective of this section is to prove the following Theorem:

**Theorem 6.** Define  $E_{first}(i) := n^{0.6} \left(\frac{dn}{dn-2i}\right)^{4d}$ . With high probability, for all  $i \le i_{trans}$  and all  $j \in \{0\} \cup [d-1]$ :

$$\left|S_i^{(j)} - ns_j\left(\frac{i}{n}\right)\right| \le E_{first}(i). \tag{12}$$

Now we define two new random variables for each *j*:

$$S_i^{(j)+} := S_i^{(j)} - ns_j(i/n) - E_{first}(i)$$
  

$$S_i^{(j)-} := S_i^{(j)} - ns_j(i/n) + E_{first}(i).$$

Next, we introduce a stopping time T, defined as the first step  $i \le i_{trans}$  for which (12) is *not* satisfied for some j; if (12) always holds, then let  $T = \infty$ . Although this stopping time is not necessarily needed to prove Theorem 6, it does make some calculations easier, and moreover, a similar stopping time *will* be necessary for the following section; hence, this serves as a good warm-up. Let variable name W be introduced to equip this stopping time to variable S, i.e.

$$W_i^{(j)+} := \begin{cases} S_i^{(j)+}, i < T \\ S_T^{(j)+}, i \ge T \end{cases} \qquad W_i^{(j)-} := \begin{cases} S_i^{(j)-}, i < T \\ S_T^{(j)-}, i \ge T. \end{cases}$$

Note that  $W_i^{(j)+}$  corresponds to the upper boundary and  $W_i^{(j)-}$  to the lower one in the sense that crossing the corresponding boundary will make the corresponding variable change signs; furthermore, the inequality of Theorem 6 holds if and only if  $W_{i_{trans}}^{(j)+} \leq 0$  and  $W_{i_{trans}}^{(j)-} \geq 0$  for each j. We now state our martingale Lemma:

**Lemma 7.** Restricted to  $i \le i_{trans}$ , for all j,  $(W_i^{(j)-})_i$  is a submartingale and  $(W_i^{(j)+})_i$  is a supermartingale.

**Proof.** Here we just prove the first part of the Lemma; the second part follows from nearly identical calculations. Fix some arbitrary  $i \le i_{trans}$ ; we need to show that

$$\mathbb{E}\left[W_{i+1}^{(j)-}-W_i^{(j)-}\mid G_i\right]\geq 0.$$

Also assume that  $T \ge i+1$ , else  $W_{i+1}^{(j)-} - W_i^{(j)-} = 0$  and we are done; it follows that  $W_i^{(j)-} = S_i^{(j)-}$ ,  $W_{i+1}^{(j)-} = S_{i+1}^{(j)-}$ , and (12) holds for the fixed i. By (4) and (5) and using Taylor's Theorem, we have, for some  $\psi \in [i, i+1]$ :

$$\mathbb{E}\left[S_{i+1}^{(j)-} - S_{i}^{(j)-} \mid G_{i}\right] = \frac{-2Y_{i}^{(j)}}{S_{i}^{(d-1)}} + O\left(\frac{1}{dn-2i}\right) + \frac{2y_{j}(i/n)}{s_{d-1}(i/n)} - \frac{d^{2}}{d\mu^{2}}\left(\frac{ns_{j}(\mu/n)}{2}\right)\Big|_{\mu=\psi} + \left(E_{first}(i+1) - E_{first}(i)\right).$$

We split the above expression (excluding  $O\left(\frac{1}{dn-2i}\right)$ ) into three summands.

1. Here we make use of the fact that  $Y_i^{(j)} = S_i^{(j)} - S_i^{(j-1)}$  and  $y_j(t) = s_j(t) - s_{j-1}(t)$ . We have (putting  $s_{j-1}$  and  $S_i^{(-1)} = 0$ ):

$$\frac{-2Y_i^{(j)}}{S_i^{(d-1)}} + \frac{2y_j(i/n)}{s_{d-1}(i/n)} = \frac{-2S_i^{(j)} + 2ns_j(i/n) + 2S_i^{(j-1)} - 2ns_{j-1}(i/n)}{ns_{d-1}(i/n)} + 2Y_i^{(j)} \left(\frac{1}{ns_{d-1}(i/n)} - \frac{1}{S_i^{(d-1)}}\right) \\
\geq \frac{-4E_{first}(i)}{ns_{d-1}(i/n)} + 2Y_i^{(j)} \left(\frac{1}{ns_{d-1}(i/n)} - \frac{1}{S_i^{(d-1)}}\right) \quad \text{by (12) and } i < T \\
\geq \frac{-4E_{first}(i)}{ns_{d-1}(i/n)} - \frac{2Y_i^{(j)}E_{first}(i)}{S_i^{(d-1)}(ns_{d-1}(i/n))} \quad \text{by (12) and } i < T \\
\geq \frac{-6E_{first}(i)}{ns_{d-1}(i/n)}.$$

2.

$$-\frac{d^2}{d\mu^2} \left( \frac{n s_j(\mu/n)}{2} \right) \Big|_{\mu=\psi} = \frac{2}{n} \left( \frac{s_{d-1}(\psi/n)(y_{j-1}(\psi/n) - y_j(\psi/n)) + y_j(\psi/n)y_{d-1}(\psi/n)}{(s_{d-1}(\psi/n))^3} \right)$$

$$= O\left( \frac{1}{dn-2i} \right) \quad \text{by (9)}.$$

3. For some  $\phi \in [i, i+1]$ :

$$E_{first}(i+1) - E_{first}(i) = \frac{dE_{first}(\mu)}{d\mu} \Big|_{\mu=\phi}$$

$$= 8d^{4d+1}n^{4d+0.6} (dn - 2\phi)^{-4d-1}$$

$$= (1+o(1))\frac{8dE_{first}(i)}{dn-2i}.$$
(13)

Now we put the three bounds together:

$$\mathbb{E}\left[Y_{i+1}^{-} - Y_{i}^{-} \mid G_{i}\right] \ge \frac{7dE_{first}(i)}{dn - 2i} - \frac{6E_{first}(i)}{ns_{d-1}(i/n)} + O\left(\frac{1}{dn - 2i}\right)$$

$$\ge \frac{dE_{first}(i) + O(1)}{dn - 2i} \quad \text{by (9)}$$

$$\ge 0.$$

Next, we need a Lipschitz condition on each of our variables. Note that  $S_{i+1}^{(j)} - S_i^{(j)}$  is either -2, -1, or 0; also, one can quickly verify that  $|s_j((i+1)/n) - s_j(i/n)| \le \frac{2}{n}$  by (5) and (9), and  $|E_{first}(i+1) - E_{first}(i)| = o(1)$  by (13). Hence, we have, for all  $i \le i_{trans}$  and all j:

$$\max\left\{ \left| W_{i+1}^{(j)+} - W_i^{(j)+} \right|, \left| W_{i+1}^{(j)-} - W_i^{(j)-} \right| \right\} \le 5.$$
 (14)

We conclude the proof of Theorem 6 by noting that, by Lemma 7 and (14), we can use the standard Hoeffding-Azuma inequality for martingales (e.g. Theorem 7.2.1 in [1]) to show that  $\mathbb{P}\left[W_{i_{trans}}^{(j)+}>0\right]$  and  $\mathbb{P}\left[W_{i_{trans}}^{(j)-}<0\right]$  are both o(1). For example, for the variable  $W_{i}^{(j)+}$  one would get

$$\mathbb{P}\left[W_{i_{trans}}^{(j)+} > 0\right] \le \exp\left\{-\frac{n^{1.2}}{50i_{trans}}\right\} = o(1).$$

## 4. Second phase

The second phase is where the more sophisticated tools will be used, including the use of critical intervals, self-correcting estimates, and a more general martingale inequality. Furthermore, this phase is broken up into d-1 sub-phases, in relation to when each of the d-1 sequences  $S^{(j)}$  (for  $j \le d-2$ ) terminate at 0. First, a few definitions: for all  $k \in \{0\} \cup [d-2]$ , define

$$i_{after}(k) := \left| \frac{dn}{2} - \ln{(n)^{d-1.01-k}} \right|.$$

These step values will govern the endpoints of the sub-phases: define for all  $k \in \{0\} \cup [d-2]$ :

$$I_k := \begin{cases} \left[i_{trans} + 1, i_{after}(0)\right], & k = 0\\ \left[i_{after}(k-1) + 1, i_{after}(k)\right], & k > 0. \end{cases}$$

Next, for all i, j, k such that  $0 \le j < d$ ,  $0 \le k < d - 1$ , and  $i \in I_k$ , define error functions

$$E_{j,k}(i) = E_j(i) := 2^k \ln(n)^{0.05} (ns_j(i/n))^{0.7}.$$

Note that, by (11), we have

$$E_j(i) = \Theta\left(\ln(n)^{-0.7d + 0.75 + 0.7j} \left(\frac{dn}{2} - i\right)^{0.7}\right). \tag{15}$$

Finally, for any  $r \in \mathbb{R}_+$  and  $\ell \in [d-2]$ , define

$$i(r,\ell) = \frac{dn}{2} - \left(\frac{\ell!}{2(d-1)!}\right) r(\ln n)^{d-1-\ell}.$$

The following Theorem will be proved by induction over the d-1 sub-phases governed by the index k:

**Theorem 8.** *For each k* ∈  $\{0\}$  ∪ [d-2]:

1. With high probability, for all integers  $j \in [0, d-1]$  and  $i \in I_k$ :

$$\left| S_i^{(j)} - ns_j \left( \frac{i}{n} \right) \right| \le 4E_j(i). \tag{16}$$

2.  $S_{i_{after}(k)}^{(k)} = 0$  with high probability. Furthermore, for any k+1-tuple  $\{r_0, r_1, \ldots, r_k\} \in (\mathbb{R}_+ \cup \{0\})^{k+1}$ :

$$\mathbb{P}\left(\bigcap_{\ell=0}^{k} \left(S_{\lfloor i(r_{\ell},\ell)\rfloor}^{(\ell)} = 0\right)\right) \to \exp\left\{-\sum_{\ell=0}^{k} r_{\ell}\right\}.$$

In the end, it is only the second statement with k = d - 2 that matters for proving Theorem 1. We make the connection here:

**Proof of Theorem** 1 from Theorem 8. First, note that  $S_{\lfloor i(r_{\ell},\ell) \rfloor}^{(\ell)} = 0$  is the same as  $T_{\ell} \leq i(r_{\ell},\ell)$ , hence by Theorem 8:

$$\mathbb{P}\left(\bigcap_{\ell=0}^{d-2} \left(T_{\ell} \leq i(r_{\ell}, \ell)\right)\right) \to \exp\left\{-\sum_{\ell=0}^{d-2} r_{\ell}\right\}.$$

Using the Principle of Inclusion-Exclusion plus a simple limiting argument, one can derive

$$\mathbb{P}\left(\bigcap_{\ell=0}^{d-2}\left(\frac{(d-1)!(dn-2T_{\ell})}{\ell!(\ln{(n)})^{d-1-\ell}}\leq r_{\ell}\right)\right)=\mathbb{P}\left(\bigcap_{\ell=0}^{d-2}\left(T_{\ell}\geq i(r_{\ell},\ell)\right)\right)\to \prod_{\ell=0}^{d-2}\left(1-e^{-r_{\ell}}\right),$$

hence the d-1-dimensional random vector with entries  $V_n^{(\ell)} = \frac{(d-1)!(dn-2T_\ell)}{\ell!(\ln{(n)})^{d-1-\ell}}$  converges in distribution to the product of d-1 independent exponential variables of mean 1.

The rest of this section is for proving the first statement of Theorem 8 (for some fixed k using induction), and Section 5 will be for proving the second statement (again, for some fixed k using induction, assuming the first statement holds for the same k). Hence, for the rest of the paper we will fix some  $k \in \{0\} \cup [d-2]$ .

First, we note that (16) holds w.h.p. for all j < k by a simple argument: by induction on the second statement of Theorem 8, w.h.p. if  $i \in I_k$  then  $S_i^{(j)} = 0$ . By (11) and by definition of  $E_j(i)$ , if  $i \in I_k$  then  $ns_j(i/n) \ll E_j(i)$ , completing the argument.

Next, we prove that (16) holds for j = d - 1 if it holds for all other values of j: by combining (1) and (7), we have

$$\left| S_i^{(d-1)} - ns_{d-1} \left( \frac{i}{n} \right) \right| = \left| \sum_{j=0}^{d-2} \left( S_i^{(j)} - ns_j \left( \frac{i}{n} \right) \right) \right|$$

$$\leq \sum_{j=0}^{d-2} 4E_j(i) \qquad \text{by (16) for } j \leq d-2$$

$$< 4E_{d-1}(i) \qquad \text{by (15)}.$$

Hence, for the rest of this section, we need to show the first statement of Theorem 8 for  $j \in [k, d-2]$ . From now on we always assume j to be in this range. We will *also* assume that, for all  $\lambda < k$ ,  $S_i^{(\lambda)} = 0$  if  $i \in I_k$  (which holds w.h.p. from above).

In this section we will make use of so-called *critical intervals*, ranges of possible values for  $S_i^{(j)}$  in which we apply a martingale argument. The lower critical interval will be

$$[ns_i(i/n) - 4E_i(i), ns_i(i/n) - 3E_i(i)],$$

and the upper critical interval will be

$$[ns_i(i/n) + 3E_i(i), ns_i(i/n) + 4E_i(i)].$$

Our goal is to show that w.h.p.  $S_i^{(j)}$  does not cross either critical interval; however, we first need to show that  $S_i^{(j)}$  sits between the critical intervals at the beginning of the phase (this is the reason why  $E_j(i)$  has the  $2^k$  factor; it makes a sudden jump in size between phases to accommodate a new martingale process), which is the statement of our first Lemma of this section:

**Lemma 9.** W.h.p., for all  $j \in [k, d-2]$  (putting  $i_{after}(-1) = i_{trans}$  for convenience of notation):

$$\left|S_{i_{after}(k-1)+1}^{(j)}-ns_{j}\left(\frac{i_{after}(k-1)+1}{n}\right)\right|<3E_{j}(i_{after}(k-1)+1).$$

**Proof.** First, recall that  $S_{i+1}^{(j)} - S_i^{(j)} \in \{-2, -1, 0\}$  and  $|ns_j((i+1)/n) - ns_j(i/n)| \le 2$  for any i and j (see paragraph above (14)). Second, consider the change in the bound itself between  $i_{after}(k-1)$  and  $i_{after}(k-1)+1$ : by definitions of  $i_{trans}$ ,  $E_{first}$ ,  $E_j$ , and by (15), we have  $1 \ll E_{first}(i_{trans}) = \Theta\left(n^{0.64}\right)$ ,  $E_j(i_{trans}+1) = \omega\left(n^{0.69}\right)$ , and  $1 \ll E_j(i_{after}(k-1)) \approx \frac{1}{2}(E_j(i_{after}(k-1)+1))$  for k>0. Hence, by induction on the first statement of Theorem 8 and by Theorem 6, the statement of the Lemma follows.

Next, like in Section 3, we define two new random variables for each j and  $i \in I_k$ :

$$S_i^{(j)+} := S_i^{(j)} - ns_j(i/n) - 4E_j(i)$$
  

$$S_i^{(j)-} := S_i^{(j)} - ns_j(i/n) + 4E_j(i).$$

We also re-introduce the stopping time T, now defined as the first step  $i \in I_k$  for which (16) is *not* satisfied for some j; if (16) always holds, then let  $T = \infty$ . Let variable name W be introduced to equip this stopping time to variable S, i.e.

$$W_i^{(j)+} := \begin{cases} S_i^{(j)+}, i < T \\ S_T^{(j)+}, i \ge T \end{cases} \qquad W_i^{(j)-} := \begin{cases} S_i^{(j)-}, i < T \\ S_T^{(j)-}, i \ge T. \end{cases}$$

Note that  $W_i^{(j)+}$  corresponds to the upper critical interval, and  $W_i^{(j)-}$  to the lower one. Furthermore, the inequality of Theorem 8 holds if and only if  $W_{i_{after}(k)}^{(j)+} \leq 0$  and  $W_{i_{after}(k)}^{(j)-} \geq 0$  for each j (here we must make use of our assumption that  $S_i^{(\lambda)} = 0$  for all  $\lambda < k$ ). The next Lemma states that, within their respective critical intervals, they are a supermartingale and submartingale respectively:

**Lemma 10.** For all 
$$i \in I_k$$
 and for all  $j \in [k, d-2]$ ,  $\mathbb{E}\left[W_{i+1}^{(j)-} - W_i^{(j)-} \mid G_i\right] \ge 0$  whenever  $W_i^{(j)-} \le E_j(i)$ , and  $\mathbb{E}\left[W_{i+1}^{(j)+} - W_i^{(j)+} \mid G_i\right] \le 0$  whenever  $W_i^{(j)+} \ge -E_j(i)$ .

**Proof.** Here we just prove the first part of the Lemma; the second part follows from nearly identical calculations. By the same logic as in the proof of Lemma 7 we work with  $S^{(j)-}$  instead of  $W^{(j)-}$  and assume that (16) holds for all j. We also have the same expected change as in Lemma 7, except with  $E_{first}(i)$  replaced with  $4E_{i}(i)$ :

$$\mathbb{E}\left[S_{i+1}^{(j)-} - S_{i}^{(j)-} \mid G_{i}\right] = \frac{-2Y_{i}^{(j)}}{S_{i}^{(d-1)}} + O\left(\frac{1}{dn-2i}\right) + \frac{2y_{j}(i/n)}{s_{d-1}(i/n)} - \frac{d^{2}}{d\mu^{2}}\left(\frac{ns_{j}(\mu/n)}{2}\right)\Big|_{\mu=\psi} + 4(E_{j}(i+1) - E_{j}(i)).$$

We split the above expression (excluding  $O\left(\frac{1}{dn-2i}\right)$ ) into three summands, assuming  $S_i^{(j)-} \le E_j(i) \iff S_i^{(j)} - ns_j(i/n) \le -3E_j(i)$  (for convenience, for the case j=0, we put  $S_i^{(j-1)}$ ,  $s_{j-1}$ , and  $E_{j-1}$  all equal to 0):

1.

$$\begin{split} \frac{-2Y_i^{(j)}}{S_i^{(d-1)}} + \frac{2y_j(i/n)}{s_{d-1}(i/n)} &= \frac{-2S_i^{(j)} + 2ns_j(i/n) + 2S_i^{(j-1)} - 2ns_{j-1}(i/n)}{ns_{d-1}(i/n)} \\ &+ 2Y_i^{(j)} \left(\frac{1}{ns_{d-1}(i/n)} - \frac{1}{S_i^{(d-1)}}\right) \\ &\geq \frac{6E_j(i) - 8E_{j-1}(i)}{ns_{d-1}(i/n)} - \frac{8S_i^{(j)}E_{d-1}(i)}{S_i^{(d-1)}(ns_{d-1}(i/n))} \quad \text{by (16)} \\ &\geq \frac{5.9E_j(i)}{ns_{d-1}(i/n)} - \frac{9S_i^{(j)}E_{d-1}(i)}{(ns_{d-1}(i/n))^2} \quad \text{by (15), (16), (11), and } i \leq i_{after}(k) \\ &\geq \frac{5.9E_j(i)}{ns_{d-1}(i/n)} - \frac{9(ns_j(i/n))E_{d-1}(i)}{(ns_{d-1}(i/n))^2} - \frac{36E_j(i)E_{d-1}(i)}{(ns_{d-1}(i/n))^2} \quad \text{by (16)} \\ &= \left(\frac{E_j(i)}{ns_{d-1}(i/n)}\right) \left(5.9 - 9\left(\frac{s_j(i/n)}{s_{d-1}(i/n)}\right)^{0.3} - \frac{36 * 2^k \ln{(n)^{0.05}}}{(ns_{d-1}(i/n))^{0.3}}\right) \\ &\geq \frac{5.8E_j(i)}{ns_{d-1}(i/n)} \quad \text{by } i \leq i_{after}(k) \text{ and (11).} \end{split}$$

2. Just as in the proof of Lemma 7:

$$-\frac{d^2}{d\mu^2} \left( \frac{n s_j(\mu/n)}{2} \right) \Big|_{\mu=\psi} = O\left( \frac{1}{dn-2i} \right).$$

3.

$$4(E_{j}(i+1) - E_{j}(i)) = 4\frac{dE_{j}(\mu)}{d\mu} \Big|_{\mu=\phi} \quad \text{for some } \phi \in [i, i+1]$$

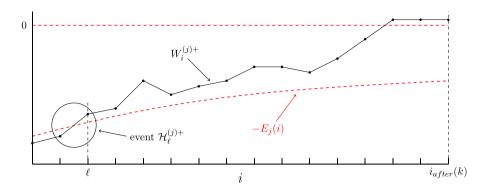
$$= (4)(2^{k}) \ln(n)^{0.05} \left(\frac{0.7}{(ns_{j}(\phi/n))^{0.3}}\right) \left(\frac{-2y_{j}(\phi/n)}{s_{d-1}(\phi/n)}\right) \quad \text{by (5)}$$

$$= (4+o(1))(2^{k}) \ln(n)^{0.05} \left(\frac{0.7}{(ns_{j}(\phi/n))^{0.3}}\right) \left(\frac{-2s_{j}(\phi/n)}{s_{d-1}(\phi/n)}\right) \quad \text{by (10)}$$

$$= \frac{-(5.6+o(1))E_{j}(\phi)}{ns_{d-1}(\phi/n)} = \frac{-(5.6+o(1))E_{j}(i)}{ns_{d-1}(i/n)}. \quad (17)$$

Now we put the above bounds together (using (11), (15), and  $i \le i_{after}(k) \le i_{after}(j)$ ):

$$\mathbb{E}\left[S_{i+1}^{(j)-} - S_i^{(j)-} \mid G_i\right] \ge \frac{0.01E_j(i)}{ns_{d-1}(i/n)} + O\left(\frac{1}{dn - 2i}\right) \\ \ge 0.$$



**Figure 1.** Visual representation of event  $\mathcal{E}_{\ell}^{(j)+}$ .

We introduce the next Lemma to get sufficiently small bounds on the one-step changes in each time step (this is known as the "bounded hypothesis" from [14]):

**Lemma 11.** For all  $i \in I_k$  and all  $j \in [k, d-2]$ ,

$$-3 < W_{i+1}^{(j)\xi} - W_i^{(j)\xi} < \ln(n)^{-d+1.06+j}$$

where "\xi" can be either "+" or "-".

**Proof.** Like in the proofs of Lemma 7 and 10, we assume that  $W^{\xi} = S^{(j)\xi}$  ( $\xi$  is + or -), else  $W_{i+1}^{\xi} - W_i^{\xi} = 0$ . Again, we have  $-2 \le S_{i+1}^{(j)} - S_i^{(j)} \le 0$ . Secondly, we have

$$\begin{aligned} &|-ns_{j}((i+1)/n) + ns_{j}(i/n) - CE_{j}(i+1) + CE_{j}(i)| \\ &\leq |-ns_{j}((i+1)/n) + ns_{j}(i/n)| + |-CE_{j}(i+1) + CE_{j}(i)| \\ &= O\left(\frac{y_{j}(i/n)}{s_{d-1}(i/n)} + \frac{E_{j}(i)}{ns_{d-1}(i/n)}\right) \quad \text{by (5), (11), and (17)} \\ &= o\left(\ln(n)^{-d+1.06+j}\right) \quad \text{by (11), (15), and } i \leq i_{after}(k) \leq i_{after}(j). \end{aligned}$$

Combining the inequalities completes the proof.

To put this all together to prove the first part of Theorem 8, we introduce a series of events: first, let  $\mathcal{E}^{(j)+}$  denote the event that  $W^{(j)+}_{i_{after}(k)} > 0$  and  $\mathcal{E}^{(j)-}$  denote the event that  $W^{(j)-}_{i_{after}(k)} < 0$ . Let  $\mathcal{E} = \left(\bigcup_{j \geq k} \mathcal{E}^{(j)+}\right) \cup \left(\bigcup_{j \geq k} \mathcal{E}^{(j)-}\right)$ ; we seek to bound  $\mathbb{P}[\mathcal{E}]$ , since  $\mathcal{E}$  is the event that (16) *doesn't* hold for some  $i \in I_k$ . Next, for all  $\ell \in I_k$ , let  $\mathcal{H}^{(j)+}_{\ell}$  be the event that  $W^{(j)+}_{\ell-1} < -E_j(\ell-1)$  and  $W^{(j)+}_{\ell} \geq -E_j(\ell)$ , and let

$$\mathcal{E}_{\ell}^{(j)+} := \mathcal{H}_{\ell}^{(j)+} \cap \left\{ W_i^{(j)+} \geq -E_j(i) \text{ for all } i \geq \ell \right\} \cap \left\{ W_{i_{after(k)}}^{(j)+} > 0 \right\}.$$

(see Fig. 1 for a visual representation of event  $\mathcal{E}_{\ell}^{(j)+}$ )

Similarly, for all  $\ell \in I_k$ , let  $\mathcal{H}_{\ell}^{(j)-}$  be the event that  $W_{\ell-1}^{(j)-} > E_j(\ell-1)$  and  $W_{\ell}^{(j)-} \le E_j(\ell)$ , and let

$$\mathcal{E}_{\ell}^{(j)-} := \, \mathcal{H}_{\ell}^{(j)-} \cap \left\{ W_i^{(j)-} \leq E_j(i) \text{ for all } i \geq \ell \right\} \cap \left\{ W_{i_{after}(k)}^{(j)-} < 0 \right\}.$$

Finally, note that, by Lemma 9, with high probability we must have

$$W_{i_{after}(k-1)+1}^{(j)+} < -E_j(i_{after}(k-1)+1) \text{ and } W_{i_{after}(k-1)+1}^{(j)-} > E_j(i_{after}(k-1)+1).$$

Furthermore, assuming these two inequalities hold (and, once again, assuming that  $S_i^{\lambda} = 0$  if  $\lambda < k$ ), then if  $W_{i_{after}(k)}^{(j)+} > 0$  for some j, one of the events  $\mathcal{E}_{\ell}^{(j)+}$  must happen; likewise, if  $W_{i_{after}(k)}^{(j)-} < 0$  for some j, one of the events  $\mathcal{E}_{\ell}^{(j)-}$  must happen; hence,  $\mathcal{E}^{(j)+} = \bigcup_{\ell} \mathcal{E}_{\ell}^{(j)+}$  and  $\mathcal{E}^{(j)-} = \bigcup_{\ell} \mathcal{E}_{\ell}^{(j)-}$ .

We are now ready to prove the first statement of Theorem 8 in full.

**Proof of the first part of Theorem** 8 *with fixed k.* First, we fix an arbitrary j (in [k, d-2]). We prove that  $\mathbb{P}[\mathcal{E}^{(j)-}] = \exp\{-\Omega(\ln(n)^{0.036})\}$ ; the proof for bounding  $\mathbb{P}[\mathcal{E}^{(j)+}]$  is nearly identical. We will use Corollary 5 to bound  $\mathbb{P}[\mathcal{E}^{(j)-}_{\ell}]$  for each fixed  $\ell$ . Given a fixed  $\ell$ , we define a modified stopping time

$$T_{mod} := \min_{i \in [\ell, i_{after}(k)]} \left\{ W_i^{(j)-} > E_j(i) \text{ or } i = T \right\}$$

(letting  $T_{mod} = \infty$  if the condition doesn't hold for any i in the range). Let variable  $W_i^{\ell}$  be the variable  $W_i^{(j)-}$  defined just on  $i \in [\ell, i_{after}(k)]$  equipped with this stopping time (we drop the "(j)-" here for convenience); i.e.

$$W_i^{\ell} := \begin{cases} W_i^{(j)-}, i < T_{mod} \\ W_{T_{mod}}^{(j)-}, i \ge T_{mod}. \end{cases}$$

Note that  $(W_i^{\ell})_i$  (over  $i \in [\ell, i_{after}(k)]$ ) is a submartingale by Lemma 10, since our new stopping time negates the need for the condition  $W_i^{(j)-} \leq E_j(i)$ ; also,  $(W_i^{\ell})_i$  satisfies Lemma 11. Since we want an upper bound for  $\mathbb{P}[\mathcal{E}_{\ell}^{(j)-}]$ , we can condition on event  $\mathcal{H}_{\ell}^{(j)-}$ , as  $\mathcal{H}_{\ell}^{(j)-} \supseteq \mathcal{E}_{\ell}^{(j)-}$ . Now let

$$A_i = -W_{\ell+i}^{\ell} + W_{\ell}^{\ell},$$
  
 $\eta = \ln(n)^{-d+1.06+j},$   
 $N = 3,$   
 $m = i_{after}(k) - \ell,$   
 $a = 0.9E_i(\ell).$ 

Note that the conditions of Corollary 5 are satisfied:  $0 = A_0$  and  $\eta < N/10$  are obvious, Lemma 11 gives us  $-\eta \le A_{i+1} - A_i \le N$ , and  $(A_i)_i$  is a supermartingale since  $(W_i^\ell)_i$  is a submartingale. We therefore implement Corollary 5, using  $m \le \frac{dn}{2} - \ell \le dns_{d-1}(\ell/n)$  (by (9)), (11), and (15):

$$\mathbb{P}[A_m \ge a] \le e^{-\frac{a^2}{3\eta Nm}} + e^{-\frac{a}{6N}} = e^{-\Omega(\ln(n)^{0.04}(ns_j(\ell/n))^{0.4})} + e^{-\Omega(\ln(n)^{0.05}(ns_j(\ell/n))^{0.7})}.$$
 (18)

To bound  $\mathbb{P}\left[\mathcal{E}_{\ell}^{(j)-}\right]$ , we show that  $\mathcal{E}_{\ell}^{(j)-}\subseteq \{A_m\geq a\}$  and apply (18) while conditioning on  $\mathcal{H}_{\ell}^{(j)-}$ . Given  $\mathcal{H}_{\ell}^{(j)-}$  happens, we have  $W_{\ell}^{\ell}=W_{\ell}^{(j)-}>0.9E_{j}(\ell)=a$  by (15), Lemma 11, and  $i\leq i_{after}(j)$ . Therefore  $\mathcal{E}_{\ell}^{(j)-}=\mathcal{H}_{\ell}^{(j)-}\cap\left\{W_{i_{after}(k)}^{\ell}<0\right\}\subseteq\{A_m\geq a\}$ , hence

$$\mathbb{P}\left[\mathcal{E}_{\ell}^{(j)-}\right] = e^{-\Omega\left(\ln{(n)^{0.04}(ns_{j}(\ell/n))^{0.4}}\right)} + e^{-\Omega\left(\ln{(n)^{0.05}(ns_{j}(\ell/n))^{0.7}}\right)}.$$

We now take a union bound to bound  $\mathbb{P}[\mathcal{E}^{(j)-}]$  (using (11) where appropriate):

$$\begin{split} & \mathbb{P}[\mathcal{E}^{(j)-}] \leq \sum_{\ell=i_{after}(k-1)+1}^{i_{after}(k)} \mathbb{P}\left[\mathcal{E}_{\ell}^{(j)-}\right] \\ & = \sum_{\ell=i_{trans}}^{i_{after}(k)} \left( \exp\left\{-\Omega\left(\ln\left(n\right)^{0.04}(ns_{j}(\ell/n))^{0.4}\right)\right\} + \exp\left\{-\Omega\left(\ln\left(n\right)^{0.05}(ns_{j}(\ell/n))^{0.7}\right)\right\} \right) \\ & = \sum_{\ell=i_{trans}}^{i_{after}(j)} \left( \exp\left\{-\Omega\left(\frac{(dn-2\ell)^{0.4}}{\ln\left(n\right)^{0.4d-0.44-0.4j}}\right)\right\} + \exp\left\{-\Omega\left(\frac{(dn-2\ell)^{0.7}}{\ln\left(n\right)^{0.7d-0.75-0.7j}}\right)\right\} \right) \\ & = \sum_{p=\lfloor \ln\left(n\right)^{d-1.01-j}}^{\lceil n^{1-1/(100d)} \rceil} \left( \exp\left\{-\Omega\left(\frac{p^{0.4}}{\ln\left(n\right)^{0.4d-0.44-0.4j}}\right)\right\} + \exp\left\{-\Omega\left(\frac{p^{0.7}}{\ln\left(n\right)^{0.7d-0.75-0.7j}}\right)\right\} \right) \\ & = \ln\left(n\right)^{d-1.01-j} \sum_{q=1}^{\infty} \left( \exp\left\{-\Omega\left(q^{0.4}\ln\left(n\right)^{0.036}\right)\right\} + \exp\left\{-\Omega\left(q^{0.7}\ln\left(n\right)^{0.043}\right)\right\} \right) \\ & = \exp\left\{-\Omega\left(\ln\left(n\right)^{0.036}\right)\right\}. \end{split}$$

We give a note for the aspects of the proof of bounding  $\mathbb{P}[\mathcal{E}^{(j)+}]$  that are different from the above: use the variable  $W_i^{(j)+}$  instead of  $W_i^{(j)-}$ , events  $\mathcal{E}_\ell^{(j)+}$  instead of  $\mathcal{E}_\ell^{(j)-}$ , and  $\mathcal{H}_\ell^{(j)+}$  instead of  $\mathcal{H}_\ell^{(j)-}$ . Define  $T_{mod}$  instead as

$$T_{mod} := \min_{i \in [\ell, i_{after}(k)]} \left\{ W_i^{(j)+} < -E_j(i) \text{ or } i = T \right\}.$$

Finally, use Corollary 4 instead of Corollary 5 (which will be slightly easier to implement).  $\Box$ 

## 5. Final phase

We continue our proof by induction of Theorem 8 with our fixed index k; now we prove the second part. We assume the first part of Theorem 8 to hold, as well as the second part of the Theorem for lesser k; for example, we have  $S_{i_{after}(k-1)}^{(k-1)} = 0$  w.h.p. In this section we focus on the d-process for a narrow domain of i. Let

$$i_{before}(k) := \left\lfloor \frac{dn}{2} - \ln(n)^{d-0.8-k} \right\rfloor.$$

We will consider the d-process starting at step  $i_{before}(k)$  assuming that (16) holds at  $i = i_{before}(k)$ ; we do not need the first part of Theorem 8 in this section otherwise. We do not use martingale arguments here, but rather we show that the distribution of the sequence of time steps at which a vertex of degree k is chosen from the d-process is similar to a uniform distribution over all possible such sequences. Theorem 8, (10), and (15) tell us that w.h.p. we will have  $\sim \frac{2(d-1)!}{k!} \ln{(n)^{0.2}}$  vertices of degree at most k (or degree equal to k; they are the same here) left when there are  $\lfloor \ln{(n)^{d-0.8-k}} \rfloor$  steps left; hence, the average distance between steps at which we remove vertices of degree k is  $\frac{k!}{2(d-1)!} \ln{(n)^{d-1-k}}$ . When there are this many steps left times r, we expect r such

vertices to remain, and for the probability that there are no vertices of degree k to be  $e^{-r}$ . Most of this section will build towards proving the following Theorem:

**Theorem 12.** Let L(n) be an integer-valued function so that  $L(n) = \Theta(\ln(n)^{0.2})$  and let  $J(n) = \lfloor \frac{dn}{2} \rfloor - i_{before}(k) \sim \ln(n)^{d-0.8-k}$ . Let H be any graph with  $i_{before}(k)$  edges which satisfies (16) at  $i = i_{before}(k)$ , has no vertices of degree at most k-1, and has L(n) vertices of degree k. Also, let  $r \in \mathbb{R}^+$  be arbitrary. Then

$$\mathbb{P}\left[S_{\lfloor \frac{dn}{2} - \frac{rJ(n)}{L(n)} \rfloor}^{(k)} = 0 \mid G_{i_{before}(k)} = H\right] \to e^{-r}.$$

First, we note that, given that (16) holds for  $i = i_{before}(k)$  and by (1), that w.h.p.  $dn - 2i_{before}(k) - S_{i_{before}(k)}^{(d-1)} = O\left(\frac{dn - 2i_{before}}{\ln{(n)}}\right)$  (consider  $S_{i_{before}(k)}^{(d-2)}$ ); hence, for all  $i \in [i_{before}(k), i_{after}(k)]$ :

$$S_{i}^{(d-1)} = dn - 2i + O\left(\frac{dn - 2i_{before}}{\ln(n)}\right) = (dn - 2i)\left(1 + O\left(\frac{1}{\ln(n)^{0.79}}\right)\right). \tag{19}$$

Let  $t_{start} = i_{before}(k)$  and  $t_{end} = \lfloor dn/2 - rJ(n)/L(n) \rfloor$ . Consider the d-process between  $t_{start}$  and  $t_{end}$ , given that  $G_{t_{start}} = H$ . At each step two vertices are chosen; now assume that the pair at each step is ordered uniformly at random, so that a sequence of  $2(t_{end} - t_{start})$  vertices is generated. We also generate a binary sequence simultaneously, each digit corresponding to a vertex: after a pair of vertices is picked for the vertex sequence, for each of the two vertices (in the order that they are randomly shuffled) append a "1" to the binary sequence if the corresponding vertex had degree k just before it was picked, and append a "0" otherwise. Let  $\mathcal{P}:\{0,1\}^{2(t_{end}-t_{start})} \rightarrow [0,1]$  be the corresponding probability function that arises from this process (note that, if  $\gamma$  is a string with more than L(n) "1"s, then  $\mathcal{P}(\gamma) = 0$ ). Note that  $\mathcal{P}$  depends on the graph H. We compare this to a second probability function  $\mathcal{Q}:\{0,1\}^{2(t_{end}-t_{start})} \rightarrow [0,1]$ , which is defined by picking a binary string with L(n) 1's and 2J(n) - L(n) 0's uniformly at random, then taking the first  $2(t_{end} - t_{start})$  digits.

For any binary sequence  $\gamma$  with  $\ell$  digits, and  $I \subset [\ell]$ , let  $\gamma_I$  be the subsequence with indices from I; for example,  $\gamma_{[a]}$  would be the first a digits of  $\gamma$ , and  $\gamma_{\{a\}}$  would just be the a-th digit (for notation's sake, let " $\gamma_{[0]}$ " be the empty string). Also let  $\|\gamma\|$  denote the number of 1's in  $\gamma$ . We now present the following Lemma:

**Lemma 13.** Let  $\alpha$  be an arbitrary  $2(t_{end} - t_{start})$  length binary sequence with at most L(n) 1's, and let  $\gamma$  be the random binary sequence according to either  $\mathcal{P}$  or  $\mathcal{Q}$ . Let  $i \in [2(t_{end} - t_{start})]$ . Then (letting  $a_{\{0\}} = 1$  for sake of notation):

$$\begin{split} & \underbrace{\mathbb{P}_{\mathcal{P}}[\gamma_{[i]} = \alpha_{[i]} \mid \gamma_{[i-1]} = \alpha_{[i-1]}]}_{\mathbb{P}_{\mathcal{Q}}[\gamma_{[i]} = \alpha_{[i]} \mid \gamma_{[i-1]} = \alpha_{[i-1]}]} \begin{cases} = 1 + O\left(\frac{1}{J(n)\ln(n)^{0.39}}\right) & \text{if } \alpha_{\{i\}} = 0 \text{ and } \alpha_{\{i-1\}} = 0 \\ = 1 + O\left(\frac{\ln(n)^{0.4}}{J(n)}\right) & \text{if } \alpha_{\{i\}} = 0 \text{ and } \alpha_{\{i-1\}} = 1 \\ = 1 + O\left(\frac{1}{\ln(n)^{0.79}}\right) & \text{if } \alpha_{\{i\}} = 1 \text{ and } \alpha_{\{i-1\}} = 0 \\ \leq 1 + O\left(\frac{1}{\ln(n)^{0.79}}\right) & \text{if } \alpha_{\{i\}} = 1 \text{ and } \alpha_{\{i-1\}} = 1. \end{cases} \end{split}$$

**Proof.** First, we consider the cases where  $\alpha_{\{i\}} = 1$ . We have

$$\mathbb{P}_{\mathcal{Q}}[\gamma_{\{i\}} = 1 \mid \gamma_{[i-1]} = \alpha_{[i-1]}] = \frac{L(n) - \|\alpha_{[i-1]}\|}{2J(n) - (i-1)}.$$
 (20)

For the probability space  $\mathcal{P}$ , we need to consider three subcases: we need to consider whether i is even or odd, and if it is even, whether  $\alpha_{\{i-1\}}$  is 0 or 1, since each step of the d-process outputs

two digits of the binary string. Let's say that  $\tau$  corresponds to the last step in the d-process before the i-th binary digit is generated (recall that pairs of digits are generated together). Then if i is odd:

$$\mathbb{P}_{\mathcal{P}}[\gamma_{\{i\}} = 1 \mid \gamma_{\{i-1\}}] = -\frac{1}{2} \mathbb{E} \left[ S_{\tau+1}^{(k)} - S_{\tau}^{(k)} \mid G_{\tau} \right] \\
= \frac{S_{\tau}^{(k)}}{S_{\tau}^{(d-1)}} \left( 1 + O\left(\frac{1}{dn - 2\tau}\right) \right) \quad \text{by (3)} \\
= \frac{S_{\tau}^{(k)}}{2 \lfloor dn/2 - \tau \rfloor} \left( 1 + O\left(\frac{1}{\ln(n)^{0.79}}\right) \right) \quad \text{by (19)} \\
= \frac{L(n) - \|\alpha_{\{i-1\}}\|}{2J(n) - (i-1)} \left( 1 + O\left(\frac{1}{\ln(n)^{0.79}}\right) \right). \tag{21}$$

If i is even and  $\alpha_{\{i-1\}} = 1$ , then  $S_{\tau}^{(k)} = L(n) - \|\alpha_{\{i-1\}}\| + 1$ . At step  $\tau$  there are  $S_{\tau}^{(k)} \left(S_{\tau}^{(d-1)} + O(1)\right)$  ordered pairs of vertices whose first vertex has degree k, and  $at \ most \ 2\binom{S_{\tau}^{(k)}}{2}$  ordered pairs of vertices both with degree k; hence:

$$\mathbb{P}_{\mathcal{P}}[\gamma_{\{i\}} = 1 \mid \gamma_{[i-1]} = \alpha_{[i-1]}] \leq \frac{S_{\tau}^{(k)} - 1}{S_{\tau}^{(d-1)} + O(1)} \\
= \frac{S_{\tau}^{(k)} - 1}{2 \lfloor dn/2 - \tau \rfloor} \left( 1 + O\left(\frac{1}{\ln(n)^{0.79}}\right) \right) \quad \text{by (19)} \\
= \frac{L(n) - \|\alpha_{[i-1]}\|}{2I(n) - (i-1)} \left( 1 + O\left(\frac{1}{\ln(n)^{0.79}}\right) \right). \tag{22}$$

Hence the final inequality of the Lemma holds by (21) and (22).

Next, consider the case where i is even and  $\alpha_{\{i-1\}} = 0$ ; here,  $S_{\tau}^{(k)} = L(n) - \|\alpha_{[i-1]}\|$  once again. At step  $\tau$  there are  $\left(S_{\tau}^{(d-1)} - S_{\tau}^{(k)}\right) \left(S_{\tau}^{(d-1)} + O(1)\right)$  ordered pairs of vertices whose first vertex has degree greater than k, and  $S_{\tau}^{(k)} \left(S_{\tau}^{(d-1)} - S_{\tau}^{(k)} + O(1)\right)$  ordered pairs of vertices for which the first vertex has degree greater k and the second vertex has degree k (one can "pick the second vertex first" to see this). Hence:

$$\mathbb{P}_{\mathcal{P}}[\gamma_{\{i\}} = 1 \mid \gamma_{[i-1]} = \alpha_{[i-1]}] = \frac{S_{\tau}^{(k)}}{S_{\tau}^{(d-1)}} \left( 1 + O\left(\frac{1}{S_{\tau}^{(d-1)}}\right) \right) \quad \text{since } S_{\tau}^{(d-1)} \gg S_{\tau}^{(k)} \text{ there}$$

$$= \frac{S_{\tau}^{(k)}}{2 \lfloor dn/2 - \tau \rfloor} \left( 1 + O\left(\frac{1}{\ln(n)^{0.79}}\right) \right) \quad \text{by (19)}$$

$$= \frac{L(n) - \|\alpha_{[i-1]}\|}{2J(n) - (i-1)} \left( 1 + O\left(\frac{1}{\ln(n)^{0.79}}\right) \right), \tag{23}$$

hence the third equality of the Lemma holds by (21) and (23).

Now consider  $\alpha_{\{i\}} = 0$ . By modifying (20) to accommodate  $\gamma_{\{i\}} = 0$ , we have

$$\mathbb{P}_{\mathcal{Q}}[\gamma_{\{i\}} = 0 \mid \gamma_{[i-1]} = \alpha_{[i-1]}] = 1 - \frac{L(n) - \|\alpha_{[i-1]}\|}{2J(n) - (i-1)}.$$
 (24)

Similarly, by modifying (21) and (23), if  $a_{\{i-1\}} = 0$  then

$$\mathbb{P}_{\mathcal{P}}[\gamma_{\{i\}} = 0 \mid \gamma_{[i-1]} = \alpha_{[i-1]}] = 1 - \frac{L(n) - \|\alpha_{[i-1]}\|}{2J(n) - (i-1)} \left(1 + O\left(\frac{1}{\ln(n)^{0.79}}\right)\right). \tag{25}$$

By modifying (21) and (22), if  $a_{\{i-1\}} = 1$ , then

$$\mathbb{P}_{\mathcal{P}}[\gamma_{\{i\}} = 0 \mid \gamma_{[i-1]} = \alpha_{[i-1]}] \ge 1 - \frac{L(n) - \|\alpha_{[i-1]}\|}{2J(n) - (i-1)} \left(1 + O\left(\frac{1}{\ln(n)^{0.79}}\right)\right) \\
= 1 + O\left(\frac{L(n) - \|\alpha_{[i-1]}\|}{2J(n) - (i-1)}\right).$$
(26)

Since  $L(n) = \Theta(\ln(n)^{0.2})$  and

$$2J(n) - (i-1) = 2\lfloor dn/2 - \tau \rfloor = \Omega(dn/2 - t_{end}) = \Omega(J(n)/\ln(n)^{0.2}),$$

then  $\frac{L(n)-\|\alpha_{[i-1]}\|}{2J(n)-(i-1)}=O\left(\frac{\ln{(n)}^{0.4}}{J(n)}\right)$ . Therefore the ratio of (25) and (24) is  $1+O\left(\frac{1}{J(n)\ln{(n)}^{0.39}}\right)$ , verifying the first inequality of the Lemma, and the ratio of (26) and (24) is  $1+O\left(\frac{\ln{(n)}^{0.4}}{J(n)}\right)$ , verifying the second inequality of the Lemma.

**Proof of Theorem** 12. First, let  $\alpha$  be an arbitrary string which satisfies the criteria in Lemma 13. By using the Lemma 13 recursively:

$$\frac{\mathbb{P}_{\mathcal{P}}[\gamma = \alpha]}{\mathbb{P}_{\mathcal{Q}}[\gamma = \alpha]} \le \exp\left\{ O\left( J(n) \frac{1}{J(n) \ln(n)^{0.39}} + L(n) \left( \frac{1}{\ln(n)^{0.79}} + \frac{\ln(n)^{0.4}}{J(n)} \right) \right) \right\} 
= 1 + o(1),$$
(27)

and if  $\alpha$  is an arbitrary string with no two consecutive 1's which satisfies the criteria in Lemma 13, then by similar logic,

$$\frac{\mathbb{P}_{\mathcal{P}}[\gamma = \alpha]}{\mathbb{P}_{\mathcal{O}}[\gamma = \alpha]} = 1 + o(1). \tag{28}$$

Let  $\mathcal{C}$  be the event that  $\gamma$  has two consecutive 1's; we consider  $\mathbb{P}[\mathcal{C} \mid G_{t_{start}} = H]$ . We consider probability space  $\mathcal{Q}$  first. Recall that  $\alpha$  is a string that has  $\sim 2J(n) = \Omega(\ln(n)^{1.2})$  characters and at most  $L(n) = \Theta(\ln(n)^{0.2})$  1's. Because  $\mathcal{Q}$  is a truncation of a uniform distribution, the probability of having two consecutive 1's will be  $O\left(\frac{(L(n))^2}{J(n)}\right) = O(\ln(n)^{-0.8})$ . Hence, by (27) we must have

$$\mathbb{P}_{\mathcal{P}}[\mathcal{C} \mid G_{t_{start}} = H] = o(1) \text{ and } \mathbb{P}_{\mathcal{Q}}[\mathcal{C} \mid G_{t_{start}} = H] = o(1).$$
(29)

We now combine (28) and (29) to prove Theorem 12 (for ease of notation, assume we are given  $G_{t_{start}} = H$ ):

$$\mathbb{P}\left[S_{\lfloor \frac{dn}{2} - \frac{rJ(n)}{L(n)} \rfloor}^{(k)} = 0\right] = \mathbb{P}_{\mathcal{P}}[\|\gamma\| = L(n)]$$

$$= \mathbb{P}_{\mathcal{P}}[\|\gamma\| = L(n) \mid \mathcal{C}] \mathbb{P}_{\mathcal{P}}[\mathcal{C}] + \mathbb{P}_{\mathcal{P}}[\|\gamma\| = L(n) \mid \overline{\mathcal{C}}] \mathbb{P}_{\mathcal{P}}[\overline{\mathcal{C}}]$$

$$= \mathbb{P}_{\mathcal{Q}}[\|\gamma\| = L(n)] + o(1) \quad \text{by (28) and (29)}$$

$$= \frac{\binom{2^{l(t_{end} - t_{start})}}{L(n)}}{\binom{2J(n)}{L(n)}} + o(1)$$

$$= \frac{\binom{2^{l(t_{end} - t_{start})}}{L(n)}}{\binom{2J(n)}{L(n)}} + o(1)$$

$$= \left(1 - \frac{r(1 + o(1))}{L(n)}\right)^{L(n)} + o(1)$$

$$= e^{-r} + o(1).$$

We can now complete the proof of the second statement of Theorem 8 at value k. Roughly speaking, we will use Theorem 12 with  $L(n) \approx \frac{2(d-1)! \ln{(n)^{0.2}}}{k!}$ , so  $\frac{dn}{2} - i(r_k, k) \approx \frac{r_k J(n)}{L(n)}$ . First, note that  $S_{i_{after}(k)}^{(k)} = 0$  (w.h.p.) comes automatically when the rest of the statement is proved (by putting  $r_\ell = 0$  for  $\ell < k$  and having  $r_k \to 0$ ). Let  $\mathcal{G}_\ell$  be the event that  $S_{\lfloor i(r_\ell,\ell)\rfloor}^{(\ell)} = 0$  and  $\mathcal{G} = \bigcap_{\ell \le k} \mathcal{G}_\ell$ , let  $\mathcal{F}$  be the event that (16) holds for  $i = i_{before}(k)$  and  $S_{i_{before}(k)}^{(j)} = 0$  for j < k, and let  $\mathcal{A} = \mathcal{F} \cap \bigcap_{\ell < k} \mathcal{G}_\ell$ . Also, let

$$\mathcal{I} = [ns_k(i_{before}(k)/n) - 4E_k(i_{before}(k)), ns_k(i_{before}(k)/n) + 4E_k(i_{before}(k))].$$

Note that, by part 1 of Theorem 8, by induction on the second part Theorem 8, and since  $i_{before}(k) > i_{after}(k-1)$ ,  $\mathcal{F}$  happens with probability 1 - o(1). Therefore:

$$\mathbb{P}[\mathcal{G}] = \mathbb{P}\left[\mathcal{G}_k \cap \mathcal{A}\right] + o(1)$$

$$= \sum_{p \in \mathcal{I}} \mathbb{P}\left[\mathcal{G} \mid \mathcal{A} \cap \left(S_{i_{before}(k)}^{(k)} = p\right)\right] \mathbb{P}\left[\mathcal{A} \cap \left(S_{i_{before}(k)}^{(k)} = p\right)\right] + o(1).$$

We can now apply (10), (15), and Theorem 12 to get

$$\mathbb{P}\left[\mathcal{G} \mid \mathcal{A} \cap \left(S_{i_{before}(k)}^{(k)} = p\right)\right] = e^{-r_{\ell}} + o(1)$$

for  $p \in \mathcal{I}$ . We note that all o(1) functions in the sum can be made to be the same by carefully reviewing the proof of Theorem 12. Therefore:

$$\begin{split} \mathbb{P}[\mathcal{G}] &= \sum_{p \in \mathcal{I}} (e^{-r_\ell} + o(1)) \mathbb{P}\left[\mathcal{A} \cap \left(S_{i_{before}(k)}^{(k)} = p\right)\right] + o(1) \\ &= e^{-r_\ell} \sum_{p \in \mathcal{I}} \mathbb{P}\left[\mathcal{A} \cap \left(S_{i_{before}(k)}^{(k)} = p\right)\right] + o(1) \\ &= e^{-r_\ell} \mathbb{P}\left[\bigcap_{\ell < k} \mathcal{G}_\ell\right] + o(1) \qquad \text{by Theorem 8} \\ &= \exp\left\{\sum_{\ell = 0}^k e^{-r_\ell}\right\} + o(1) \qquad \text{by induction on Theorem 8,} \end{split}$$

proving Theorem 8.

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