

A NOTE ON THE OPTIMAL REPLACEMENT PROBLEM

LAM YEH,* *The Chinese University of Hong Kong*

Abstract

In this note, we study a new repair replacement model for a deteriorating system, in which the successive survival times of the system form a geometric process and are stochastically non-increasing, whereas the consecutive repair times after failure also constitute a geometric process but are stochastically non-decreasing. Two kinds of replacement policy are considered, one based on the working age of the system and the other one determined by the number of failures. The explicit expressions of the long-run average costs per unit time under these two kinds of policy are calculated.

REPLACEMENT PROBLEM; STOCHASTIC MONOTONICITY; RENEWAL PROCESSES

1. Introduction

Most repair replacement models assume that a failure system after repair will yield a functioning system which is ‘as good as new’ i.e., the system is renewed after each failure. An alternative assumption is to assume that a failure system after repair will function again, but has the same failure rate and the same effective age at the time of failure: this leads to the minimal repair model. In both cases, the repair times are often negligible (see e.g. Barlow and Proschan (1965) or Ascher and Feingold (1984)), although there exist a large number of repair replacement models where the repair times are taken into account (see e.g. Morimura (1970)).

However, for deteriorating systems, the problem is different from that described above. For example, in machine maintenance problems, after repair, because of the deterioration, the functioning time of a machine will become shorter and shorter so the total functioning time or the total life of the machine must be finite. On the other hand, in view of the ageing and accumulative wear, the repair time will become longer and longer and tend to infinity: i.e., finally the machine is non-repairable. Therefore, we have to consider a repair replacement model for deteriorating systems as follows: the successive survival times are non-increasing and will eventually die out, so that the total life of the system is finite; moreover, the consecutive repair times are non-decreasing and will finally tend to infinity.

One possible way of modelling these deteriorating systems is to use the non-homogeneous Poisson process. Another possible way is to introduce the following geometric process.

Definition 1. Given a sequence of random variables X_1, X_2, \dots if for some $a > 0$, $\{a^{n-1}X_n, n = 1, 2, \dots\}$ forms a renewal process, then $\{X_n, n = 1, 2, \dots\}$ is a geometric process. a is called the parameter of the geometric process.

Definition 2. A geometric process is called a non-increasing geometric process if $a \geq 1$ and a non-decreasing geometric process if $a \leq 1$.

Note, if $a = 1$, then the geometric process reduces to a renewal process.

It is easy to show that for a geometric process, if $a > 1$, then it is stochastically decreasing and converges to 0 in probability; furthermore, if $a > 1$, then $U = \sum_{n=1}^{\infty} X_n$ converges in the

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* Postal Address: Department of Statistics, The Chinese University of Hong Kong, Shatin, N.T., Hong Kong.

sense of mean square with $P(U < \infty) = 1$. Whereas, if $a < 1$, then it is stochastically increasing and tends to infinity in probability (see Lam Yeh (1987) for more details). Therefore, for the deteriorating systems considered above, if its successive survival times are increasing geometrically with a constant rate, then we can model them by using a non-increasing geometric process. Otherwise, if its successive survival times are decreasing in a general way, we can still use a non-increasing geometric process as an approximation for modelling the successive survival times; this is exactly like using a linear function as an approximation of a general function. Similarly, we can use a non-decreasing geometric process for formulating its consecutive repair times either precisely or approximately.

2. Model

Now, we are able to state the new repair replacement model as below.

Assumption 1. At the beginning, a new system is used.

Assumption 2. Whenever the system fails, we can repair it. Let X_n be the survival time after the $(n - 1)$ th repair, then $\{X_n, n = 1, 2, \dots\}$ forms a non-increasing geometric process with parameter $a \geq 1$ and $E(X_1) = \lambda > 0$.

Assumption 3. Let Y_n be the repair time after the n th failure, then $\{Y_n, n = 1, 2, \dots\}$ constitutes a non-decreasing geometric process with parameter $0 < b \leq 1$ and $E(Y_1) = \mu \geq 0, \mu = 0$ means that the repair time is negligible.

Assumption 4. $\{X_n, n = 1, 2, \dots\}$ and $\{Y_n, n = 1, 2, \dots\}$ are independent.

Assumption 5. The repair cost rate is C , the reward rate whenever the system is working is r , without loss of generality, we assume that $r = 1$.

Two kinds of replacement policy are considered in this model.

(i) A replacement policy T is a policy of which we replace the system whenever the working age of the system reaches T .

The working age T of a system at time t is the cumulative survival time by time t , i.e.,

$$(1) \quad T = \begin{cases} t - V_n, & U_n + V_n \leq t < U_{n+1} + V_n \\ U_{n+1}, & U_{n+1} + V_n \leq t < U_{n+1} + V_{n+1} \end{cases}$$

where

$$U_n = \sum_{i=1}^n X_i, \quad V_n = \sum_{i=1}^n Y_i \quad \text{and} \quad U_0 = 0, \quad V_0 = 0.$$

(ii) A replacement policy N is a policy of which we replace the system at the time of N th failure since the last replacement.

Assumption 6. The system may be replaced sometime by a new and identical one. The replacement cost under policy T is C_1 ; the replacement cost under policy N is C_2 .

Under replacement policy T or N , the problem is to determine an optimal replacement policy T^* or N^* respectively, such that the long-run average cost per unit time is minimized.

3. Long-run average cost

Let T_1 be the first replacement time, in general, for $n \geq 2$, let T_n be the time between $(n - 1)$ th replacement and n th replacement, then obviously, $\{T_n, n = 1, 2, \dots\}$ forms a renewal process. It is well known that the long-run average cost per unit time is given by

$$(2) \quad \frac{\text{the expected cost incurred in a cycle}}{\text{the expected length of a cycle}}$$

where a cycle is the time between two consecutive replacements (see, e.g. Ross (1970)).

Under the replacement policy T , denote the length of a cycle by W , then

$$(3) \quad W = T + V_K, \quad U_K < T \leq U_{K+1}, \quad K = 0, 1, \dots$$

From Definition 1 and Assumption 3, it follows that $E(Y_k) = \mu/b^{k-1}$. Then

$$\begin{aligned} E(W) &= E\left(T + \sum_{k=1}^K Y_k\right) \\ &= E\left[T + \sum_{k=1}^{\infty} Y_k I(U_k < T)\right] \\ &= T + \mu \sum_{k=1}^{\infty} \frac{1}{b^{k-1}} F_k(T - 0) \end{aligned}$$

where I is the indicator and F_k is the distribution of U_k .

Thus, from (2) the long-run average cost per unit time $C_1(T)$ under policy T is given by

$$(4) \quad C_1(T) = \frac{C\mu \sum_{k=1}^{\infty} \frac{1}{b^{k-1}} F_k(T - 0) + C_1 - T}{T + \mu \sum_{k=1}^{\infty} \frac{1}{b^{k-1}} F_k(T - 0)} = C_1^*(T) - 1,$$

where

$$(5) \quad C_1^*(T) = \frac{(C + 1)\mu \sum_{k=1}^{\infty} \frac{1}{b^{k-1}} F_k(T - 0) + C_1}{T + \mu \sum_{k=1}^{\infty} \frac{1}{b^{k-1}} F_k(T - 0)}.$$

Similarly, under replacement policy N , the long-run average cost per unit time $C_2(N)$ is given by

$$(6) \quad C_2(N) = \frac{C\mu \sum_{k=1}^{N-1} \frac{1}{b^{k-1}} + C_2 - \lambda \sum_{k=1}^N \frac{1}{a^{k-1}}}{\lambda \sum_{k=1}^N \frac{1}{a^{k-1}} + \mu \sum_{k=1}^{N-1} \frac{1}{b^{k-1}}}, = C_2^*(N) - 1$$

where

$$(7) \quad C_2^*(N) = \frac{(C + 1)\mu \sum_{k=1}^{N-1} \frac{1}{b^{k-1}} + C_2}{\lambda \sum_{k=1}^N \frac{1}{a^{k-1}} + \mu \sum_{k=1}^{N-1} \frac{1}{b^{k-1}}}, \quad N = 1, 2, \dots$$

Finally, we can determine the optimal replacement policy T^* and N^* by minimizing $C_1(T)$ (or $C_1^*(T)$) and $C_2(N)$ (or $C_2^*(N)$) respectively. Furthermore, the minimization procedure can be done by analytical or numerical methods.

In practice, we prefer to adopt the optimal policy N^* rather than use the optimal policy T^* , because of the much simpler form of (6) or (7).

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