



Short Communication

Longer-term impacts of sugar-sweetened beverage taxes on fast-food beverage prices: evidence from Oakland, California, 2-year post-tax

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Abstract

Objective: To evaluate the effect of a sugar-sweetened beverage (SSB) tax implemented in Oakland, California, in July 2017, on prices of beverages sold in fast-food restaurants 2-year post-tax.

Design: Using a difference-in-differences (DID) approach, we analysed beverage price data collected from fast-food restaurants 1-month pre-tax and 2-year post-tax in Oakland (intervention site) and Sacramento, California (comparison site). Separate linear regression models were used to estimate the impact of the tax on prices of bottled regular soda, bottled diet soda, bottled unsweetened beverages and fountain drinks.

Setting: Oakland and Sacramento, California, USA.

Participants: Chain and non-chain fast-food restaurants (*n* 85).

Results: DID estimates indicate that in fast-food restaurants, on average, the price of bottled regular soda increased by 1.44 cents/oz (95% CI 0.50, 2.73) (tax pass-through rate of 144%) and the price of bottled diet soda increased by 1.17 cents/oz (95% CI 0.07, 2.13). No statistically significant differences were found between bottled regular and diet soda price increases. Price effects for unsweetened beverages and fountain drinks were not statistically significant. Further, the estimated price change for fountain drinks was nearly zero.

Conclusions: Findings suggest that the effectiveness of SSB taxes in discouraging SSB consumption may be limited in fast-food restaurants in Oakland, California, because there were similar price increases in taxed and untaxed bottled soda and no changes in fountain drink prices.

Keywords

Sugar-sweetened beverages
Excise tax
Tax pass-through
Obesity
Fast food

In recent decades, American household food expenditures have shifted from food at home to food away from home; by 2010, food away from home spending exceeded food at home spending and has continued to increase, particularly in fast-food restaurants⁽¹⁾. Recent estimates suggest that Americans, on average, spend 18% of household food expenditures at fast-food restaurants⁽¹⁾ and, on a given day, 37% frequent a fast-food restaurant⁽²⁾. Consuming fast-food meals, which are high in calories, sodium, saturated fat and sugar⁽³⁾, is associated with higher energy intake and lower diet quality among children and adults^(4,5) and obesity among adults⁽⁶⁾. Further, adult fast-food

consumers obtain approximately 26% of their calories from fast food⁽⁶⁾. US studies have highlighted disparities, noting men, younger adults and non-Hispanic blacks are more likely to consume fast food^(2,6) and adverse impacts of fast-food consumption on diet are larger for lower- and middle-income populations and non-Hispanic blacks^(4,5).

Fast-food consumption is also associated with greater sugar-sweetened beverage (SSB) consumption^(4,7), which is linked to obesity, type 2 diabetes and CVD⁽⁸⁾. SSB are the leading source of added sugar in the American diet^(9,10) and are widely available in fast-food restaurants. In fact, between 2003 and 2016, 14% of SSB consumed in the

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USA were included with a fast-food meal purchase⁽⁶⁾. Fast-food restaurants often employ marketing strategies to encourage SSB purchases. SSB are usually bundled with meals⁽³⁾, and pricing strategies are often used to incentivise larger portion sizes; examples include reducing the price/unit as size increases (i.e. supersized pricing), offering unlimited consumption for a fixed price (e.g. free refills) and charging one price for all sizes⁽¹¹⁾. These practices are likely profitable as soda, for example, is inexpensive to produce, especially when sold as a fountain drink (costing approximately 1 cent/ounce (oz))⁽¹²⁾.

To reduce SSB consumption and improve health outcomes, SSB taxes are increasingly being implemented worldwide⁽¹³⁾. The primary channel through which SSB taxes aim to discourage consumption is by raising SSB prices. Additionally, public health campaigns associated with SSB taxes may raise awareness of the detrimental health effects of SSB, and tax revenue may fund public health initiatives. Studies evaluating tax pass-through, the extent to which prices increase as the result of a tax, have mostly examined food stores and have found variation in pass-through rates, including by store type, beverage type and package size⁽¹⁴⁾. To our knowledge, only four studies have estimated pass-through to beverages sold in fast-food restaurants and none has done so outside the USA. A Boulder, Colorado, study estimated a pass-through rate of 49 and 69 % to fountain drink prices at 1- and 3-month post-tax, respectively⁽¹⁵⁾. Two Seattle, Washington, studies estimated an overall pass-through rate of 82 and 113 % to taxed beverages at 6-month and 1-year post-tax, respectively, and found no significant changes in prices of untaxed beverages^(16,17). Most recently, an Oakland, California, study estimated 82 % pass-through to bottled regular soda but found no impact for fountain drinks at 1-year post-tax⁽¹⁸⁾.

This study is the first to provide longer-run evidence of SSB tax pass-through in fast-food restaurants. Specifically, we estimated price changes of bottled regular soda, bottled diet soda, bottled unsweetened beverages and fountain drinks 2 years after a 1 cent/oz excise tax on SSB (≥ 25 kcal/12 fluid oz) was implemented in Oakland, California, on 1 July 2017.

Methods

Study sample and measures

Study sample and measures are summarised in this section; for further details see Marinello *et al.*⁽¹⁸⁾. Chain and non-chain fast-food restaurants were selected with geographic random sampling in Oakland, California, the intervention site, and Sacramento, California, the comparison site. Data were collected in-person 1-month pre-tax (May/June 2017) and 2-year post-tax (June 2019) using the Beverage Tax Fast-Food Restaurant Observation Form⁽¹⁹⁾, which has high inter-rater reliability⁽²⁰⁾. The price measure was equal to each product's posted price, defined as the

regular price, unless there was a reduced-price sale (other sales were not considered because they usually do not have a constant price/unit).

Data collection and analytical sample

In both sites, baseline data were collected at sixty-five restaurants 1-month pre-tax. At 2-year post-tax, fifty-seven and fifty-four of these restaurants were audited again in Oakland and Sacramento, respectively. Restaurants were not audited 2-year post-tax if data collectors were asked to leave (n 1 Oakland; n 2 Sacramento), no audit form products were available (n 2 Oakland; n 1 Sacramento), the restaurant was closed (n 4 Oakland; n 7 Sacramento) or the restaurant no longer met fast-food restaurant criteria or became a different restaurant (n 1 Oakland; n 1 Sacramento). Analyses were conducted for four beverage types: bottled regular soda, bottled diet soda, bottled unsweetened beverages (hereafter referred to as regular soda, diet soda and unsweetened beverages) and fountain drinks. Data for regular soda, diet soda and unsweetened beverages were collected by brand and size (e.g. Coke 12 oz), while fountain drinks were collected by size (e.g. medium), where oz were restaurant-specific.

In total, 166, 128, 200 and 408 products were available for regular soda, diet soda, unsweetened beverages and fountain drinks, respectively. To obtain the analytical sample, observations were excluded if information necessary for calculating price/oz was missing. Missing price or sales data resulted in the exclusion of 44, 37, 44 and 43 observations for regular soda, diet soda, unsweetened beverages and fountain drinks, respectively; 126 (75 %) of these observations were missing because prices were not shown, and the remainder were missing for unknown reasons. For fountain drinks, information on available oz was carried over across time periods to missing values within the same restaurant and size category. Attempts to collect or confirm data were made via telephone if oz were missing in all time periods or there were inconsistencies over time. After this process, twenty-nine fountain drink observations were excluded due to missing oz.

The primary analysis used a balanced sample (restricted to products with price/oz data at baseline and 2-year post-tax within a given restaurant). Balancing eliminated differences in product composition, which could affect price/oz (see sample sizes in Table 2). A secondary analysis used the unbalanced sample (see sample sizes in Table 3). While the unbalanced, unsweetened beverage sample included water, milk, 100 % juice and unsweetened tea, the balanced sample only included water and milk. Restaurant characteristics of the balanced analytical samples are provided in Table 1.

Statistical analysis

A difference-in-differences analysis of the intervention site (Oakland) relative to the comparison site (Sacramento) was



used to estimate the causal effect of the tax on beverage prices in fast-food restaurants. The primary assumption required for an unbiased estimate is that, in the absence of the tax, the difference in beverage prices between Oakland and Sacramento would have remained constant during the post-tax period (i.e. parallel trends). Previous work assessed this assumption by testing differences in soda price trends in food stores between Oakland and Sacramento 1-year pre-tax; results revealed no statistically significant differences⁽¹⁸⁾.

Separate linear regression models with bootstrapped standard errors were estimated for regular soda, diet soda, unsweetened beverages and fountain drinks. Observations were clustered at the restaurant level to account for correlations between products from the same restaurant. Models included an indicator for the treatment period (2-year post-tax) and an interaction between this indicator and an indicator for the intervention site (Oakland), as well as restaurant and product fixed effects. Inclusion of restaurant and product fixed effects eliminated potential bias from unobserved time-constant restaurant and/or product characteristics correlated with price. In the fountain drink model, product fixed effects were computed based on categories <25 oz, 25–34 oz and ≥35 oz. The data were analysed in Stata/SE 15.0 (StataCorp).

Results

Table 1 shows mean price/oz in Oakland and Sacramento at baseline and 2-year post-tax by beverage type for the balanced sample. In both sites and time periods, the mean price/oz was highest for unsweetened beverages, followed by bottled soda, and then fountain drinks.

Difference-in-differences regression estimates for the balanced analytical samples are shown in Table 2. The estimated changes in price for regular and diet soda were 1.44 (95 % CI 0.50, 2.73) and 1.17 (95 % CI 0.07, 2.13) cents/oz, respectively; the regular soda estimate represents a tax pass-through rate of 144 %. The difference in price effects between regular and diet soda was not statistically significant at the 5 % level. Results from the unsweetened beverage model suggest that prices may have increased; however, the estimate was not statistically significant. The estimated tax pass-through to fountain drinks was virtually zero and not statistically significant.

Results for the unbalanced samples, shown in Table 3, were nearly identical for regular soda, diet soda and fountain drinks. For unsweetened beverages, the estimated change in price was substantially larger than the balanced sample estimate (which may have been due to differences in product composition), but remained statistically insignificant.

Table 1 Mean price per ounce of beverages in fast-food restaurants in Oakland, CA, and Sacramento, CA, and characteristics of fast-food restaurants before and 2 years after implementation of the Oakland sugar-sweetened beverage tax

	Oakland, CA								Sacramento, CA								
	Pre-tax				2-year post-tax				Pre-tax				2-year post-tax				
	Mean	SD	n	%	Mean	SD	n	%	Mean	SD	n	%	Mean	SD	n	%	
Price (¢/oz)*																	
Bottled regular soda	10.69	2.17			12.51	2.49			9.42	2.01			9.80	1.90			
Bottled diet soda	10.68	1.96			12.30	1.98			9.52	1.98			9.97	2.29			
Bottled unsweetened beverages	14.09	6.00			14.32	5.79			16.76	5.95			16.28	5.17			
Fountain drinks	8.05	2.90			8.59	3.12			8.29	2.21			8.89	2.45			
Restaurant characteristics†																	
Chain restaurant‡			23	61			23	61			27	57			27	57	
Restaurant type																	
Burger and fries			12	32			12	32			12	26			12	26	
Mexican/Latin American			5	13			5	13			10	21			10	21	
Fried chicken/fried fish			7	18			7	18			3	6			3	6	
Sandwich			6	16			6	16			12	26			12	26	
Pizzeria/Italian			4	11			4	11			6	13			6	13	
Chinese/Pan-Asian			3	8			3	8			2	4			2	4	
Other			1	3			1	3			2	4			2	4	
Free water accessible to customers			17	46			20	53			34	72			22	48	
Has fountain machine			27	71			27	71			40	85			40	85	
Free refills offered			18	67			18	67			34	87			17	44	
Self-serve machine			18	67			18	67			36	92			35	90	

*The number of bottled regular soda products (with number of restaurants in parentheses) per time period was 20 (15) and 16 (15) in Oakland and Sacramento, respectively. The number of bottled diet soda products (with number of restaurants in parentheses) per time period was 14 (11) and 8 (8) from Oakland and Sacramento, respectively. The number of bottled unsweetened products (with number of restaurants in parentheses) per time period was 8 (6) and 27 (13) from Oakland and Sacramento, respectively. The number of fountain drink products (with number of restaurants in parentheses) per time period was 58 (26) and 88 (38) in Oakland and Sacramento, respectively.

†Due to some missing data on restaurant characteristics, denominators are lower for some items. Summary statistics on restaurant characteristics are shown for restaurants included in any of the four analyses: 38 in Oakland and 47 in Sacramento, respectively.

‡Restaurants were defined as a chain if they offered franchise opportunities or had corporate headquarters.

Table 2 Balanced sample difference-in-differences tax pass-through estimates 2-year post-tax in fast-food restaurants with 95 % CI†

Beverage type	<i>n</i> Products	<i>n</i> Restaurants	Change in price in Oakland relative to Sacramento (¢/oz)	95 % CI
Bottled regular soda	72	30	1.44*	0.50, 2.73
Bottled diet soda	44	19	1.17*	0.07, 2.13
Bottled unsweetened beverages	70	19	0.71	-0.18, 2.30
Fountain drinks	292	64	-0.06	-0.49, 0.38

*Estimates are significant at the $P < 0.05$ level.

†Each row contains results and sample sizes from separate difference-in-differences linear regressions by beverage category, controlling for restaurant and product fixed effects. Standard errors were bootstrapped and clustered on restaurant, and bias-corrected CI are shown. There are a total of 85 unique restaurants across the four analytical samples. Of these restaurants, 35 % (n 30), 22 % (n 19), 22 % (n 19) and 75 % (n 64) were included in the regular soda, diet soda, unsweetened beverage and fountain drink models, respectively.

Table 3 Unbalanced sample difference-in-differences tax pass-through estimates 2-year post-tax in fast-food restaurants with 95 % CI†

Beverage type	<i>n</i> Products	<i>n</i> Restaurants	Change in price in Oakland relative to Sacramento (¢/oz)	95 % CI
Bottled regular soda	122	59	1.43*	0.39, 2.57
Bottled diet soda	91	54	1.08*	0.00, 2.08
Bottled unsweetened beverages	156	55	1.46	-0.19, 3.68
Fountain drinks	336	79	0.02	-0.44, 0.46

*Estimates are significant at the $P < 0.05$ level.

†Each row contains results and sample sizes from separate difference-in-differences linear regressions by beverage category, controlling for restaurant and product fixed effects. Standard errors were bootstrapped and clustered on restaurant, and bias-corrected CI are shown.

Discussion

This study is the first to estimate longer-term impacts of any SSB tax on prices of beverages sold in fast-food restaurants, which are a major source of SSB. Two years following the implementation of the Oakland, California, tax, regular soda prices had increased by 1.44 cents/oz (144 % tax pass-through) and diet soda prices increased by 1.17 cents/oz. No statistically significant changes in unsweetened beverage or fountain drink prices were found; additionally, the fountain drink estimate was almost zero.

Some differences were found between this study and previous findings on tax pass-through in fast-food restaurants. Compared with our study findings for Oakland at 1-year post-tax⁽¹⁸⁾, pass-through to regular soda increased substantially (82 to 144 %), the price increase of diet soda became statistically significant and pass-through to fountain drinks remained non-statistically significant. The Boulder study found partial pass-through to fountain drink prices⁽¹⁵⁾, whereas we did not. Both Seattle studies found evidence of pass-through to taxed beverages but no evidence of increases in prices of untaxed beverages, whereas we found higher prices for untaxed diet soda^(16,17). However, comparisons to the Seattle studies are difficult because those studies did not estimate pass-through by beverage type or separately for fountain drinks. The present findings are similar to an Oakland food store evaluation, which estimated 127 % pass-through to regular soda, a price increase of 0.78 cents/oz for diet soda, and no statistically significant changes in single-serving unsweetened beverage prices⁽²¹⁾. However, our findings differ from another Oakland food store evaluation that only found partial pass-through to regular soda (63 %)⁽²²⁾.

This study highlights important considerations for understanding the effectiveness of the Oakland SSB tax in reducing SSB consumption at fast-food restaurants. First, it appears that restaurants are not price differentiating between regular and diet soda (since prices of both increased post-tax implementation), which eliminates any potential price incentive for consumers to substitute from calorically sweetened soda (taxed) to non-calorically sweetened soda (untaxed). However, the increase in regular and diet soda prices may encourage consumers to switch to unsweetened beverages, such as bottled water. Second, there is no indication that restaurants are increasing fountain drink prices. Given fountain drink profit margins are exceedingly high⁽¹²⁾, restaurants may be absorbing the tax. It is also possible that restaurants are spreading the tax to non-beverage menu items.

This study has a number of strengths including a long follow-up period and the use of a matched comparison site, which allowed us to account for secular trends and time-constant confounders. Nevertheless, this study is limited by small samples of (1) unsweetened beverages, preventing the examination of price effects by beverage type, and (2) diet soda, though we did find a statistically significant effect. Additionally, this study only assessed certain brands and beverage types and did not evaluate potential effects on meal prices. Finally, the study results may not be generalisable to other jurisdictions.

Conclusion

Overall, this study provides evidence on longer-term, 2-year post-tax effects of SSB taxes in fast-food restaurants.



Study results suggest that the Oakland SSB tax raised the prices of regular and diet soda and had no impact on fountain drink prices. Findings highlight the importance of examining longer-term effects and suggest that SSB taxes may not effectively reduce fast-food restaurant SSB consumption. Policymakers may want to consider complementary restaurant measures to reduce consumption; examples include imposing higher taxes on fountain drink syrup, restricting fountain drink cup sizes and banning free refills.

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References

- Saksena MJ, Okrant AM, Anekwe TD *et al.* (2018) *America's Eating Habits: Food Away from Home*. Washington, DC: United States Department of Agriculture, Economic Research Service.
- Fryar CD, Hughes JP, Herrick KA *et al.* (2018) Fast food consumption among adults in the United States, 2013–2016. Hyattsville, MD: National Center for Health Statistics. Data Brief Number 322. <https://www.cdc.gov/nchs/data/databriefs/db322-h.pdf> (accessed April 2020).
- Vercammen KA, Frelief JM, Moran AJ *et al.* (2019) Calorie and nutrient profile of combination meals at U.S. fast food and fast casual restaurants. *Am J Prev Med* **57**, e77–e85.
- Powell LM & Nguyen BT (2013) Fast-Food and full-service restaurant consumption among children and adolescents: effect on energy, beverage, and nutrient intake. *JAMA Pediatr* **167**, 14–20.
- Nguyen BT & Powell LM (2014) The impact of restaurant consumption among US adults: effects on energy and nutrient intakes. *Public Health Nutr* **17**, 2445–2452.
- Liu J, Rehm CD, Micha R *et al.* (2020) Quality of meals consumed by US adults at full-service and fast-food restaurants, 2003–2016: persistent low quality and widening disparities. *J Nutr* **150**, 873–883.
- Sharkey JR, Johnson CM & Dean WR (2011) Less-healthy eating behaviors have a greater association with a high level of sugar-sweetened beverage consumption among rural adults than among urban adults. *Food Nutr Res* **55**, 5819.
- Malik VS, Popkin BM, Bray GA *et al.* (2010) Sugar-sweetened beverages, obesity, type 2 diabetes mellitus, and cardiovascular disease risk. *Circulation* **121**, 1356–1364.
- Bowman S, Clemens J, Friday J *et al.* (2019) *Added Sugars in Adults' Diet: What We Eat in America, NHANES 2015–2016*. Washington, DC: United States Department of Agriculture, Food Surveys Research Group.
- Bowman S, Clemens J, Friday J *et al.* (2019) *Added Sugars in American Children's Diet: What We Eat in America, NHANES 2015–2016*. Washington, DC: United States Department of Agriculture, Food Surveys Research Group.
- Haws KL, Liu PJ, Dallas SK *et al.* (2020) Any size for a dollar: the effect of any-size-same-price versus standard pricing on beverage size choices. *J Consum Psychol* **30**, 392–401.
- Nestle M (2015) *Taking on Big Soda (and Winning)*. New York, NY: Oxford University Press.
- Global Food Research Program & University of North Carolina (2019) Sugary drink taxes around the world. https://globalfoodresearchprogram.web.unc.edu/wp-content/uploads/sites/10803/2020/08/SugaryDrink_tax_maps_2020_August_REV.pdf (accessed August 2020).
- World Bank (2020) *Taxes on Sugar-Sweetened Beverages: Summary of International Evidence and Experiences*. Washington, DC: World Bank.
- Cawley J, Crain C, Frisvold D *et al.* (2018) *The Pass-Through of the Largest Tax on Sugar-Sweetened Beverages: The Case of Boulder, Colorado*. Cambridge, MA: National Bureau of Economic Research.
- Public Health-Seattle & King County (2019) *6 Month Report: Store Audits – The Evaluation of Seattle's Sweetened Beverage Tax*. Seattle, WA: Public Health Seattle & King County. <https://www.seattle.gov/Documents/Departments/CityAuditor/auditreports/6%20Month%20Store%20Audit%20Report%20.pdf> (accessed January 2020).
- Saelens BE, Rowland MG, Qu P *et al.* (2020) *Twelve Month Report: Store Audits & Child Cohort – The Evaluation of Seattle's Sweetened Beverage Tax*. Seattle, WA: Public Health Seattle & King County. http://www.seattle.gov/Documents/Departments/CityAuditor/auditreports/SBT_12MonthReport.pdf (accessed April 2020).
- Marinello S, Pipito AA, Leider J *et al.* (2020) The impact of the Oakland sugar-sweetened beverage tax on bottled soda and fountain drink prices in fast-food restaurants. *Prev Med Rep* **17**, 101034.
- Illinois Prevention Research Center at the University of Illinois at Chicago (n.d.) *Beverage Tax Fast-Food Restaurant Observation Form*. <https://p3rc.uic.edu/resources/tools/> (accessed July 2020).
- Li Y, Leider J, Pipito AA *et al.* (2019) *Development and Reliability Testing of a Fast-Food Restaurant Observation Form for Use in Beverage Tax Evaluations*. Chicago, IL: Illinois Prevention Research Center, University of Illinois at Chicago. Research Brief Number 112. https://p3rc.uic.edu/wp-content/uploads/sites/561/2019/11/Development-Reliability-Testing-of-Fast-Food-Observation-Form_Illinois-PRC-Brief-No-112.pdf (accessed May 2020).
- Falbe J, Lee MM, Kaplan S *et al.* (2020) Higher sugar-sweetened beverage retail prices after excise taxes in Oakland and San Francisco. *Am J Public Health* **110**, 1017–1023.
- Cawley J, Frisvold D, Hill A *et al.* (2020) Oakland's sugar-sweetened beverage tax: impacts on prices, purchases and consumption by adults and children. *Econ Hum Biol* **37**, 100865.